

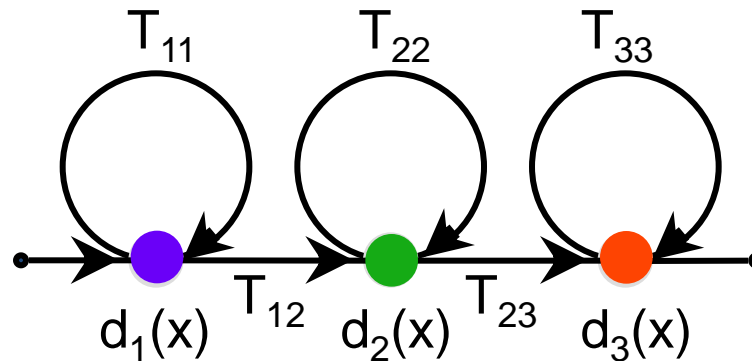
# Design and Implementation of Speech Recognition Systems

*Spring 2013*

Class 8/10: HMMs  
18/25 Feb 2013

# Recap: Generalized Templates

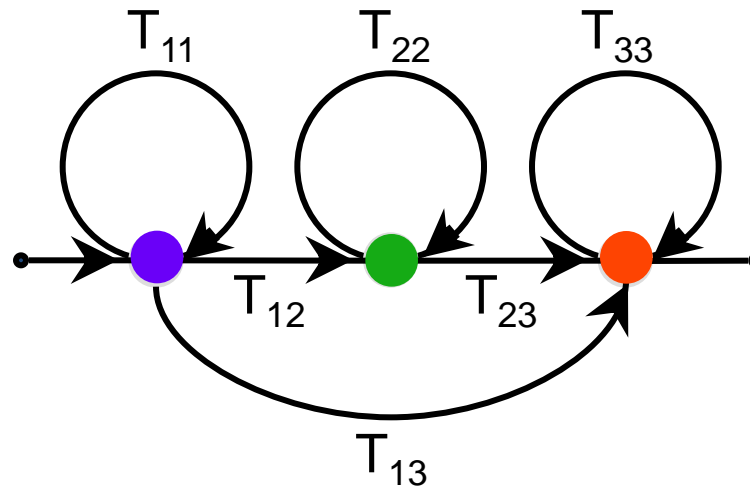
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- A set of “states”
  - A distance function associated with each state
- A set of transitions
  - Transition-specific penalties

# Recap: HMMs

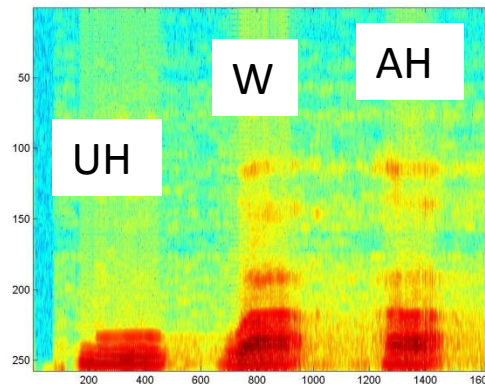
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- Identical to generalized templates in principle
- “Distance” functions at states replaced by “probability distribution function” for state
- Transition “penalties” replaced by transition probabilities
- Maximize probability of observation
  - Instead of minimizing cost
- The entire structure may be viewed as *one* generalization of the DTW models we have discussed thus far

# The HMM Process

- The HMM models the process underlying the observations as going through a number of states
  - E.g., to produce the sound “W”, it first goes through a state where it produces the sound “UH”, then goes into a state where it transitions from “UH” to “AH”, and finally to a state where it produced “AH”



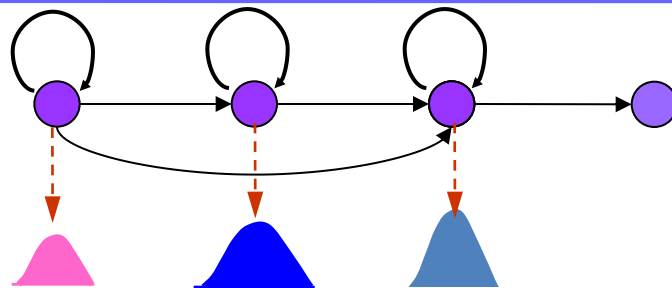
- The true underlying process is the vocal tract here
  - Which roughly goes from the configuration for “UH” to the configuration for “AH”

# HMMs are abstractions

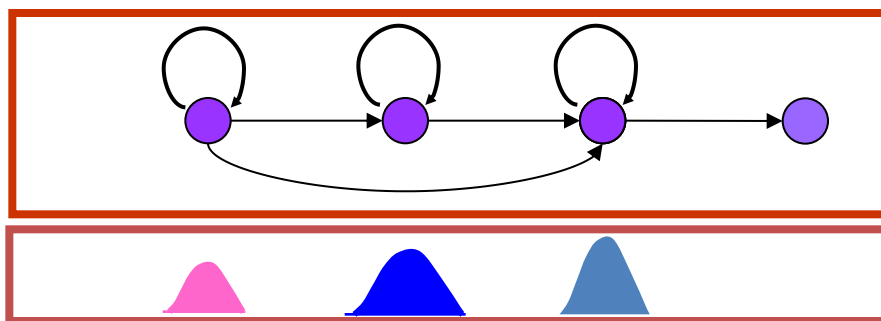
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- The states are not directly observed
  - Here states of the process are analogous to configurations of the vocal tract that produces the signal
  - We only hear the speech; we do not see the vocal tract
  - i.e. the states are *hidden*
- The interpretation of states is not always obvious
  - The vocal tract actually goes through a *continuum* of configurations
  - The model represents all of these using only a fixed number of states
- The model *abstracts* the process that generates the data
  - The system goes through a finite number of states
  - When in any state it can either remain at that state, or go to another with some probability
  - When at any states it generates observations according to a distribution associated with that state

# Hidden Markov Models



- A Hidden Markov Model consists of two components
  - A state/transition backbone that specifies how many states there are, and how they can follow one another
  - A set of probability distributions, one for each state, which specifies the distribution of all vectors in that state



Markov chain

Data distributions

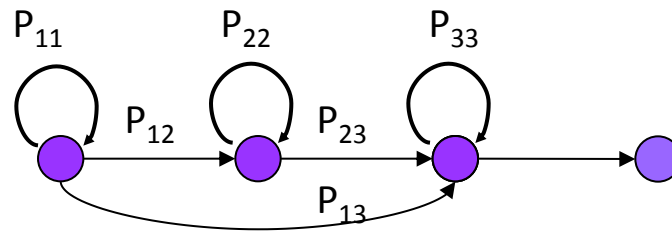
- This can be factored into two separate probabilistic entities
  - A probabilistic Markov chain with states and transitions
  - A set of data probability distributions, associated with the states

# Equivalence to DTW templates

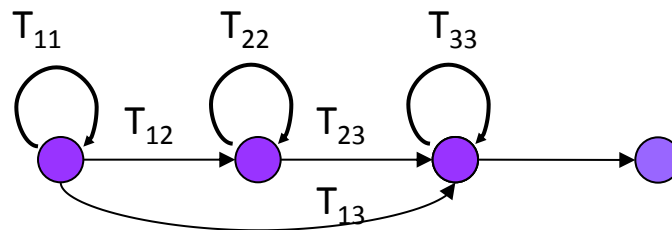
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- HMM – inference equivalent to DTW modified to use a *probabilistic* function, for the local node or edge “costs” in the trellis
  - Edges have transition probabilities
  - Nodes have *output* or *observation probabilities*
    - They provide the probability of the observed input
    - The output probability may be a Gaussian
  - Goal is to find the template with highest probability of matching the input
- Probability values associated with transitions and edges are called *likelihoods*

# Likelihoods and Cost: Transition



Markov chain



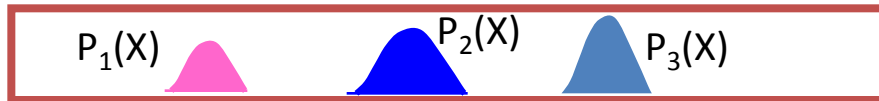
DTW Template

- Transitions in the HMM have associated probabilities
  - $P_{11}$ ,  $P_{12}$  etc
- They can be converted to “scores” through a logarithm
  - $T_{11} = \log(P_{11})$
- Or to “costs” through a negative logarithm
  - $T_{11} = -\log(P_{11})$



# Likelihoods and Cost: Nodes

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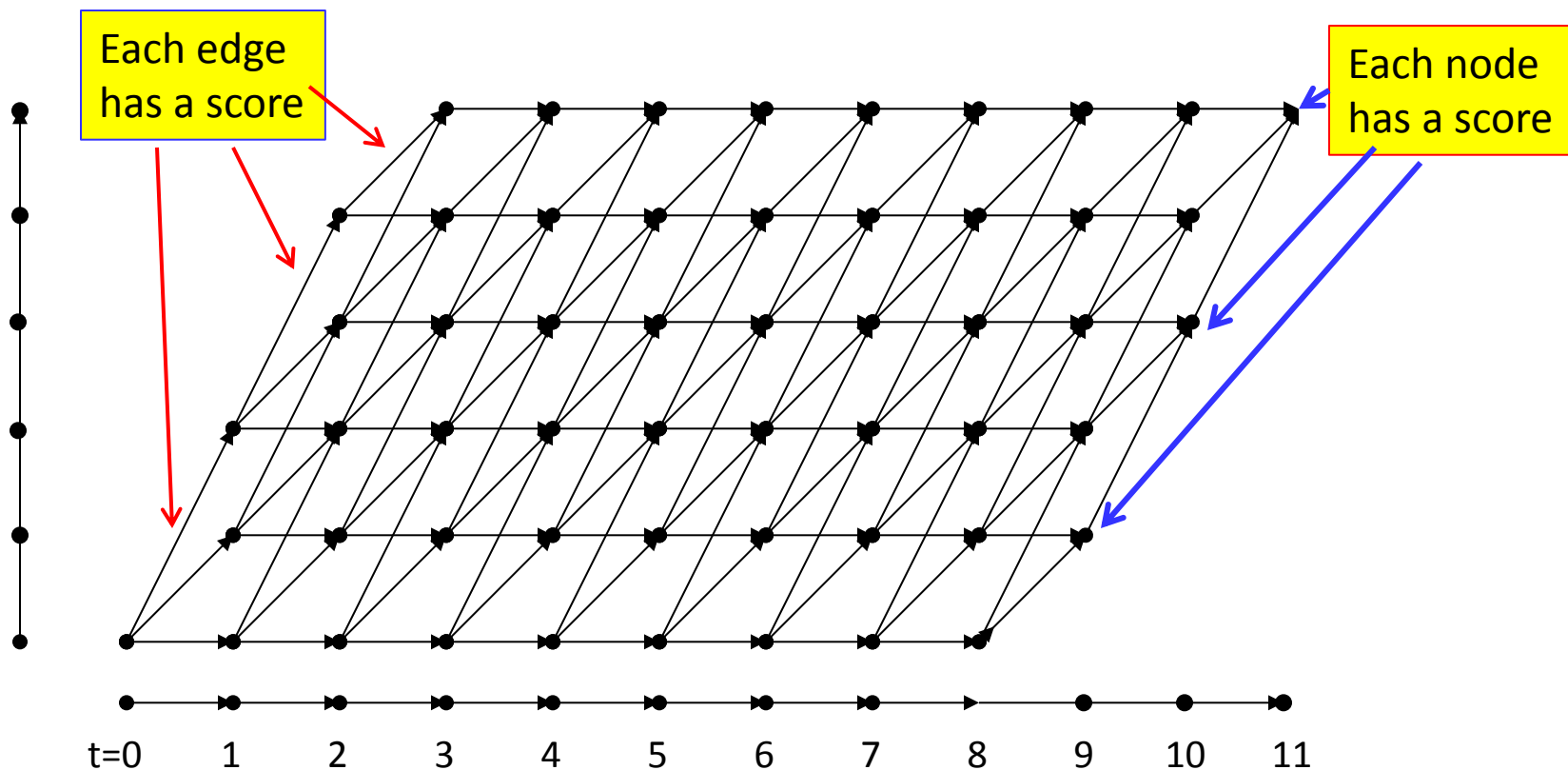


Data distributions

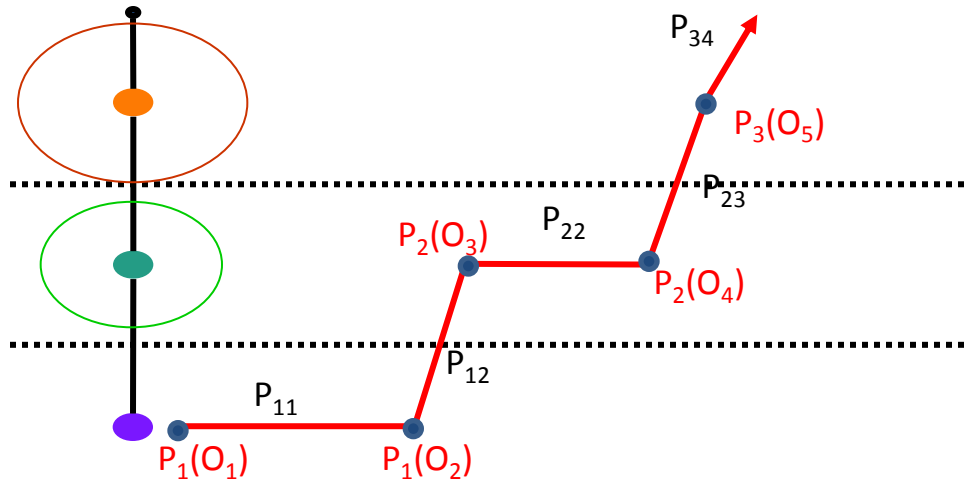
- States in the HMM have probability distributions associated with them
  - E.g Gaussians
    - Whose means and variances have been obtained from the segments associated with the node
- Nodes in the *trellis* have a probabilities associated with them
  - $P_i(O)$
  - $i$  is the “state” / template node
  - $O$  is the observation associated with any node in the trellis
- Node probabilities may be converted to:
  - Scores:  $N_i(O) = \log(P_i(O))$
  - Or Costs:  $N_i(O) = -\log(P_i(O))$

# Computation is still done with a Trellis

- Node and edge *scores* defined for trellis



# Log Likelihoods



- Use probabilities or likelihoods instead of cost
  - Scores combine multiplicatively along a path
  - Path Score =  $P_1(O_1) \cdot P_{11} \cdot P_1(O_2) \cdot P_{12} \cdot P_2(O_3) \cdot P_{22} \cdot P_2(O_4) \cdot P_{23} \cdot P_3(O_5) \cdot P_{23}$
- Alternately use log probabilities as scores:  $N_i(O) = \log(P_i(O))$ ,  $T_{11} = \log(P_{11})$ 
  - Scores add as in DTW
  - Path Score =  $N_1(O_1) + T_{11} + N_1(O_2) + T_{12} + N_2(O_3) + T_{22} + N_2(O_4) + T_{23} + N_3(O_5) + T_{23}$ 
    - Replace all “Min” operations in DTW by “Max”
- Alternately use *negative* log probabilities as cost:  $N_i(O) = \log(P_i(O))$ ,  $T_{11} = -\log(P_{11})$ 
  - Cost adds as in DTW
  - Computation remains identical to DTW (with edge costs factored in)

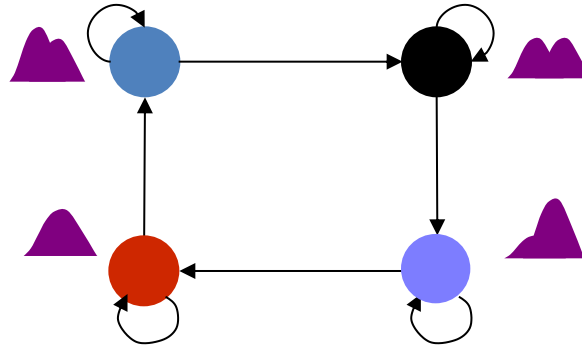
# HMM as a statistical model

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- An HMM is a statistical model for a time-varying process
- The process is always in one of a countable number of
- When the process visits in any state, it generates an observation by a random draw from a distribution associated with that state
- The process constantly moves from state to state. The probability that the process will move to any state is determined solely by the current state
  - i.e. the dynamics of the process are Markovian
- The entire model represents a probability distribution over the sequence of observations
  - It has a specific probability of generating any particular sequence
  - The probabilities of all possible observation sequences sums to 1

# How an HMM models a process

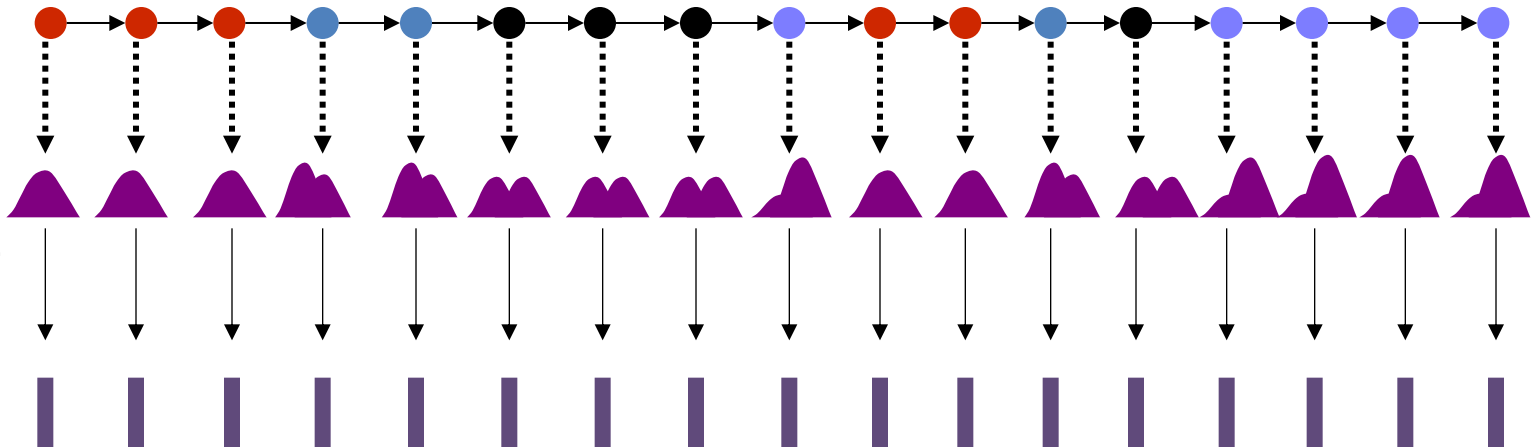
HMM assumed to be  
generating data



state  
sequence

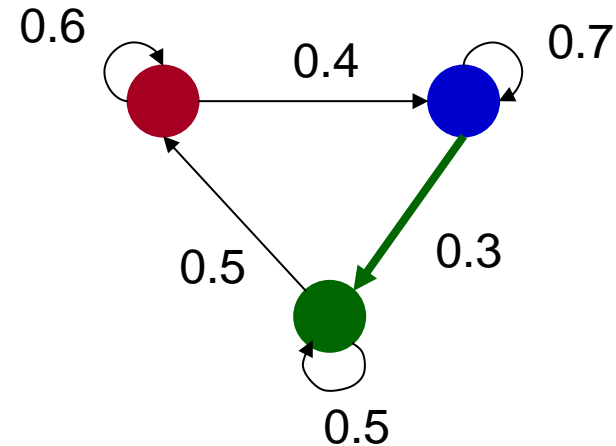
state  
distributions

observation  
sequence

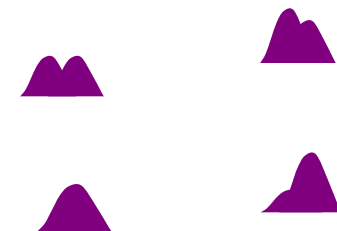


# HMM Parameters

- The *topology* of the HMM
  - No. of states and allowed transitions
  - E.g. here we have 3 states and cannot go from the blue state to the red
- The transition probabilities
  - Often represented as a matrix as here
  - $T_{ij}$  is the probability that when in state  $i$ , the process will move to  $j$
- The probability of being at a particular state at the first instant
- The *state output distributions*



$$T = \begin{pmatrix} .6 & .4 & 0 \\ 0 & .7 & .3 \\ .5 & 0 & .5 \end{pmatrix}$$



# HMM state output distributions

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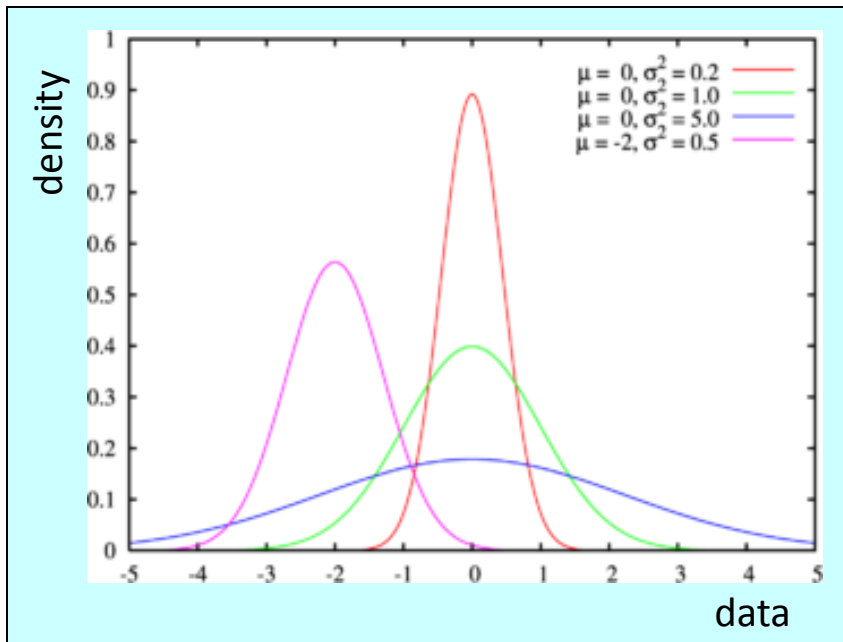
- The state output distribution represents the distribution of data produced from any state
- In the previous lecture we assume the state output distribution to be Gaussian
  - Albeit largely in a DTW context

$$P(v) = \text{Gaussian}(v; m, C) = \frac{1}{\sqrt{2\pi|C|}} e^{-0.5(v-m)^T C^{-1}(v-m)}$$

- In reality, the distribution of vectors for any state need not be Gaussian
  - In the most general case it can be arbitrarily complex
  - The Gaussian is only a coarse representation of this distribution
- If we model the output distributions of states better, we can expect the model to be a better representation of the data

# Node Score: The Gaussian Distribution

- What does a Gaussian distribution look like?
- For a single (scalar) variable, it is a bell-shaped curve representing the density of data around the mean
- Example:



Four different scalar Gaussian distributions, with different means and variances

The mean is represented by  $\mu$ , and variance by  $\sigma^2$

$\mu$  and  $\sigma$  are the *parameters* of the Gaussian distribution

(Taken from Wikipedia)



# The Scalar Gaussian Function

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- The Gaussian density function (the bell curve) is:

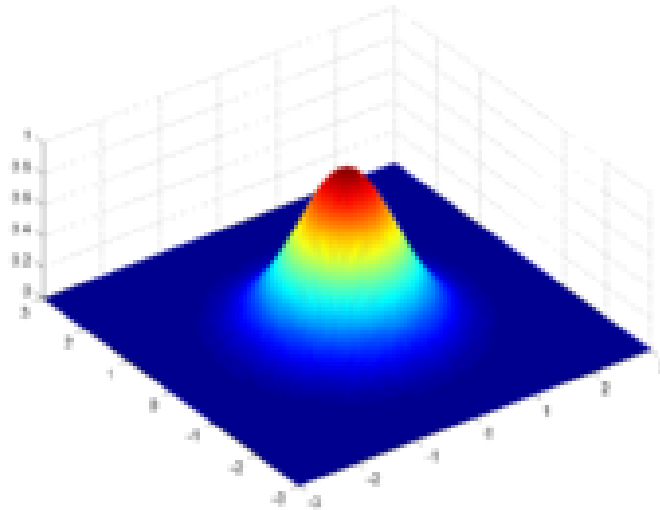
$$p(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2\sigma^2}(x-\mu)^2}$$

- $p(x)$  is the density function of the variable  $x$ , with mean  $\mu$  and variance  $\sigma^2$
- The attraction of the Gaussian function (regardless of how appropriate it is!) comes from how easily the mean and variance can be estimated from *sample data*  $x_1, x_2, x_3 \dots x_N$ 
  - $\mu = \sum_i x_i / N$
  - $\sigma^2 = \sum_i (x_i - \mu)^2 / N = \sum_i (x_i^2 - \mu^2) / N$

# The 2-D Gaussian Distribution

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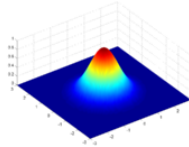
- Speech data are not scalar values, but vectors!
- Needs multi-variate (multi-dimensional) Gaussians
- Figure: A Gaussian for 2-D data
  - Shown as a 3-D plot



- Distributions for higher dimensions are tough to visualize!

# The Multidimensional Gaussian Distribution

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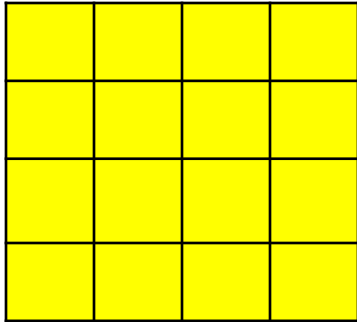
- Instead of variance, the multidimensional Gaussian has a *covariance matrix*
- The multi-dimensional Gaussian distribution of a vector variable  $x$  with mean  $\mu$  and covariance  $\Sigma$  is given by:

$$f(x) = \frac{1}{\sqrt{(2\pi)^D |C|}} \exp\left(-0.5(x - \mu)^T C^{-1}(x - \mu)\right)$$

- where  $D$  is the vector dimensionality
- The complexity in a full multi-dimensional Gaussian distribution comes from the covariance matrix, which accounts for *dependencies* between the dimensions

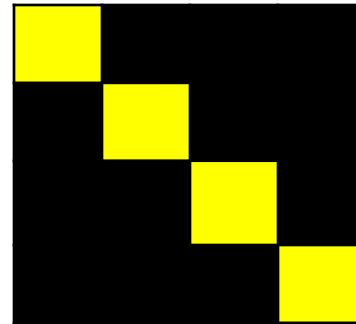
# The Diagonal Covariance Matrix

Full covariance:  
all elements are  
non-zero



$$-0.5(\mathbf{x}-\boldsymbol{\mu})^T \mathbf{C}^{-1}(\mathbf{x}-\boldsymbol{\mu})$$

Diagonal covariance:  
off-diagonal elements  
are zero



$$-\sum_i (x_i - \mu_i)^2 / 2\sigma_i^2$$

- In speech recognition, we frequently assume that the feature vector dimensions are all *independent* of each other
- *Result:* The covariance matrix is reduced to a diagonal form
  - The exponential term becomes, simply:  
 $(\sum_i (x_i - \mu_i)^2 / \sigma_i^2) / 2$ ,  $i$  going over all vector dimensions
  - The determinant of the diagonal  $\Sigma$  matrix is easy to compute
- Further, each  $\sigma_i^2$  (the  $i$ -th diagonal element in the covariance matrix) is easily estimated from  $x_i$  and  $\mu_i$  like a scalar

# Gaussian Mixtures

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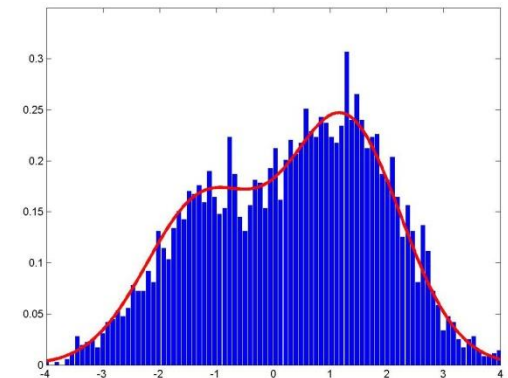
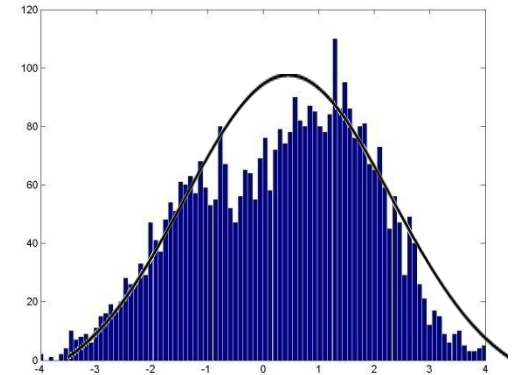
- A Gaussian Mixture is literally a mixture of Gaussians. It is a weighted combination of several Gaussian distributions

$$P(v) = \sum_{i=0}^{K-1} w_i \text{Gaussian}(v; \mu_i, C_i)$$

- $v$  is any data vector.  $P(v)$  is the probability given to that vector by the Gaussian mixture
- $K$  is the number of Gaussians being mixed
- $w_i$  is the mixture weight of the  $i^{\text{th}}$  Gaussian.  $\mu_i$  is its mean and  $C_i$  is its covariance
- The Gaussian mixture distribution is also a distribution
  - It is positive everywhere.
  - The total volume under a Gaussian mixture is 1.0.
  - Constraint: the mixture weights  $w_i$  must all be positive and sum to 1

# Gaussian Mixtures

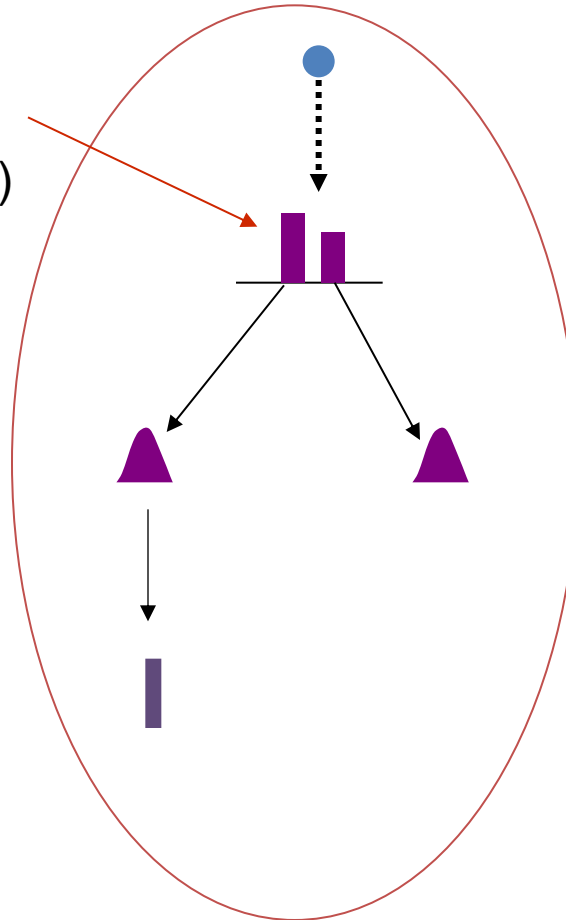
- A Gaussian mixture can represent data distributions far better than a simple Gaussian
- The two panels show the histogram of an unknown random variable
- The first panel shows how it is modeled by a simple Gaussian
- The second panel models the histogram by a mixture of two Gaussians
- Caveat: It is hard to know the optimal number of Gaussians in a mixture distribution for any random variable



# Generating an observation from a Gaussian mixture state distribution

First draw the identity of the Gaussian from the a priori probability distribution of Gaussians (mixture weights)

Then draw a vector from the selected Gaussian



# HMMs with Gaussian mixture state distributions

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- The parameters of an HMM with Gaussian mixture state distributions are:
  - $\pi$  the set of initial state probabilities for all states
  - $T$  the matrix of transition probabilities
  - A Gaussian mixture distribution for every state in the HMM. The Gaussian mixture for the  $i^{\text{th}}$  state is characterized by
    - $K_i$ , the number of Gaussians in the mixture for the  $i^{\text{th}}$  state
    - The set of mixture weights  $w_{i,j}$   $0 < j < K_i$
    - The set of Gaussian means  $\mu_{i,j}$   $0 < j < K_i$
    - The set of Covariance matrices  $C_{i,j}$   $0 < j < K_i$



# Three Basic HMM Problems

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- Given an HMM:
  - What is the probability that it will generate a specific observation sequence
  - Given a observation sequence, how do we determine which observation was generated from which state
    - The state segmentation problem
  - How do we *learn* the parameters of the HMM from observation sequences

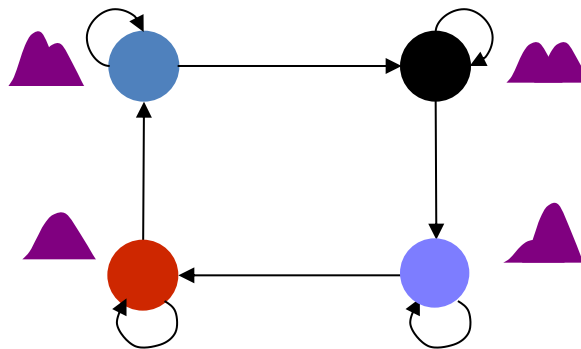
# Computing the Probability of an Observation Sequence

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- Two aspects to producing the observation:
  - Progressing through a sequence of states
  - Producing observations from these states

# Progressing through states

HMM assumed to be  
generating data



state  
sequence



- The process begins at some state (red) here
- From that state, it makes an allowed transition
  - To arrive at the same or any other state
- From that state it makes another allowed transition
  - And so on

# Probability that the HMM will follow a particular state sequence

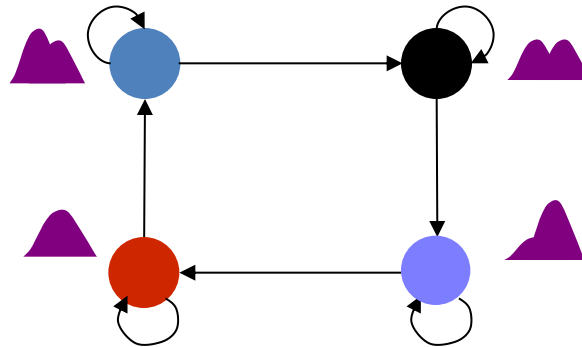
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$$P(s_1, s_2, s_3, \dots) = P(s_1) P(s_2 | s_1) P(s_3 | s_2) \dots$$

- $P(s_1)$  is the probability that the process will initially be in state  $s_1$
- $P(s_i / s_j)$  is the transition probability of moving to state  $s_i$  at the next time instant when the system is currently in  $s_j$ 
  - Also denoted by  $P_{ij}$  earlier
  - Related to edge scores in DTW as  $T_{ij} = -\log(P(s_i / s_j))$

# Generating Observations from States

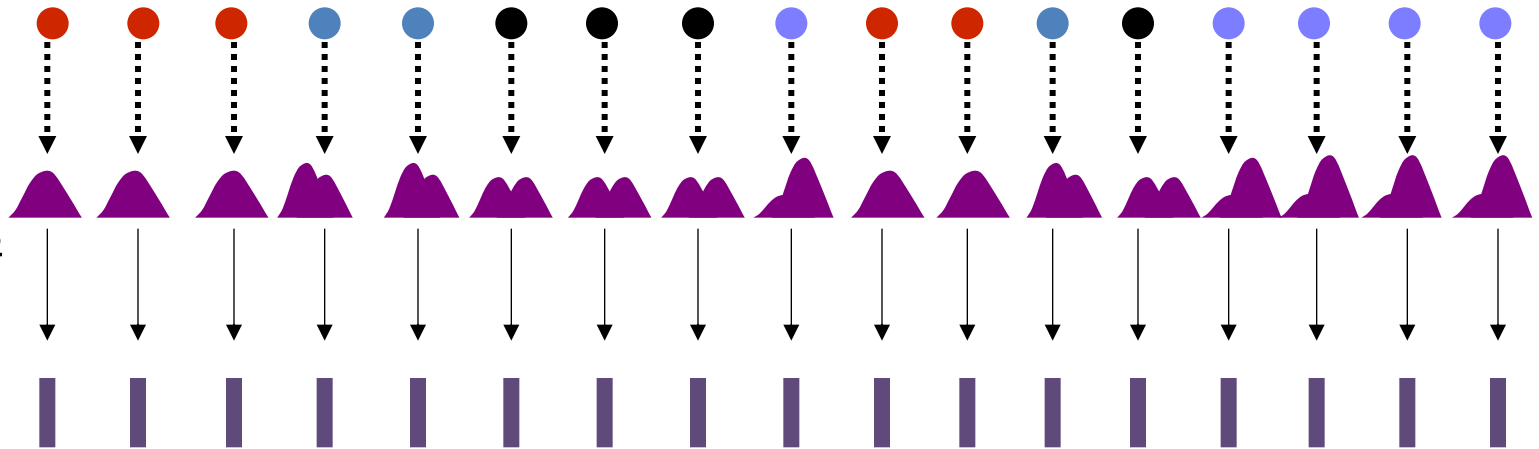
HMM assumed to be  
generating data



state  
sequence

state  
distributions

observation  
sequence



- At each time it generates an observation from the state it is in at that time

Probability that the HMM will generate a particular observation sequence given a state sequence (state sequence known)

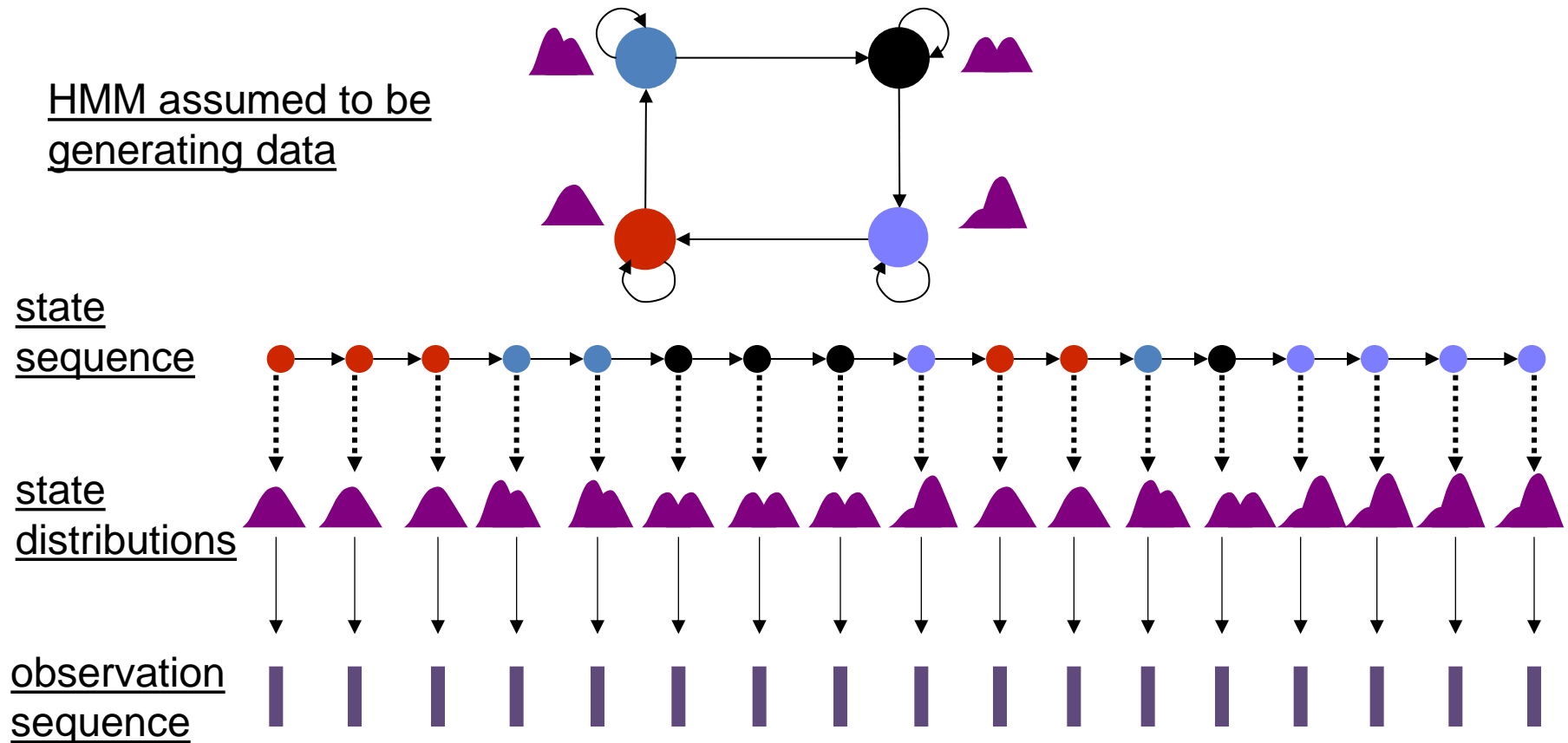
$$P(o_1, o_2, o_3, \dots | s_1, s_2, s_3, \dots) = P(o_1 | s_1) P(o_2 | s_2) P(o_3 | s_3) \dots$$



Computed from the Gaussian or Gaussian mixture for state  $s_1$

- $P(o_i / s_i)$  is the probability of generating observation  $o_i$  when the system is in state  $s_i$ 
  - Related to node scores in DTW trellis as:  
 $N_i(O) = -\log(P(o_i / s_i))$

# Progressing through States and Producing Observations



- At each time it produces an observation and makes a transition

Probability that the HMM will generate a particular state sequence and, from it, generate a particular observation sequence

$$P(o_1, o_2, o_3, \dots, s_1, s_2, s_3, \dots) =$$

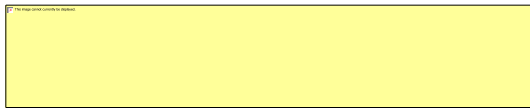
$$P(o_1|s_1)P(o_2|s_2)P(o_3|s_3)\dots P(s_1)P(s_2|s_1)P(s_3|s_2)\dots$$



# Probability of Generating an Observation Sequence

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- If only the observation is known, the precise state sequence followed to produce it is not known
- All possible state sequences must be considered



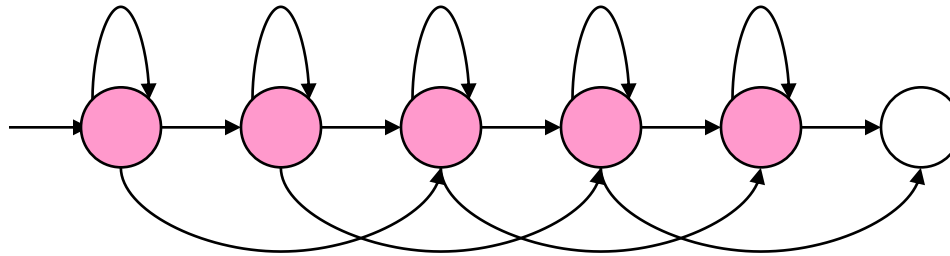
# Computing it Efficiently

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- Explicit summing over all state sequences is not efficient
  - A very large number of possible state sequences
  - For long observation sequences it may be intractable
- Fortunately, we have an efficient algorithm for this:  
The forward algorithm
- At each time, for each state compute the total probability of all state sequences that generate observations until that time and end at that state

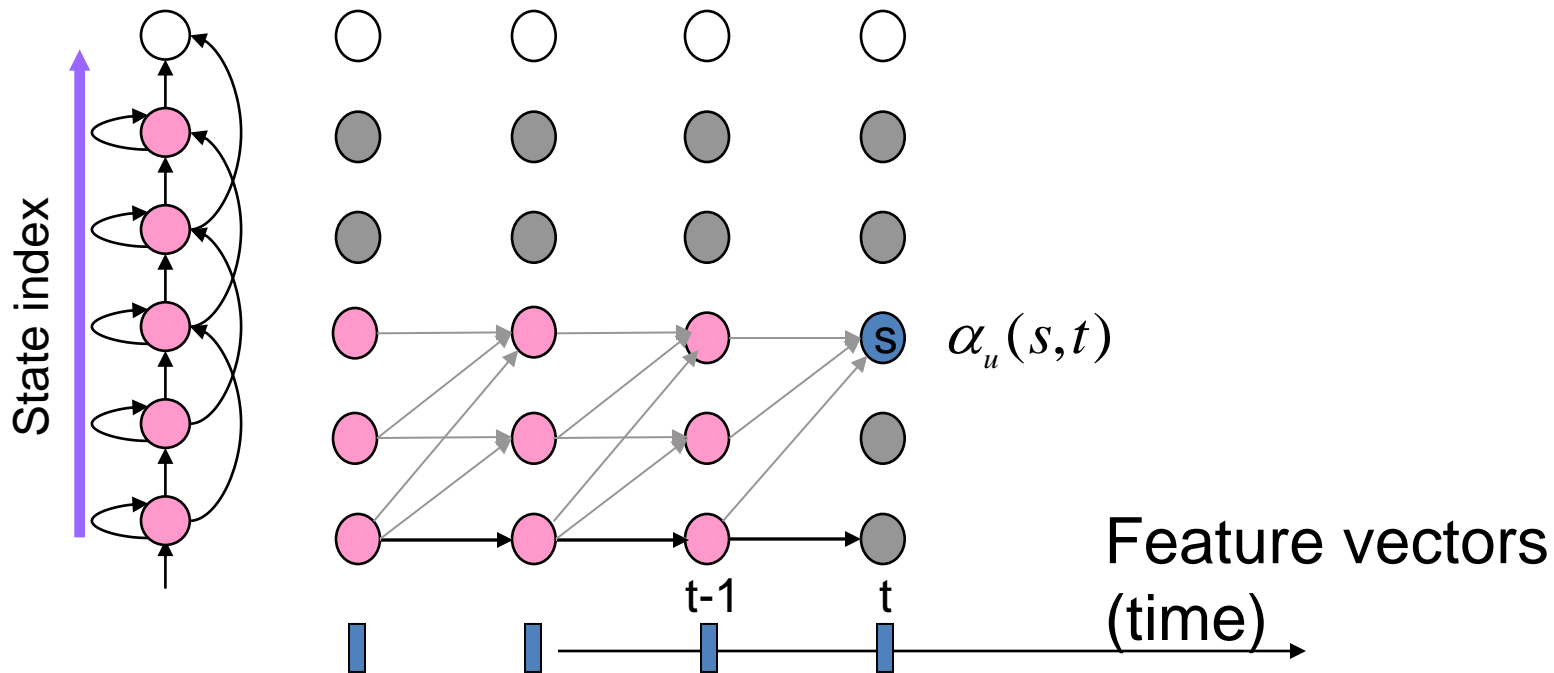
# Illustrative Example

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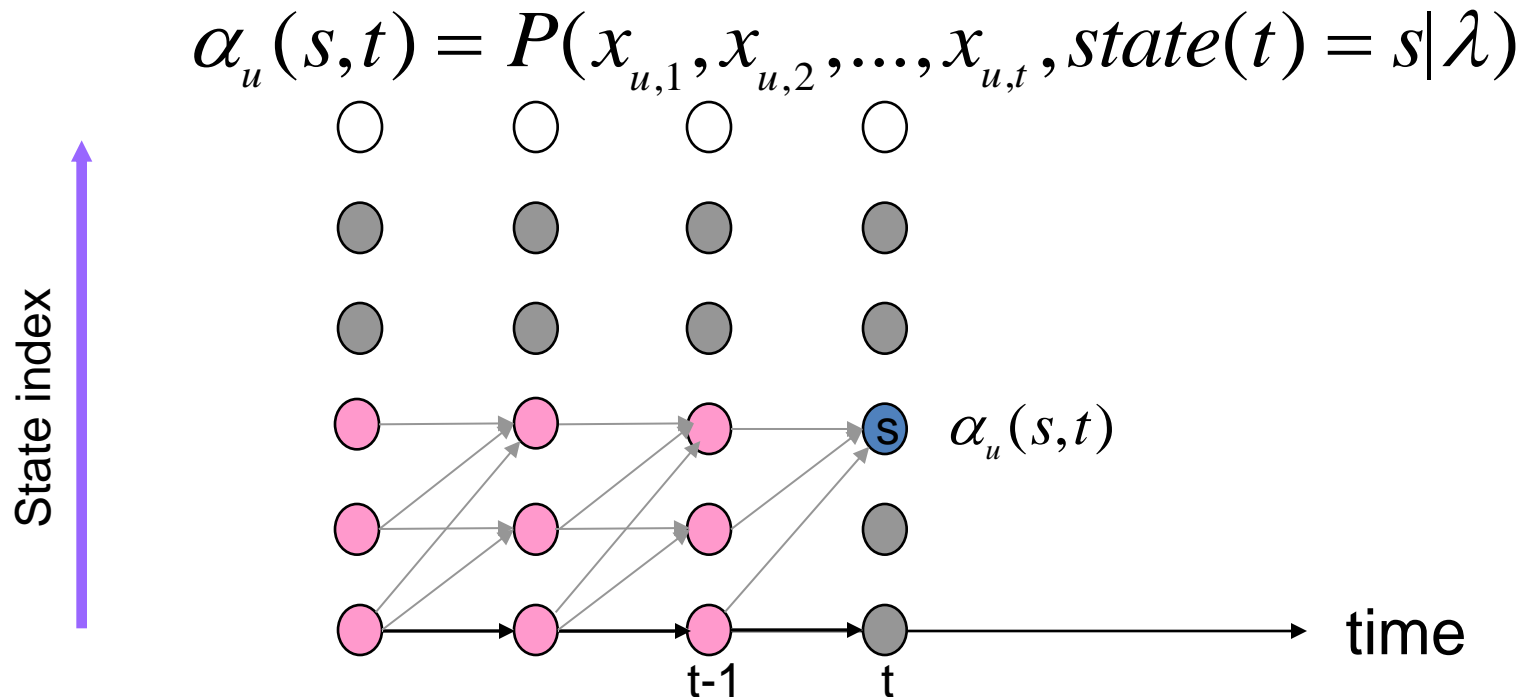
- Consider a generic HMM with 5 states and a “terminating state”. We wish to find the probability of the best state sequence for an observation sequence assuming it was generated by this HMM
  - $P(s_i) = 1$  for state 1 and 0 for others
  - The arrows represent transition for which the probability is not 0.
    - May represent it as  $P(s_i | s_j) = a_{ij}$
    - We sometimes also represent the state output probability of  $s_i$  as  $P(o_t | s_i) = b_i(t)$  for brevity

# Diversion: The HMM Trellis



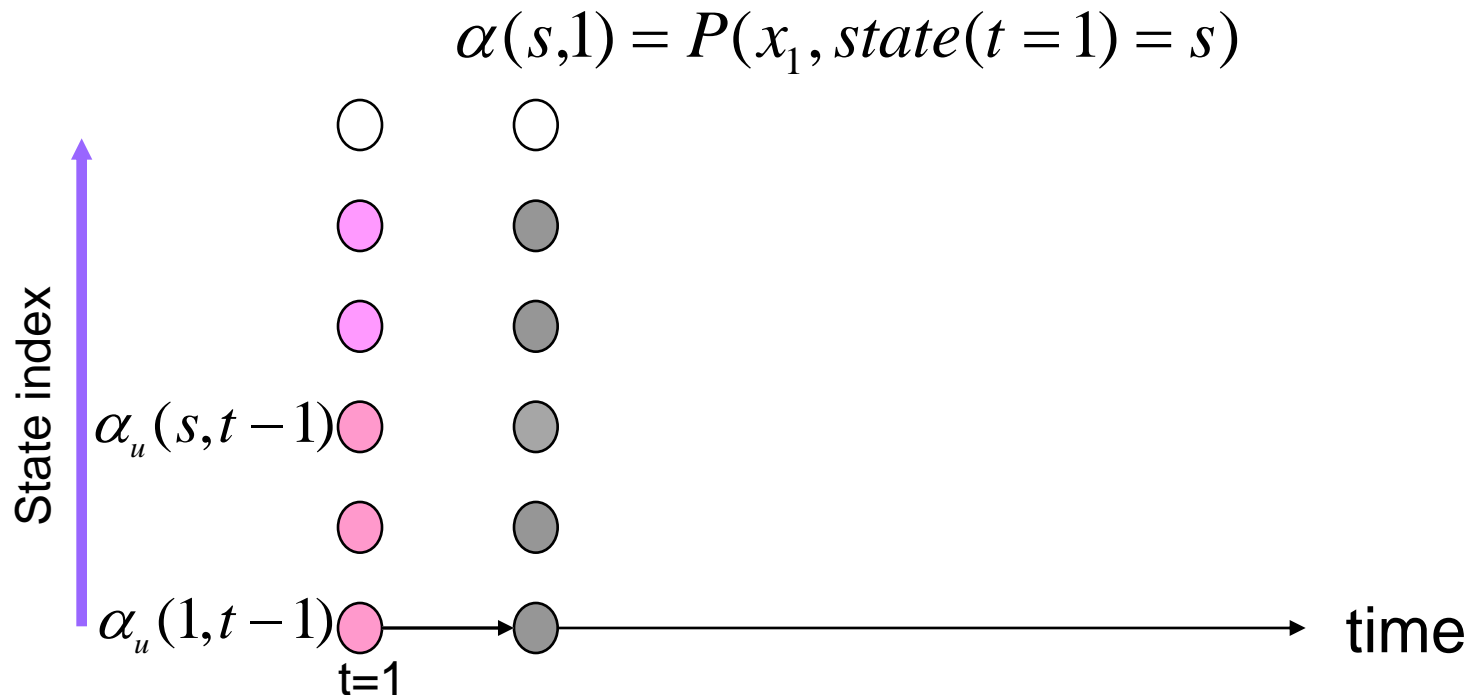
- The trellis is a graphical representation of all possible state sequences through the HMM to produce a given observation
  - Analogous to the DTW search graph / trellis
- The Y-axis represents HMM states, X axis represents observations
- Edges in trellis represent valid transitions in the HMM over a single time step
- Every node represents the event of a particular observation being generated from a particular state

# The Forward Algorithm



- $\alpha_u(s, t)$  is the total probability of ALL state sequences that end at state  $s$  at time  $t$ , and all observations until  $x_t$

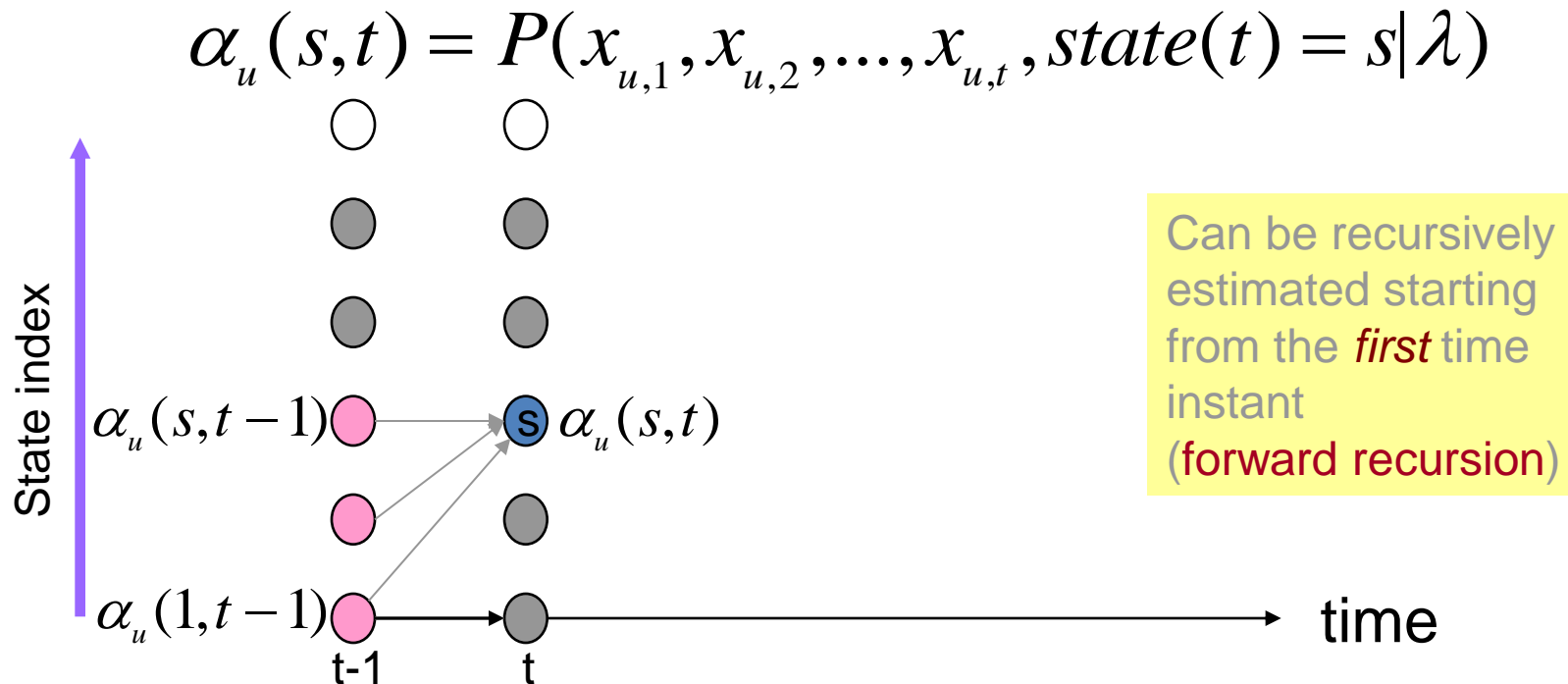
# The Forward Algorithm at $t=1$



$$\alpha(s,1) = \pi(s)P(x_1 | s)$$

- $\alpha(s,1)$  is simply the probability of being in state  $s$  at  $t=1$  *and* generating observation  $x_1$  from  $s$

# The Forward Algorithm

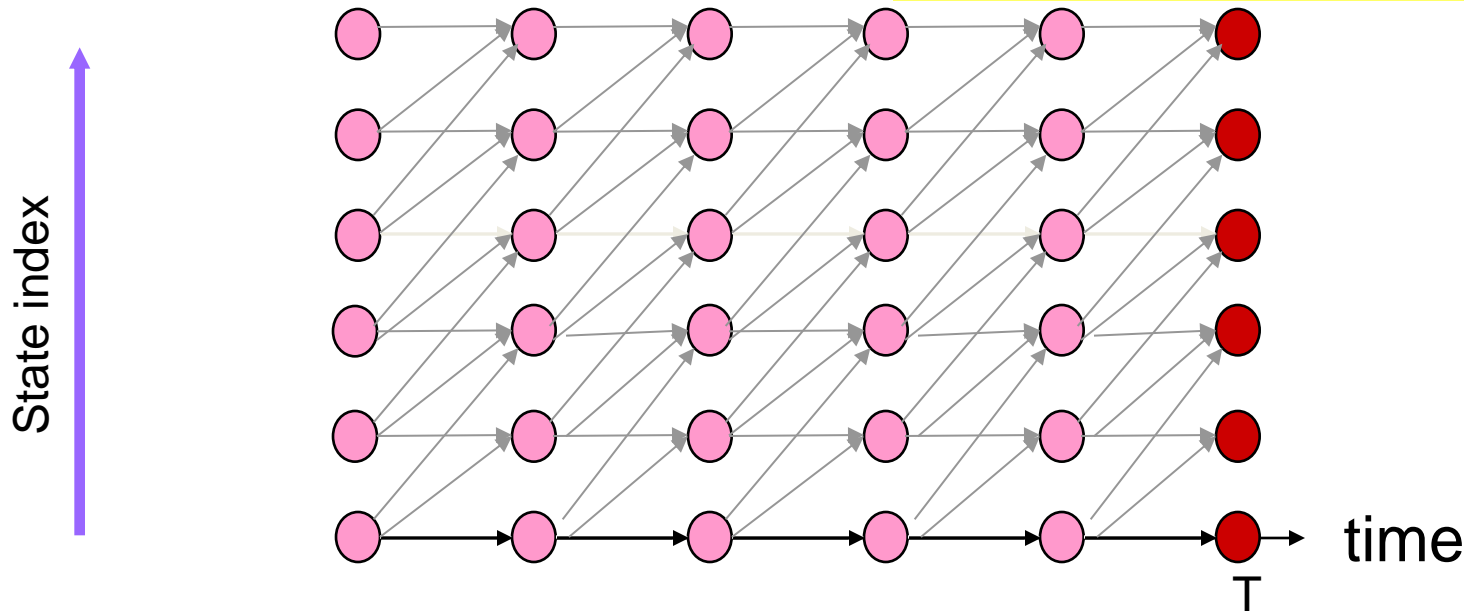


$$\alpha_u(s, t) = \sum_{s'} \alpha_u(s', t-1) P(s|s') P(x_{u,t} | s)$$

- $\alpha_u(s, t)$  can be recursively computed in terms of  $\alpha_u(s', t')$ , the forward probabilities at time t-1

# The Forward Algorithm

$$Totalprob = \sum_s \alpha_u(s, T)$$



- In the final observation the alpha at each state gives the probability of all state sequences ending at that state
- The total probability of the observation is the sum of the alpha values at all states



# The forward algorithm

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1. Initialize all alpha terms at  $t=1$ :

$$\alpha(s,1) = \pi(s)P(x_1 | s)$$

2. Recursive estimate alphas for all subsequent time steps

$$\alpha(s,t) = \sum_{s'} \alpha(s',t-1)P(s | s')P(x_t | s)$$

3. Compute overall probability

$$Totalprob = \sum_s \alpha(s,T)$$

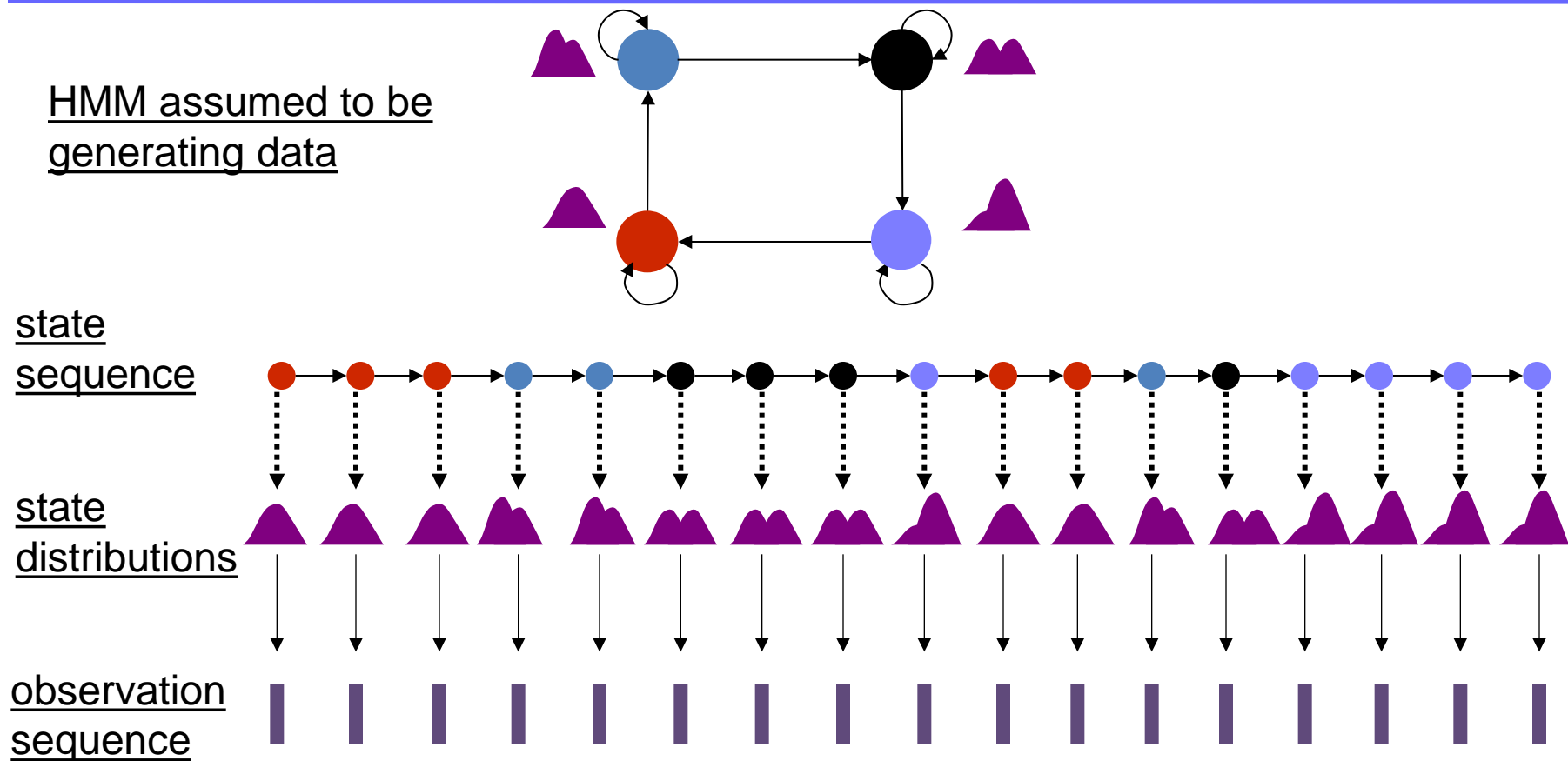
## Problem 2: The state segmentation problem

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- Given only a sequence of observations, how do we determine which sequence of states was followed in producing it?

# The HMM as a generator

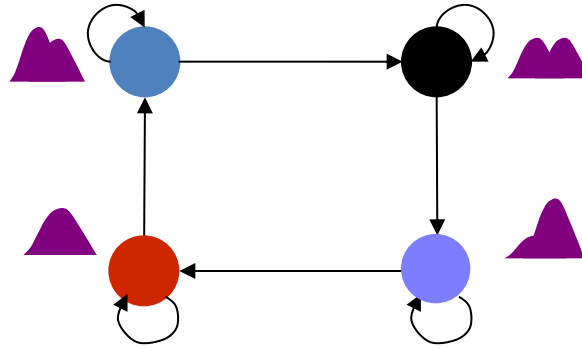
HMM assumed to be  
generating data



- The process goes through a series of states and produces observations from them

# States are Hidden

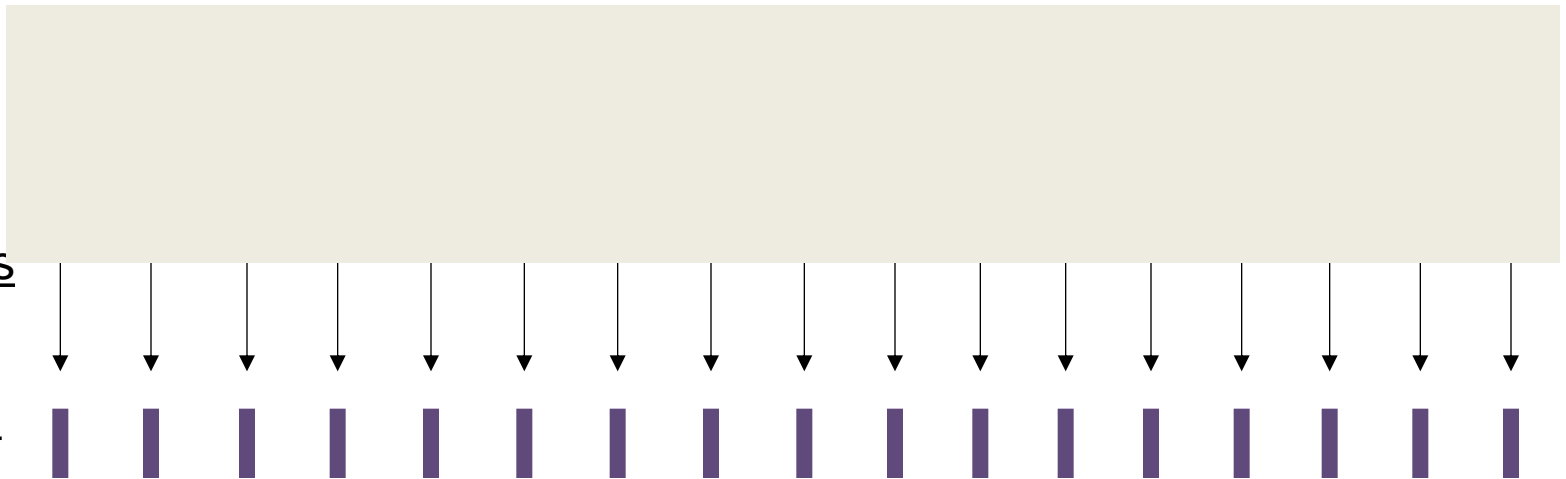
HMM assumed to be  
generating data



state  
sequence

state  
distributions

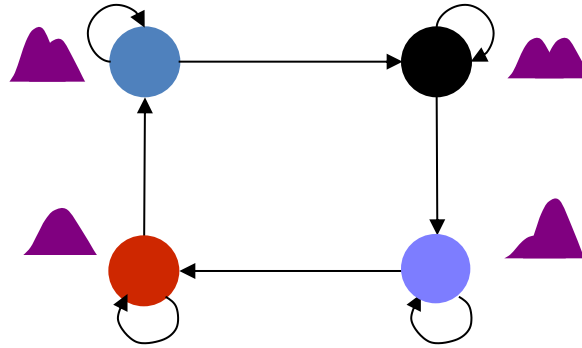
observation  
sequence



- The observations do not reveal the underlying state

# The state segmentation problem

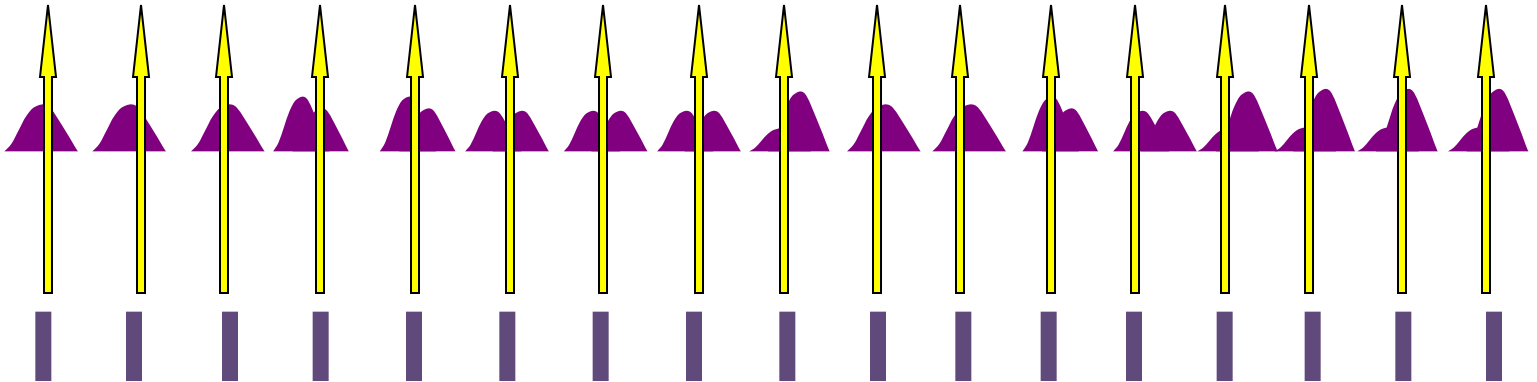
HMM assumed to be  
generating data



state  
sequence



state  
distributions



observation  
sequence

- State segmentation: Estimate state sequence given observations

# Estimating the State Sequence

---

- Any number of state sequences could have been traversed in producing the observation
  - In the worst case *every* state sequence may have produced it
- Solution: Identify the most *probable* state sequence
  - The state sequence for which the probability of progressing through that sequence and generating the observation sequence is maximum
  - i.e.  $P(o_1, o_2, o_3, \dots, s_1, s_2, s_3, \dots)$  is maximum

# Estimating the state sequence

---

- Once again, exhaustive evaluation is impossibly expensive
- But once again a simple dynamic-programming solution is available

$$P(o_1, o_2, o_3, \dots, s_1, s_2, s_3, \dots) =$$

$$P(o_1 | s_1) P(o_2 | s_2) P(o_3 | s_3) \dots P(s_1) P(s_2 | s_1) P(s_3 | s_2) \dots$$

- Needed:

$$\arg \max_{s_1, s_2, s_3, \dots} P(o_1 | s_1) P(s_1) P(o_2 | s_2) P(s_2 | s_1) P(o_3 | s_3) P(s_3 | s_2)$$

# Estimating the state sequence

---

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$$P(o_1, o_2, o_3, \dots, s_1, s_2, s_3, \dots) =$$

$$P(o_1 | s_1) P(o_2 | s_2) P(o_3 | s_3) \dots P(s_1) P(s_2 | s_1) P(s_3 | s_2) \dots$$

- Needed:

$$\arg \max_{s_1, s_2, s_3, \dots} P(o_1 | s_1) P(s_1) P(o_2 | s_2) P(s_2 | s_1) P(o_3 | s_3) P(s_3 | s_2)$$



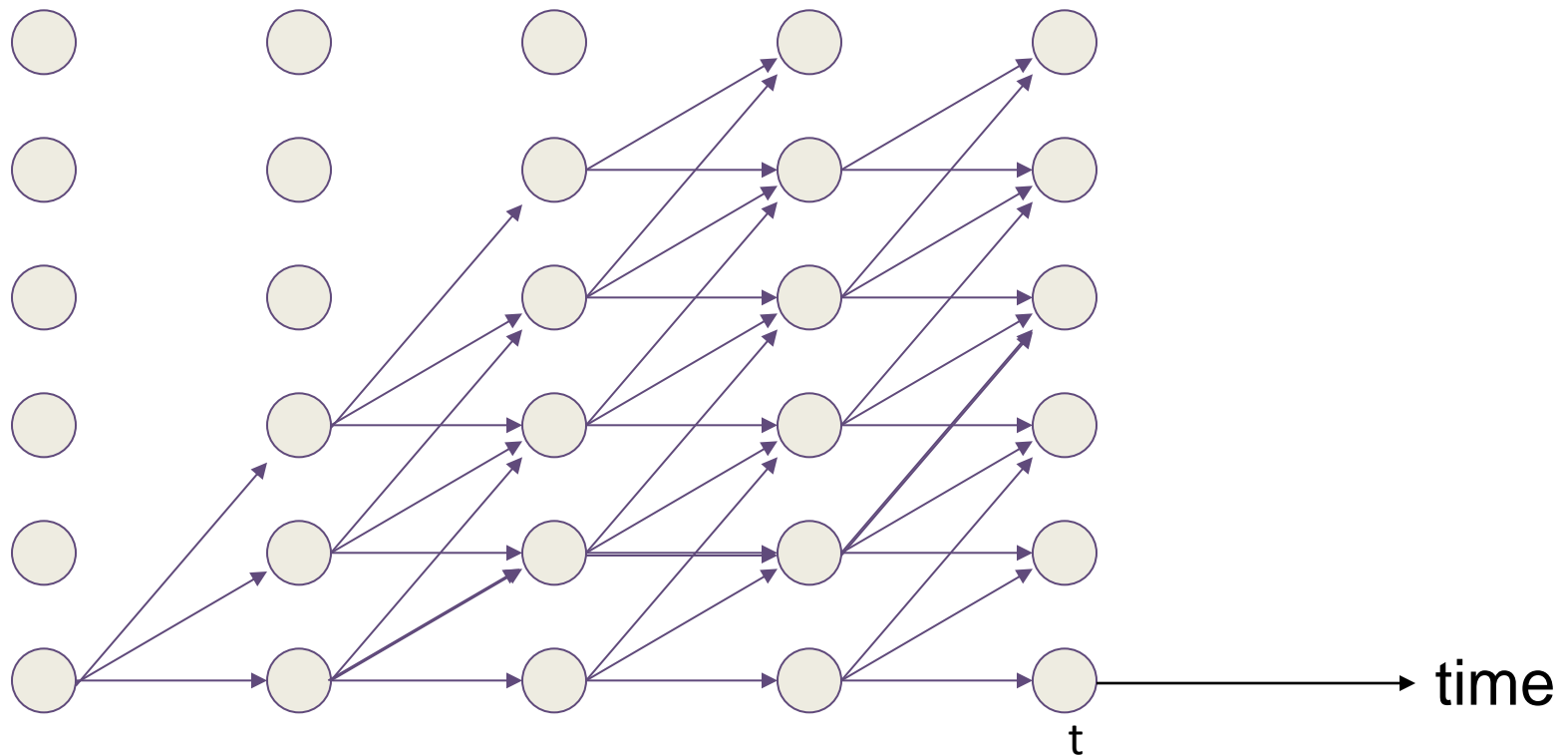
# The state sequence

---

- The probability of a state sequence  $?, ?, ?, ?, s_x, s_y$  ending at time  $t$  is simply
  - $P(?, ?, ?, ?, s_x, s_y) = P(?, ?, ?, ?, s_x) P(o_t | s_y) P(s_y | s_x)$
- The *best* state sequence that ends with  $s_x, s_y$  at  $t$  will have a probability equal to the probability of the best state sequence ending at  $t-1$  at  $s_x$  times  $P(o_t | s_y) P(s_y | s_x)$ 
  - Since the last term is independent of the state sequence leading to  $s_x$  at  $t-1$

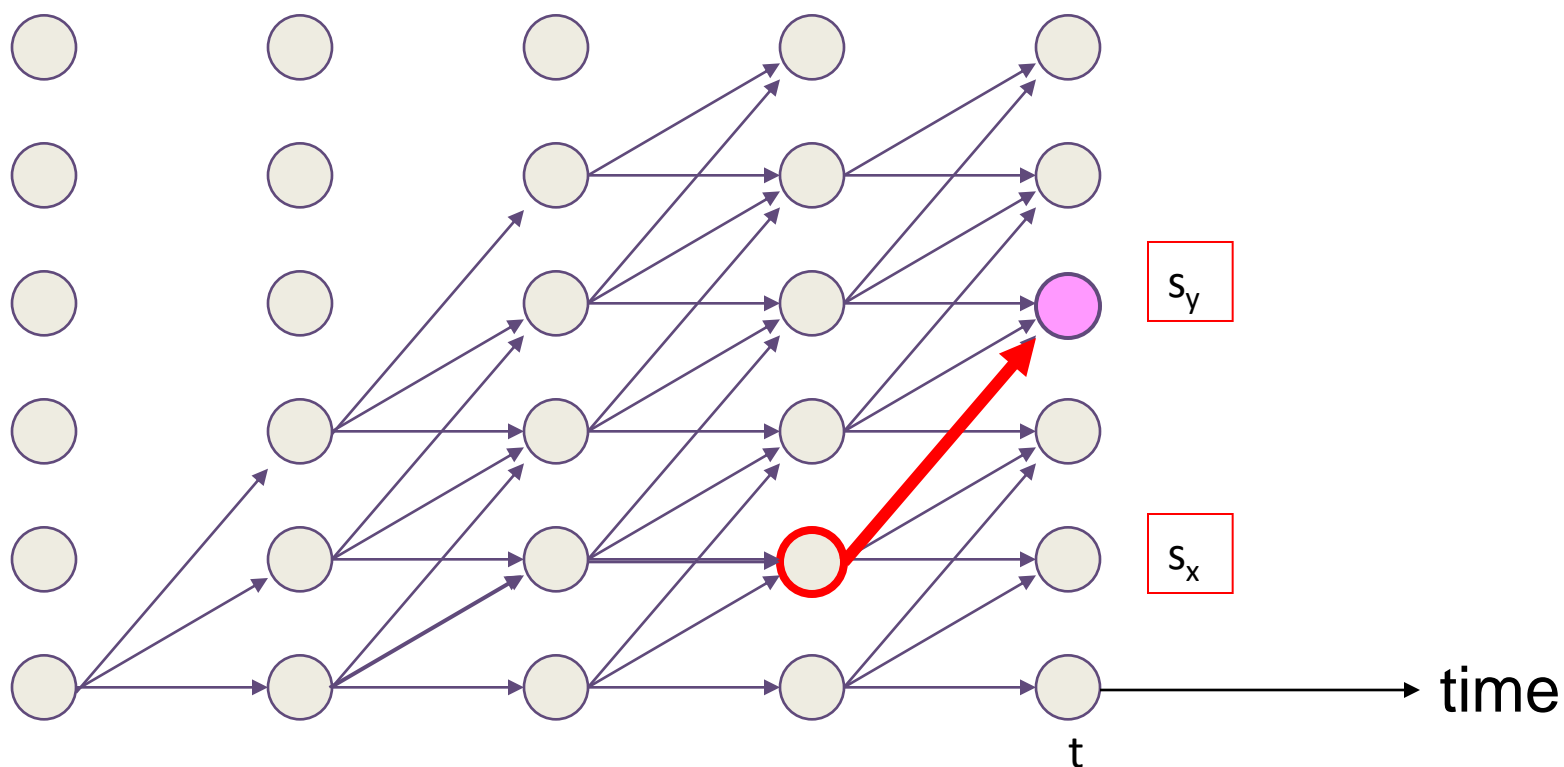
# Trellis

- The graph below shows the set of all possible state sequences through this HMM in five time instants



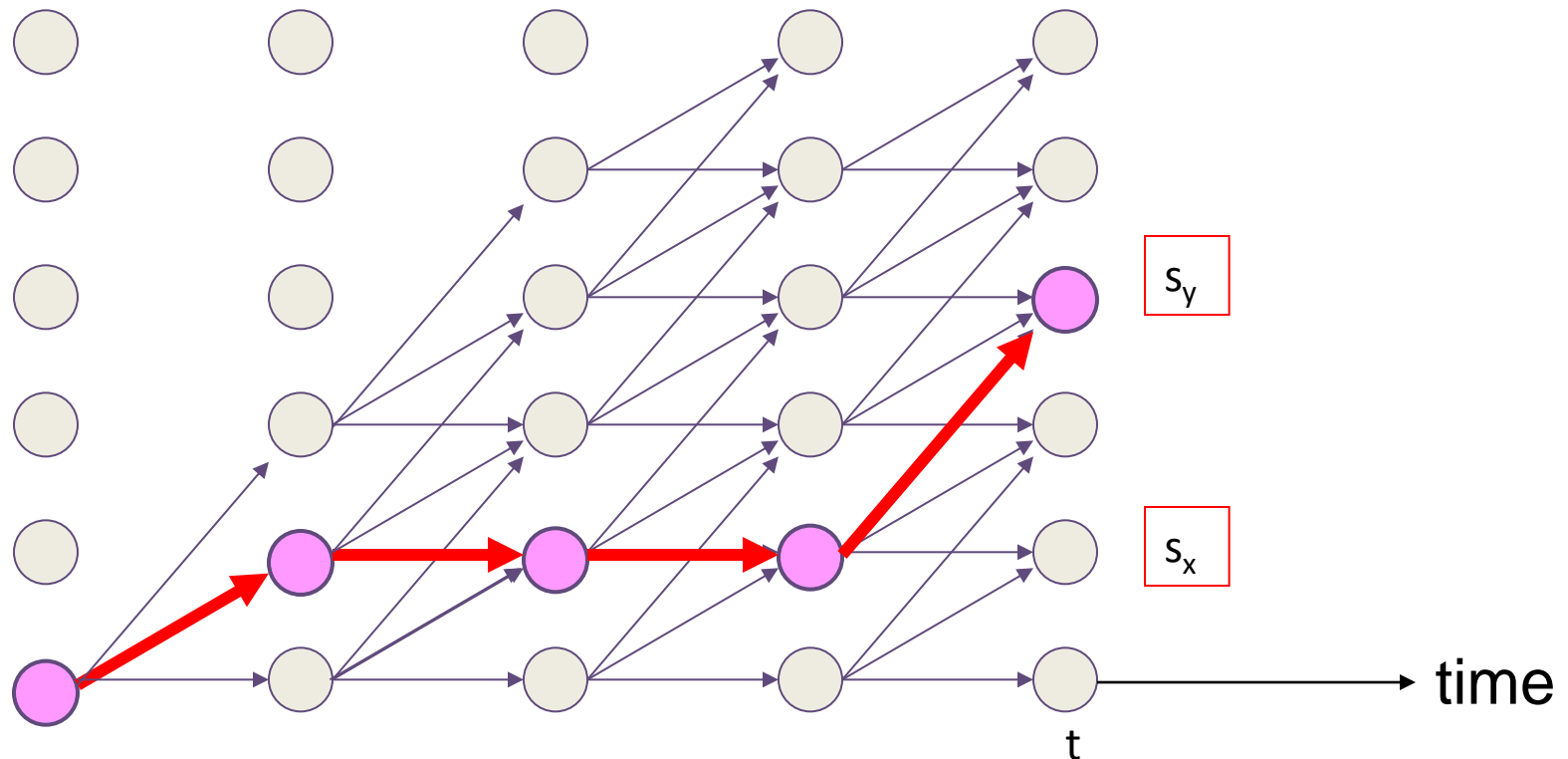
# The cost of extending a state sequence

- The cost of extending a state sequence ending at  $s_x$  is only dependent on the transition from  $s_x$  to  $s_y$ , and the observation probability at  $s_y$



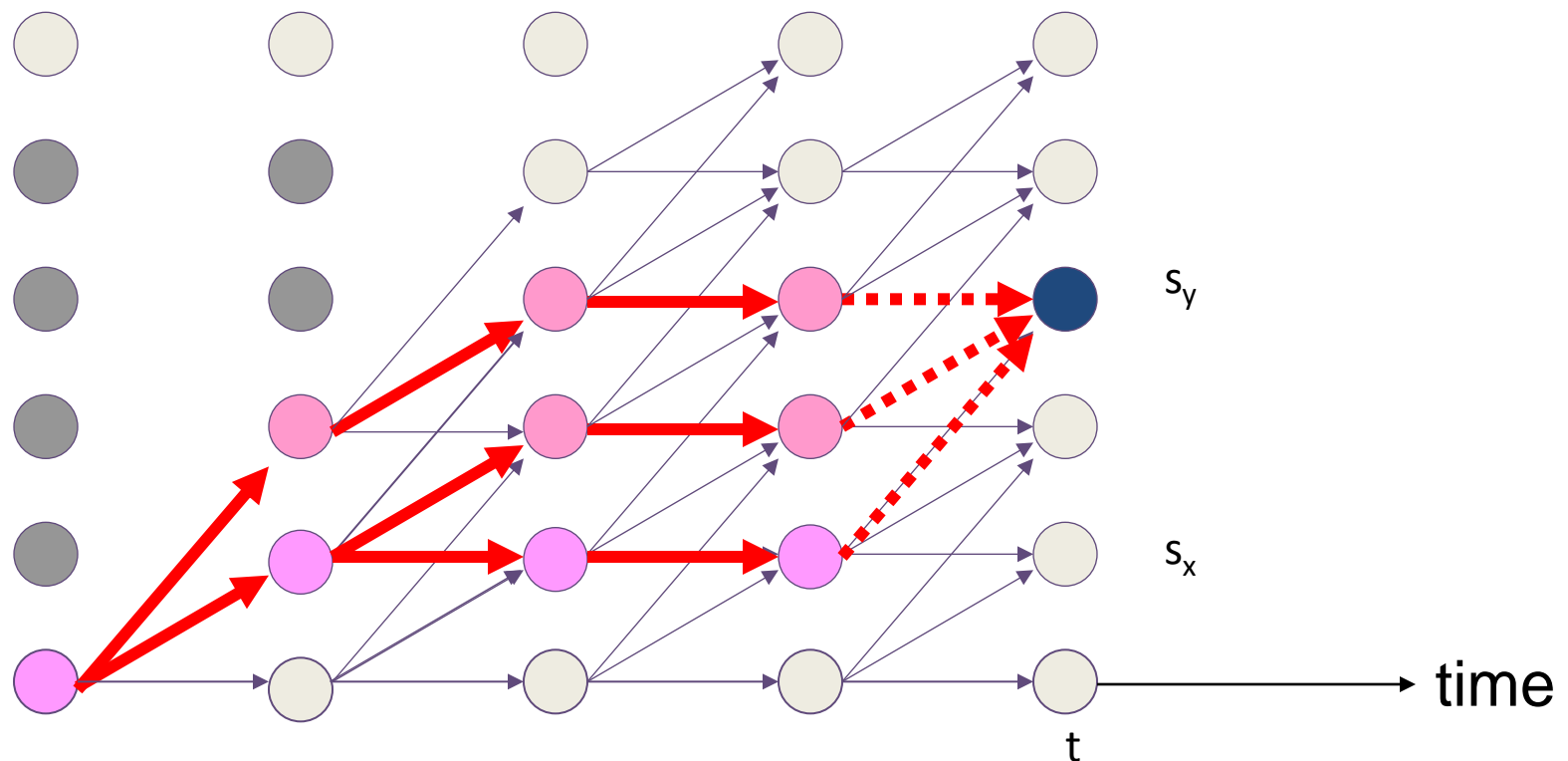
# The cost of extending a state sequence

- The best path to  $s_y$  through  $s_x$  is simply an extension of the best path to  $s_x$



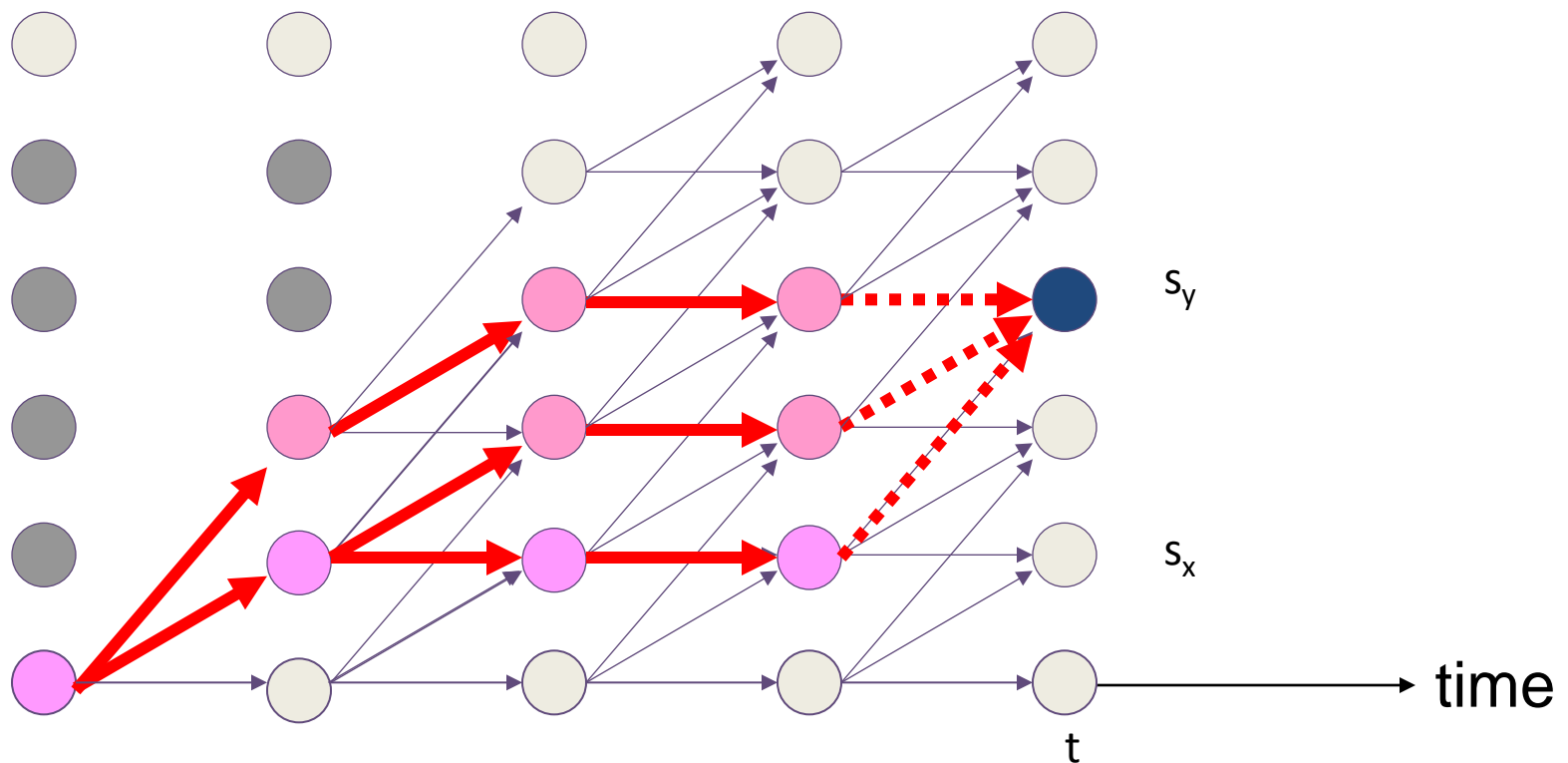
# The Recursion

- The overall best path to  $s_x$  is an extension of the best path to one of the states at the previous time



# The Recursion

- Bestpath  $\text{prob}(s_y, t) =$   
 $\text{Best} (\text{Bestpath prob}(s_{?}, t) * P(s_y | s_{?}) * P(o_t | s_y))$



# Finding the best state sequence

---

- This gives us a simple recursive formulation to find the overall best state sequence:
  1. The best state sequence  $X_{1,i}$  of length 1 ending at state  $s_i$  is simply  $s_i$ .
    - The probability  $C(X_{1,i})$  of  $X_{1,i}$  is  $P(o_1 | s_i) P(s_i)$
  2. The best state sequence of length  $t+1$  is simply given by
    - $(\operatorname{argmax}_{X_{t,i}} C(X_{t,i}) P(o_{t+1} | s_j) P(s_j | s_i)) s_i$
  3. The best overall state sequence for an utterance of length  $T$  is given by
    - $\operatorname{argmax}_{X_{t,i} s_j} C(X_{T,i})$
    - The state sequence of length  $T$  with the highest overall probability

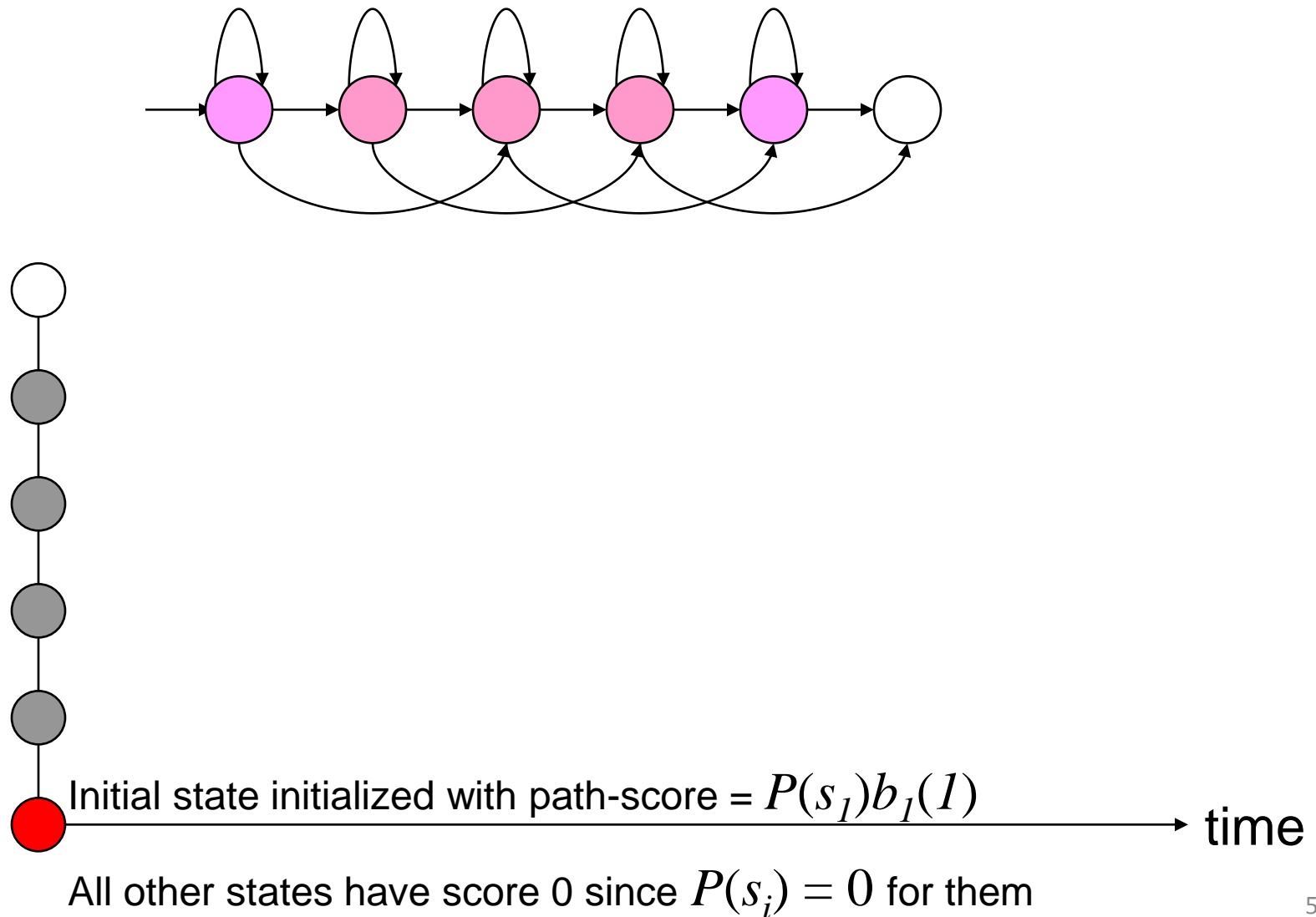
# Finding the best state sequence

---

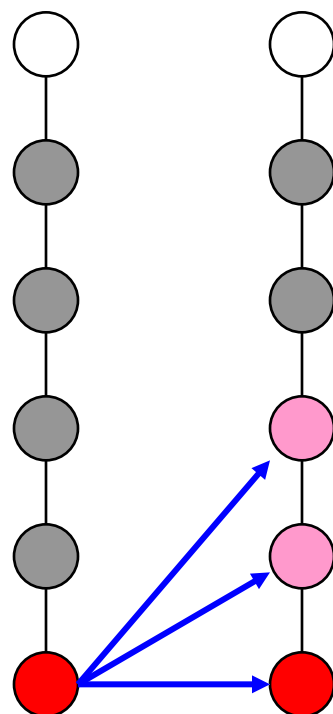
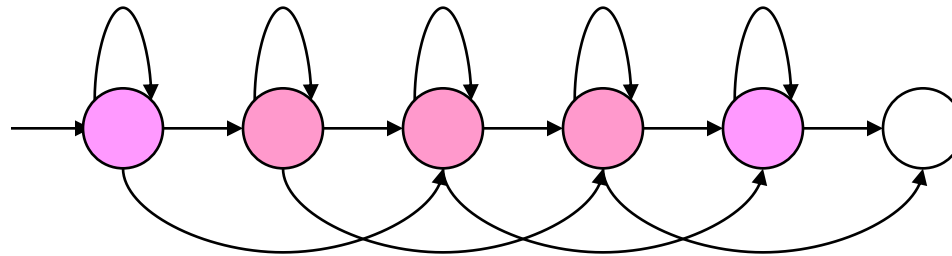
- The simple algorithm just presented is called the VITERBI algorithm in the literature
  - After A.J.Viterbi, who invented this dynamic programming algorithm for a completely different purpose: decoding error correction codes!
- The Viterbi algorithm can also be viewed as a breadth-first graph search algorithm
  - The HMM forms the Y axis of a 2-D plane
    - Edge costs of this graph are transition probabilities  $P(s|s)$ . Node costs are  $P(o|s)$
  - A linear graph with every node at a time step forms the X axis
  - A trellis is a graph formed as the crossproduct of these two graphs
  - The Viterbi algorithm finds the best path through this graph



# Viterbi Search (contd.)



# Viterbi Search (contd.)



- State with best path-score
- State with path-score < best
- State without a valid path-score

$$P_j(t) = \max_i [P_i(t-1) a_{ij} b_j(t)]$$

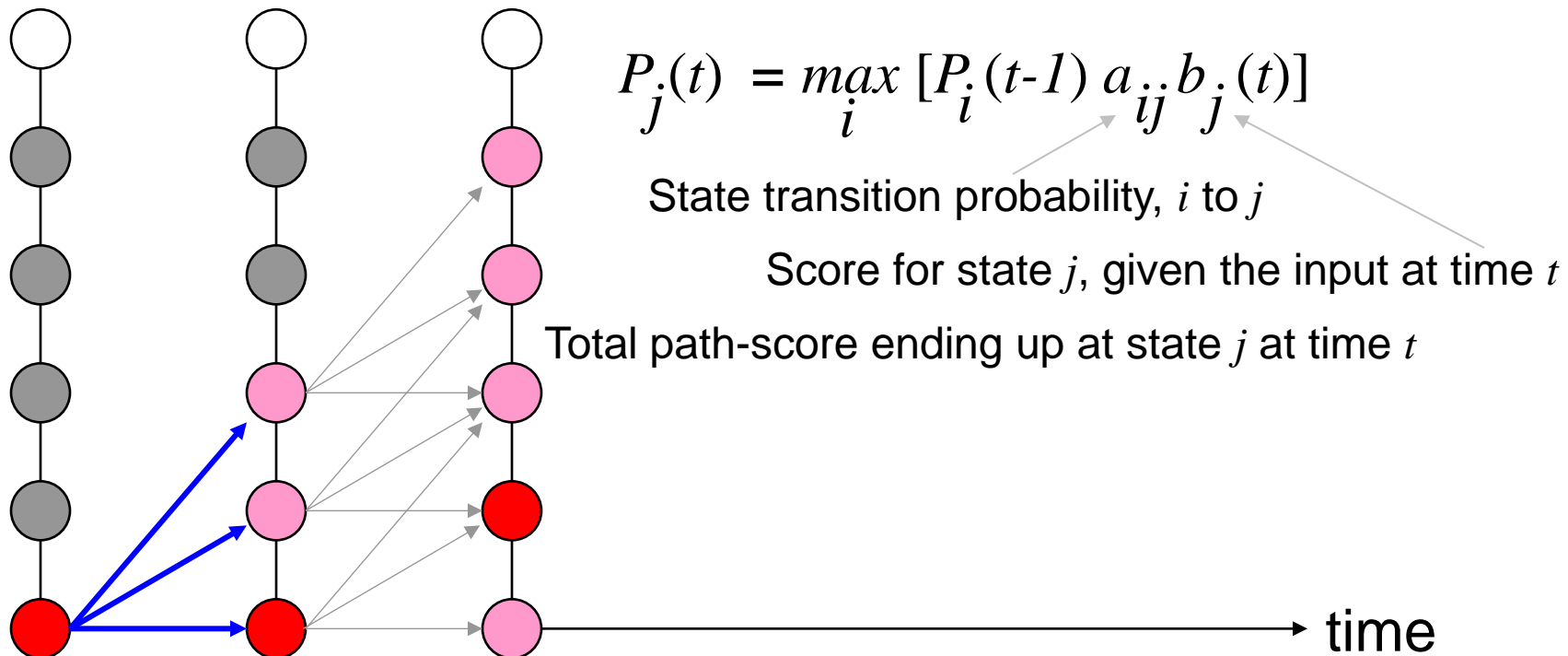
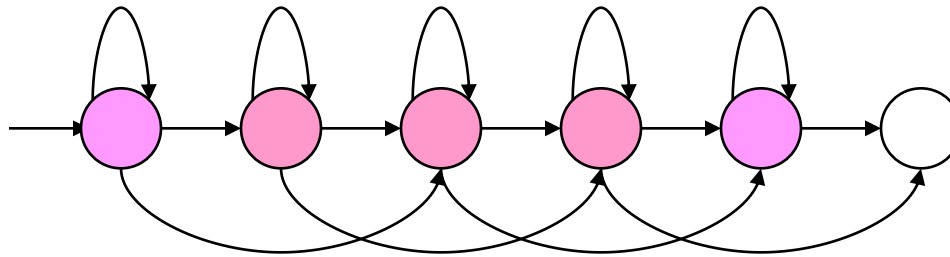
State transition probability,  $i$  to  $j$

Score for state  $j$ , given the input at time  $t$

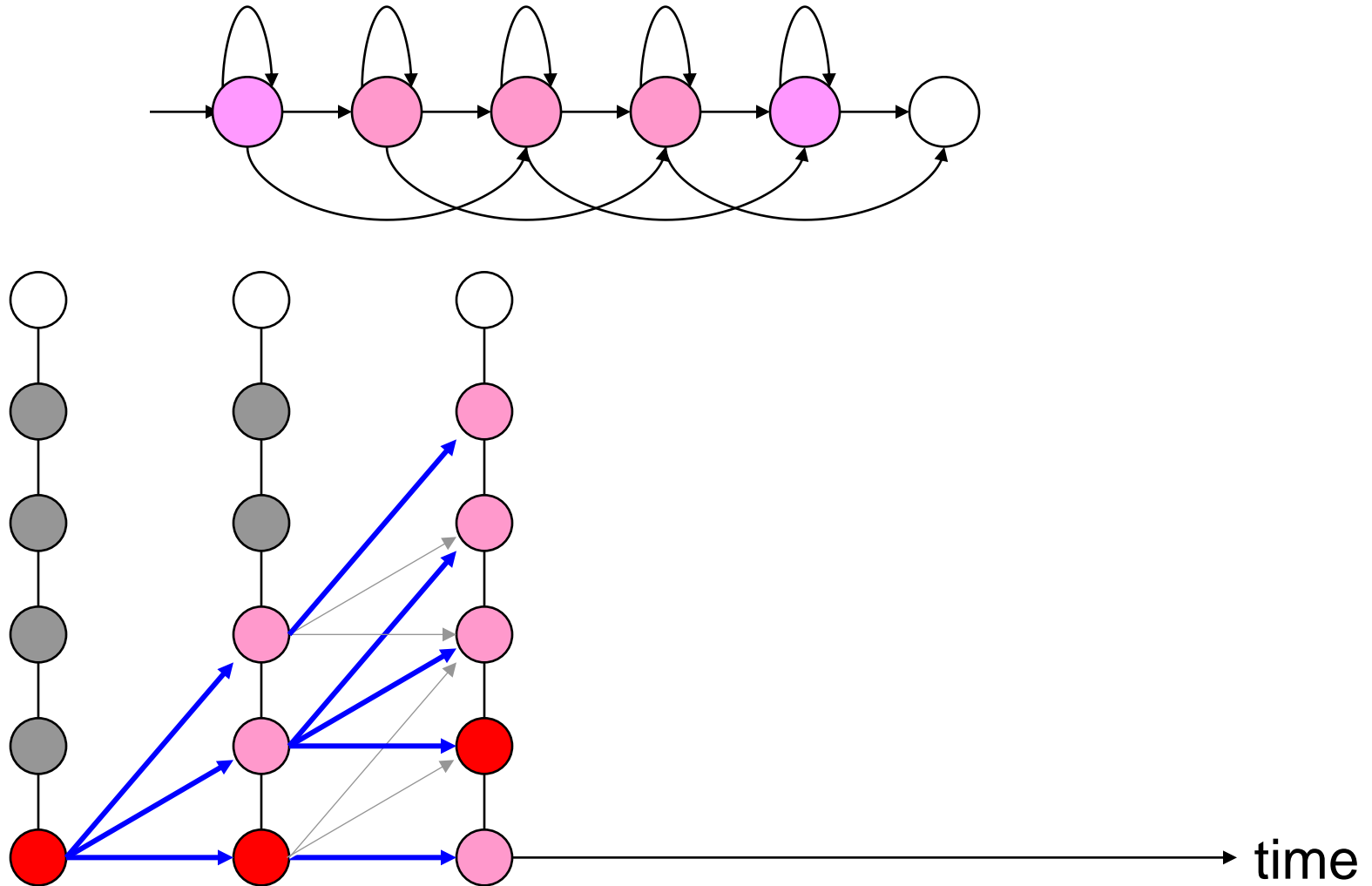
Total path-score ending up at state  $j$  at time  $t$

time

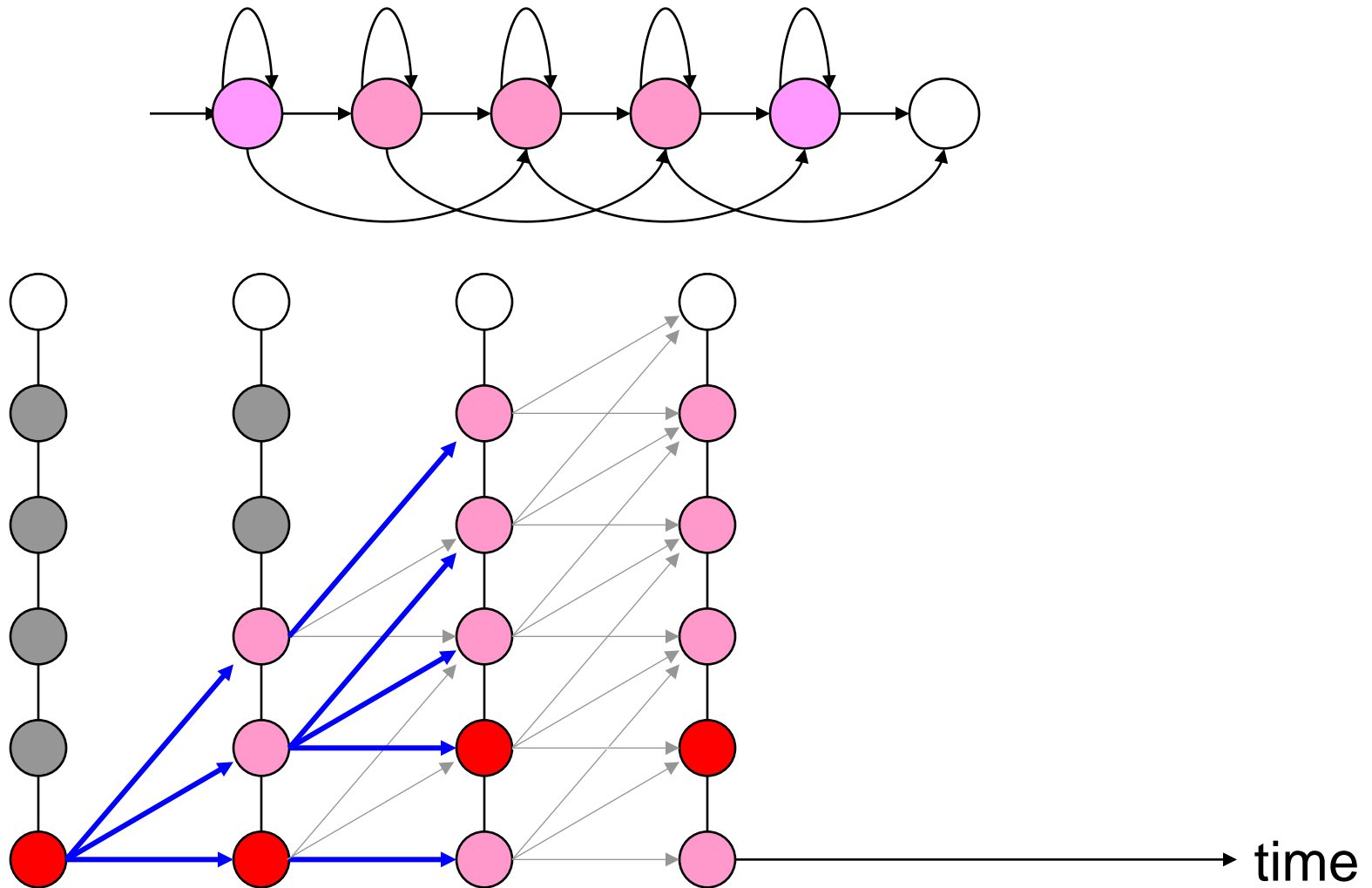
# Viterbi Search (contd.)



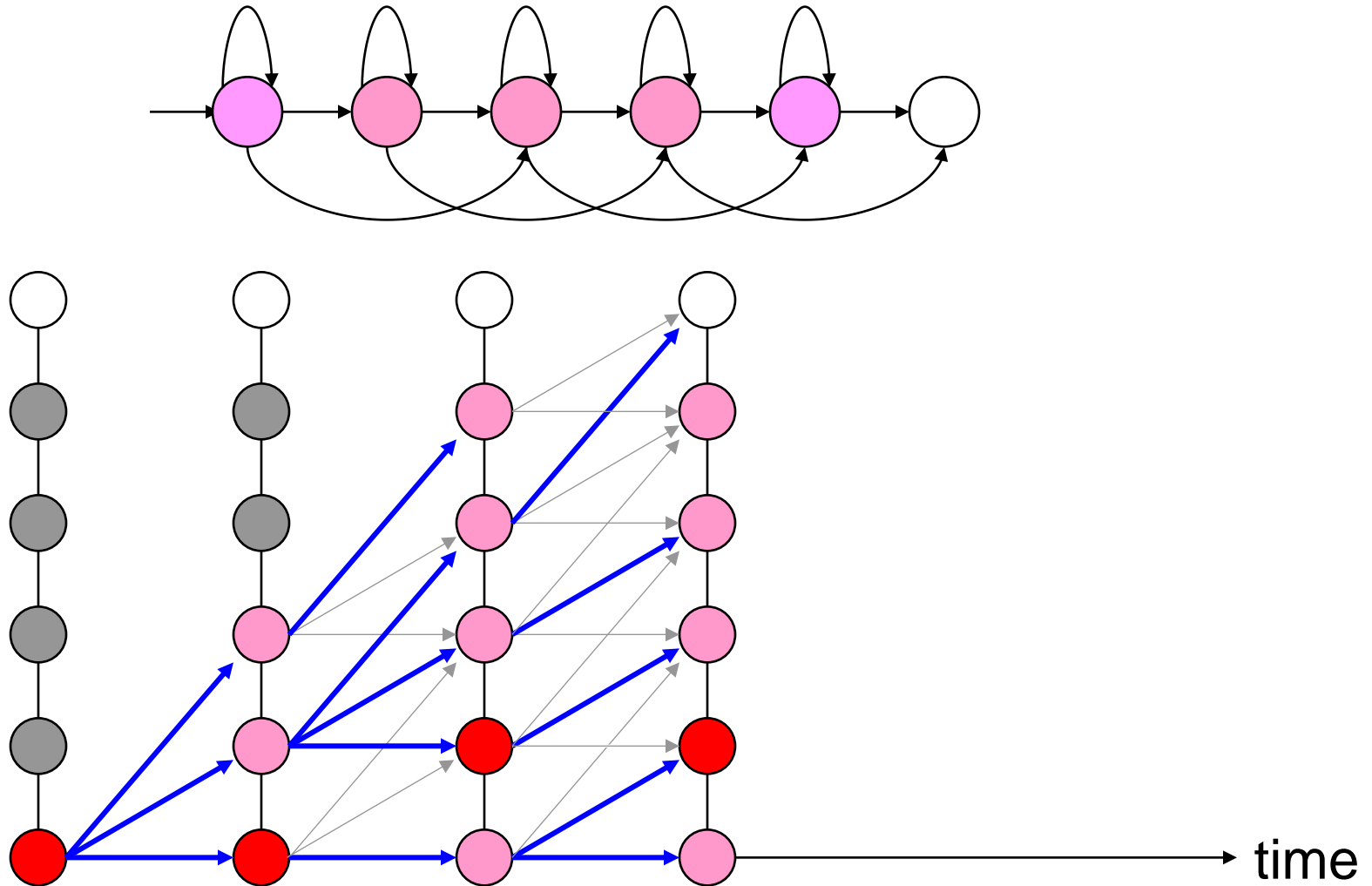
# Viterbi Search (contd.)



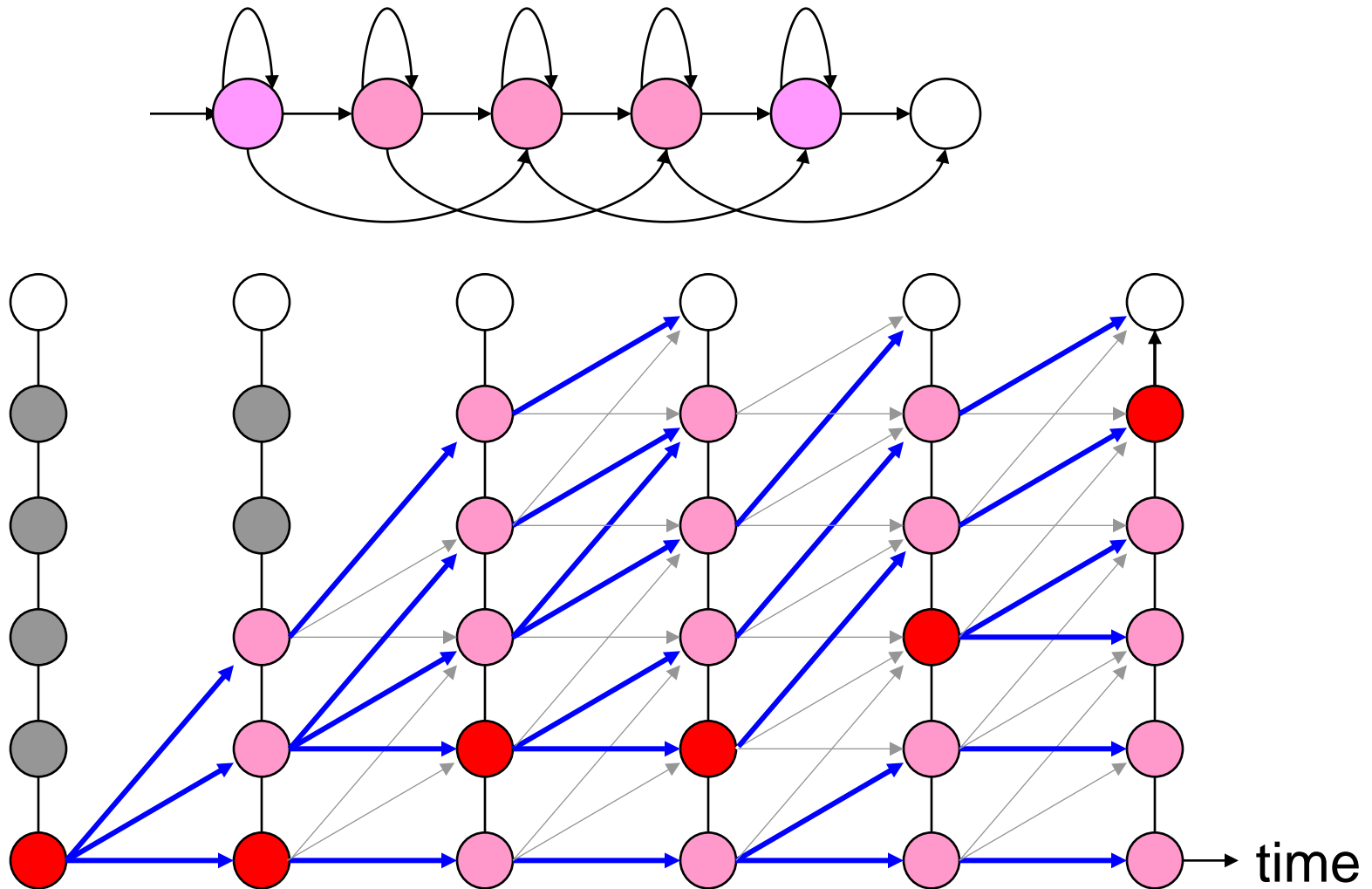
# Viterbi Search (contd.)



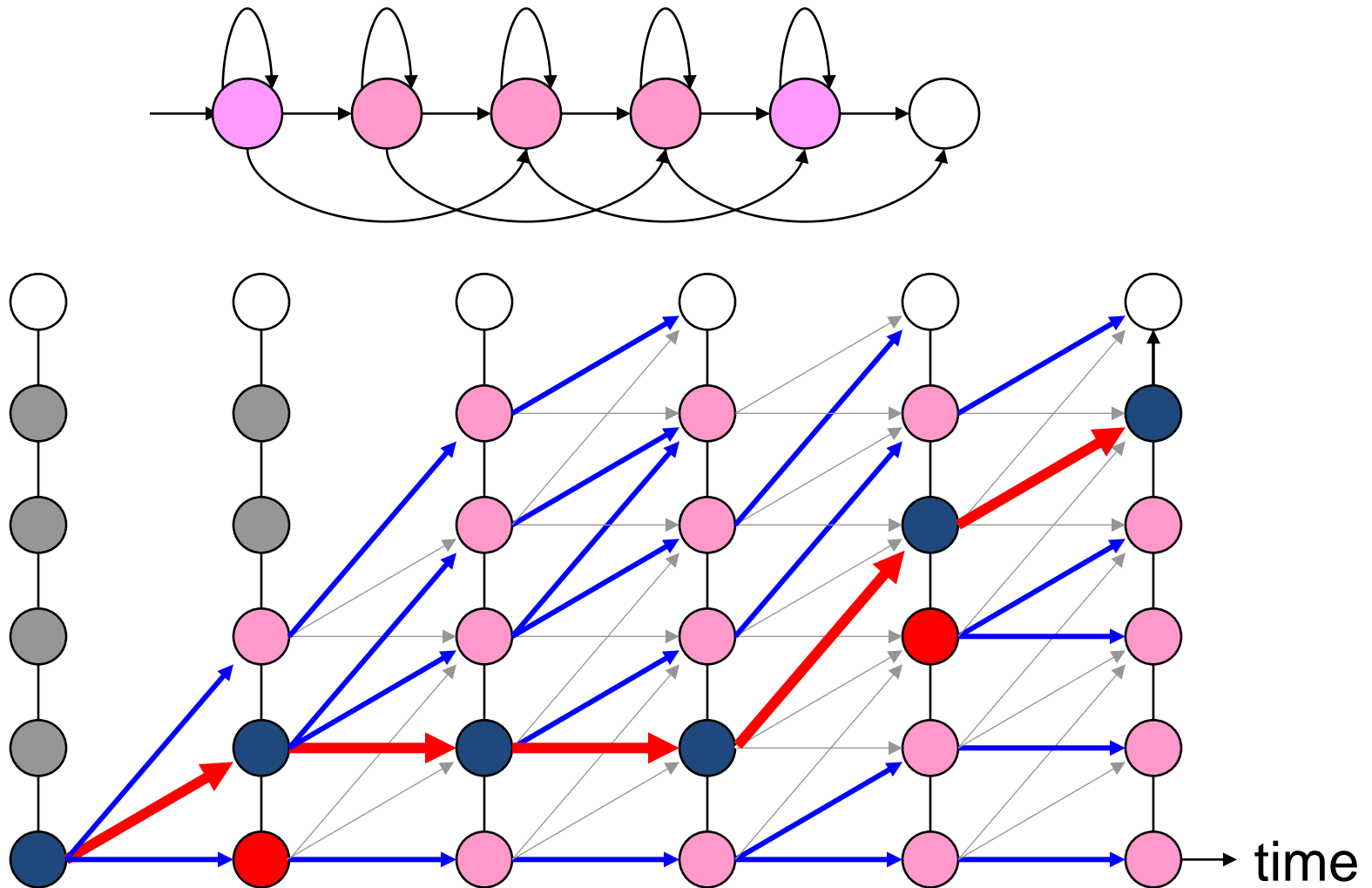
# Viterbi Search (contd.)



# Viterbi Search (contd.)

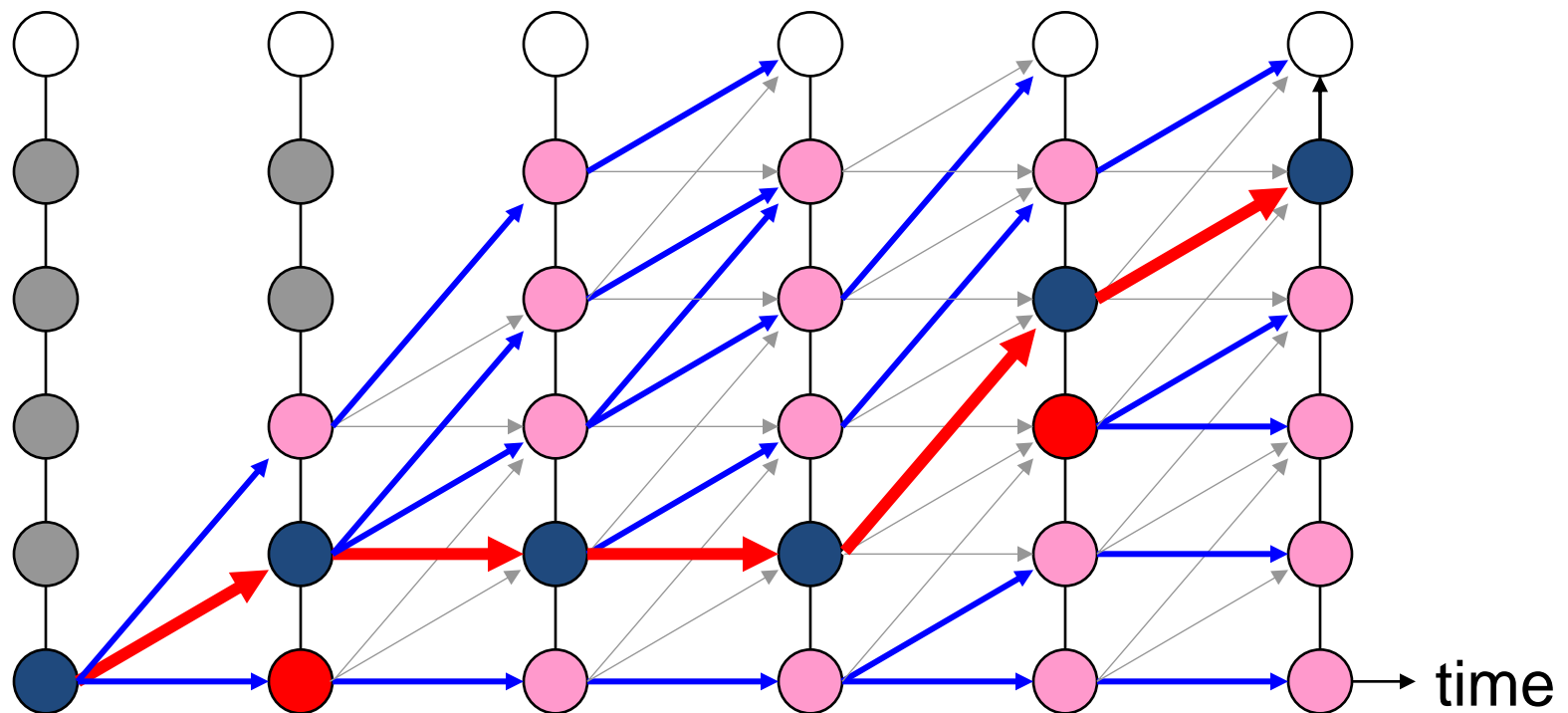


# Viterbi Search (contd.)





THE BEST STATE SEQUENCE IS THE ESTIMATE OF THE STATE SEQUENCE FOLLOWED IN GENERATING THE OBSERVATION



# Viterbi and DTW

---

- The Viterbi algorithm is identical to the string-matching procedure used for DTW that we saw earlier
- It computes an estimate of the state sequence followed in producing the observation
- *It also gives us the probability of the best state sequence*

# Problem3: Training HMM parameters

---

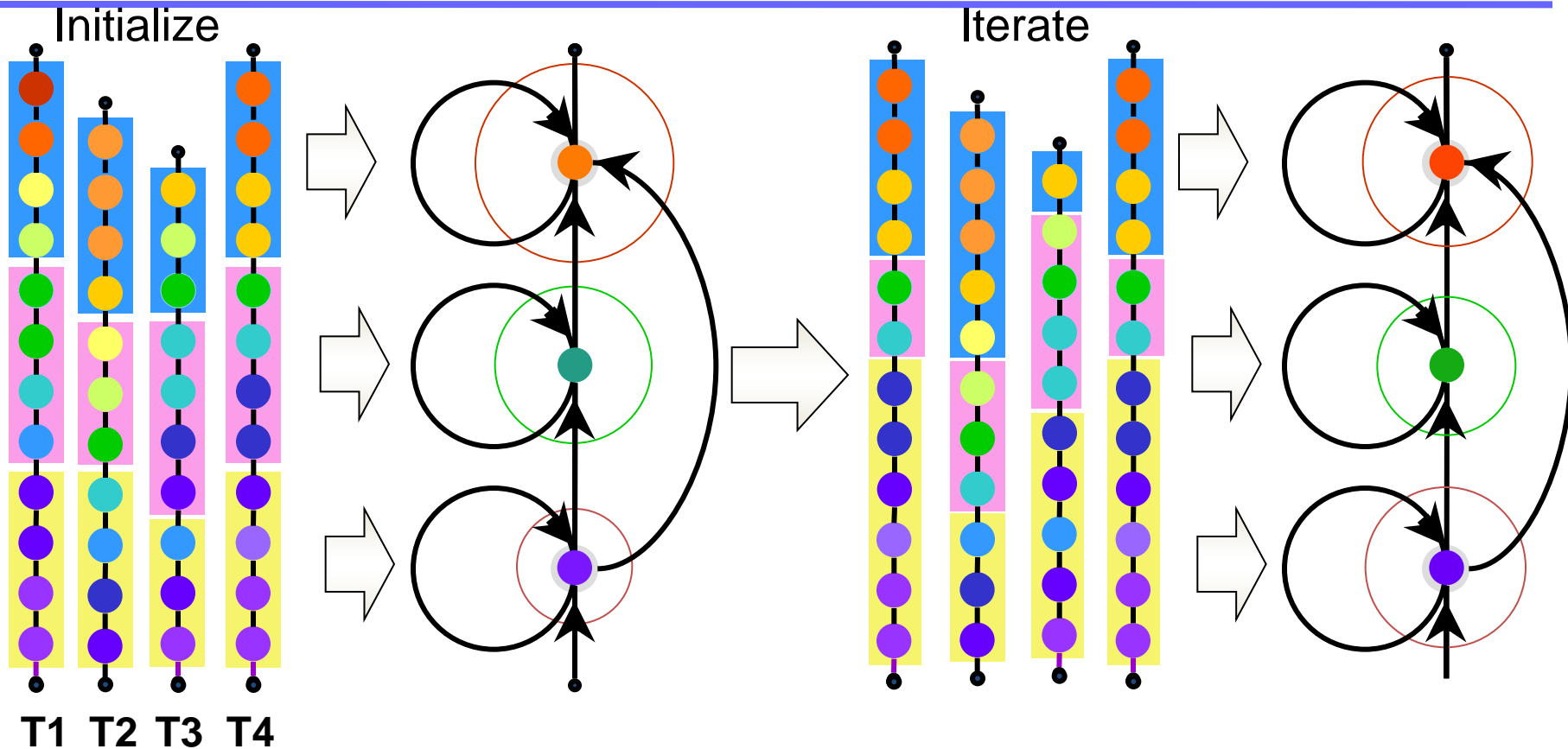
- We can compute the probability of an observation, and the best state sequence given an observation, using the HMM's parameters
- But where do the HMM parameters come from?
- They must be learned from a collection of observation sequences
- We have already seen one technique for training HMMs: The segmental K-means procedure

# Modified segmental K-means AKA Viterbi training

---

- The entire segmental K-means algorithm:
  1. Initialize all parameters
    - State means and covariances
    - Transition probabilities
    - Initial state probabilities
  2. Segment all training sequences
  3. Reestimate parameters from segmented training sequences
  4. If not converged, return to 2

# Segmental K-means

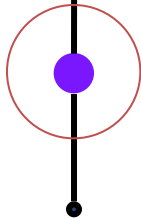
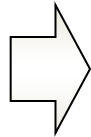
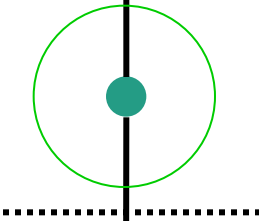
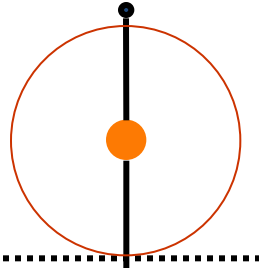
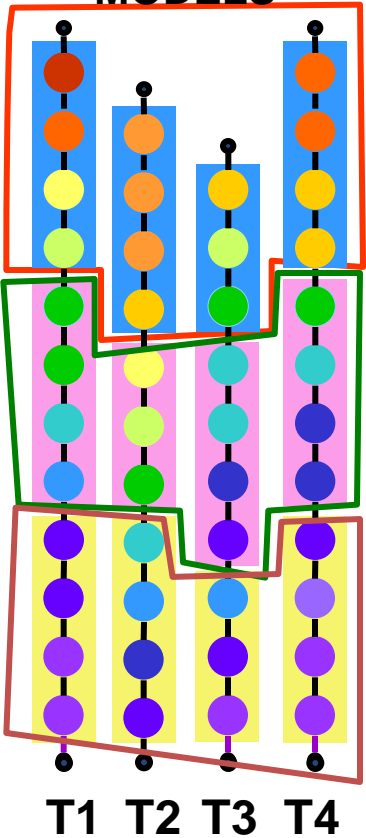


The procedure can be continued until convergence

Convergence is achieved when the total best-alignment error for all training sequences does not change significantly with further refinement of the model

# Training by segmentation: Hard Assignment

## MODELS



$$m_j = \frac{1}{\sum_{i \in \text{segment}(j)} 1} \sum_{i \in \text{segment}(j)} x_i$$

$$C_j = \frac{1}{\sum_{i \in \text{segment}(j)} 1} \sum_{i \in \text{segment}(j)} (x_i - m_j)(x_i - m_j)^T$$

Each vector belongs uniquely to a segment

$$P(x | j) = \frac{1}{\sqrt{2\pi |C|}} \exp(-0.5(x - m_j)^T C^{-1}(x - m_j))$$

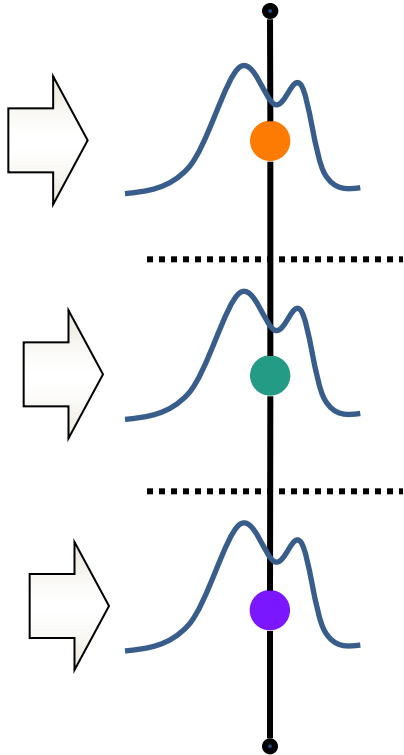
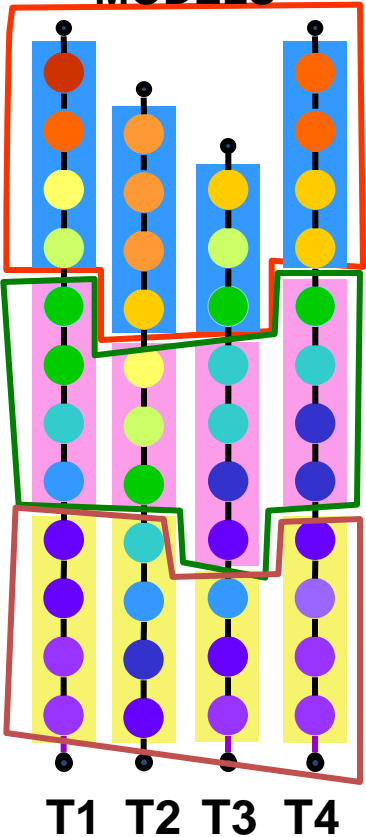
Assumes state output distribution is Gaussian

$$d_j(v) = -\log(P(x | j))$$

$$P(x | j) = \frac{1}{\sqrt{2\pi |C|}} \exp\left(-\sum_d \frac{(x_d - m_{j,d})^2}{2\sigma_d^2}\right)$$

# Training by segmentation: Hard Assignment with Gaussian Mixtures

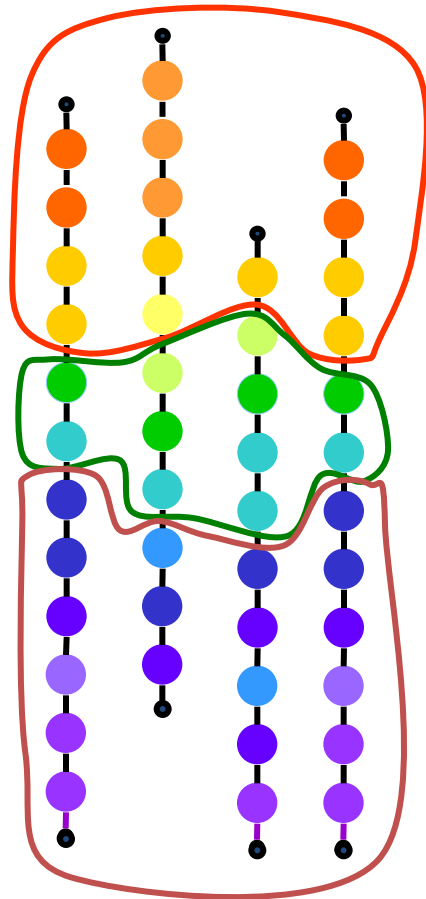
MODELS



Dealing with Gaussian Mixtures at states?

The distribution of vectors in any state is assumed to be a Gaussian mixture

# Training by segmentation: Hard Assignment with Gaussian Mixtures



Assume the distribution of each collection of vectors is a Gaussian mixture

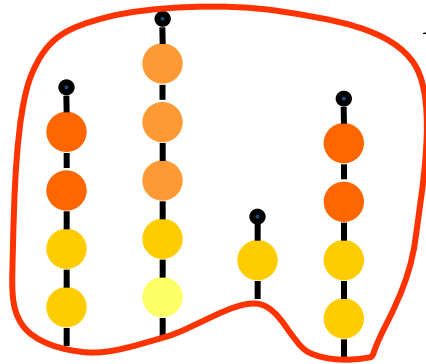
$$P(x | j) = \sum_k \frac{w_k}{\sqrt{\prod_l 2\pi\sigma_{j,k,l}^2}} \exp\left(-0.5 \sum_l \frac{(x_l - m_{j,k,l})^2}{\sigma_{j,k,l}^2}\right)$$

$$d_j(v) = -\log(P(x | j))$$

Above equation assumes Gaussian covariance matrices are diagonal



# Training a GMM by hard counting



How does one learn the parameters of the GMM?

$$P(x | j) = \sum_k \frac{w_k}{\sqrt{\prod_l 2\pi\sigma_{j,k,l}^2}} \exp\left(-0.5 \sum_l \frac{(x_l - m_{j,k,l})^2}{\sigma_{j,k,l}^2}\right)$$

$$d_j(v) = -\log(P(x | j))$$

# Gaussian Mixtures

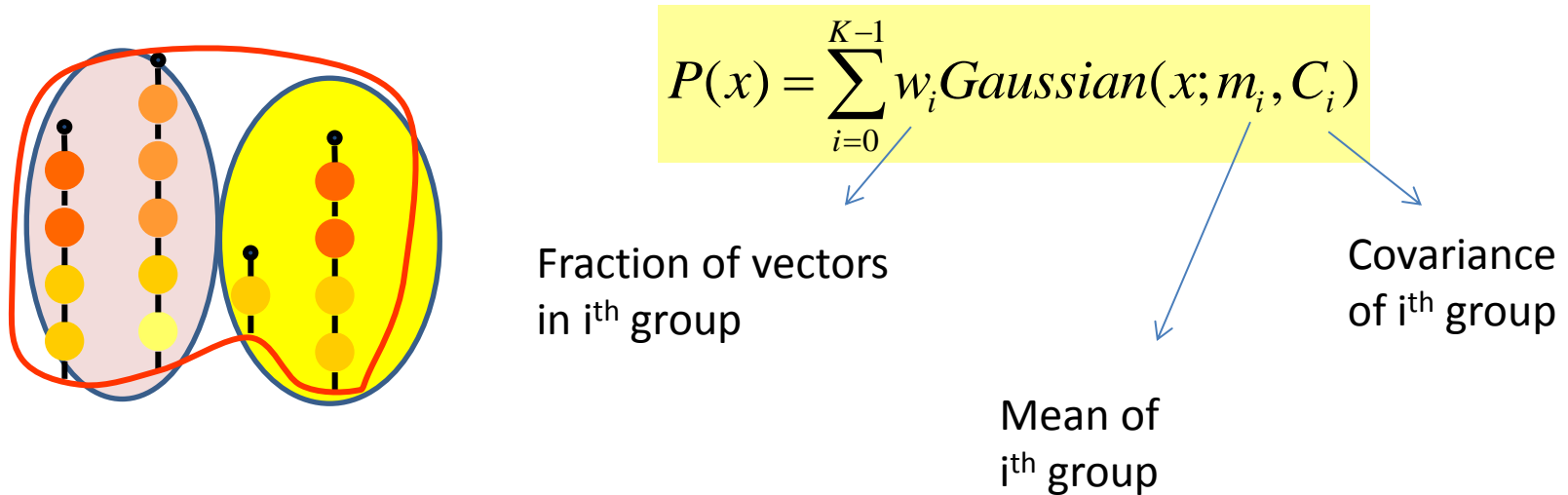
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- A Gaussian Mixture is literally a mixture of Gaussians. It is a weighted combination of several Gaussian distributions

$$P(x) = \sum_{i=0}^{K-1} w_i \text{Gaussian}(x; m_i, C_i)$$

- $x$  is any data vector.  $P(x)$  is the probability given to that vector by the Gaussian mixture
- $K$  is the number of Gaussians being mixed
- $w_i$  is the mixture weight of the  $i^{\text{th}}$  Gaussian.  $m_i$  is its mean and  $C_i$  is its covariance

# Gaussian Mixtures: A “hard” perspective

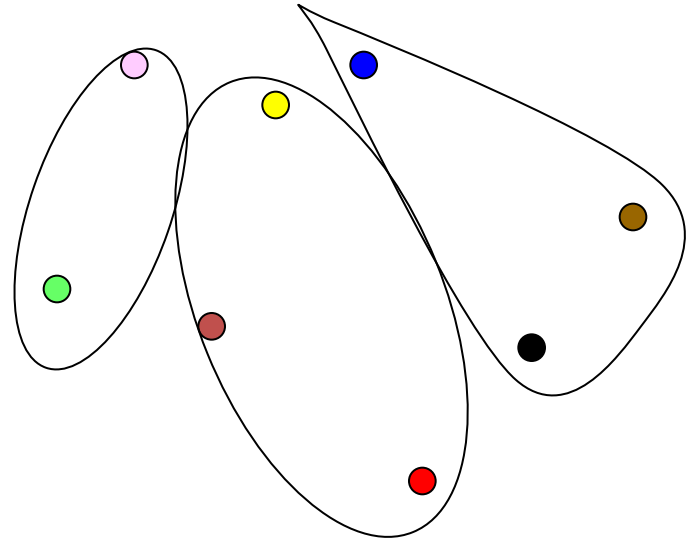


- Data from each Gaussian can be clearly grouped into clusters
- The parameters of the GMM are the parameters of individual clusters
- We can use a clustering algorithm to find the clusters
  - K-means

# The K-means algorithm

---

- The K-means algorithm is an iterative algorithm for clustering similar data from a data set
  - Similarity defined in terms of distance between clusters and data
    - E.g. distance from cluster mean
    - Negative log probability of the vector given by the distribution of the cluster
- The algorithm tries to find the most consistent clusters of data
  - Consistency in terms of specified distance measure



# K-Means training Gaussian Mixtures

- The K-means algorithm can be used to estimate Gaussian mixture distributions for a data set
- Each of the K Gaussians is assumed to represent a separate cluster of the data
- The  $j^{\text{th}}$  cluster is characterized by
  - Its covariance  $C_j$
  - Its mean vector  $m_j$
  - A mixture weight  $w_j$  that specifies what portion of the total data belongs to that cluster
- Define the distance between a vector and the  $j^{\text{th}}$  cluster as

$$d(v, j) = 0.5 \log \left( (2\pi)^D |C_j| \right) + 0.5 (v - m_j)^T C_j^{-1} (v - m_j) - \log(w_j)$$

$-\log P(v, j) = -\log P(j) - \log P(v|j)$ ,  $P()$  is a Gaussian

D = dimension of vectors

# K-Means: Estimating parameters for a cluster

- The parameters for a cluster are its mixture weight, mean vector and covariance matrix. These are computed as follows:

$$m_j = \frac{1}{N_j} \sum_{v: j(v)=j} v$$

$j(v)$  is the cluster that vector  $v$  is assigned to

- $N_j$  is the number of vectors that have been tagged as belonging to cluster  $j$
- The summation is over all vectors who have been tagged as belonging to  $j$

$$C_j = \frac{1}{N_j} \sum_{v: j(v)=j} (v - m_j)(v - m_j)^T$$

$$w_j = \frac{N_j}{N}$$

- $N$  is the total number of training vectors for all clusters

# The K-means algorithm

---

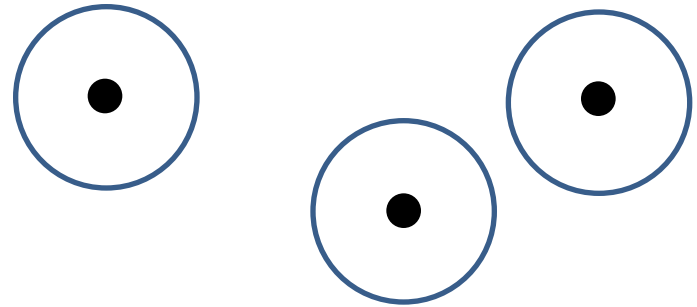
- Initialize all clusters somehow (the number of clusters is assumed)
- For each training vector, find the closest cluster
- Reassign training vectors to their closest clusters
- Iterate the above two steps until the total distance of all training vectors from their clusters converges
  - Convergence can be proved for most distance measures

# K-means

---

## 1. Initialize cluster parameters

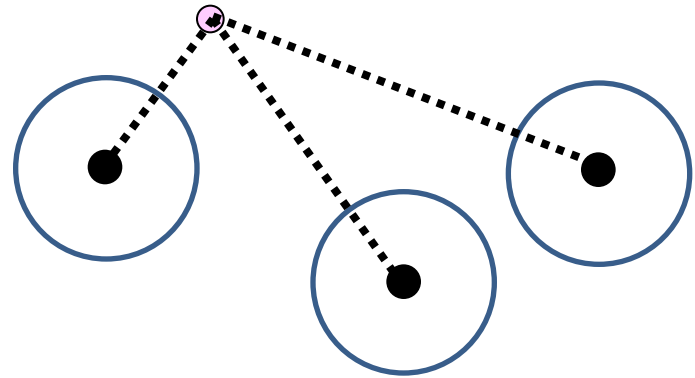
How? We'll return to this shortly





# K-means

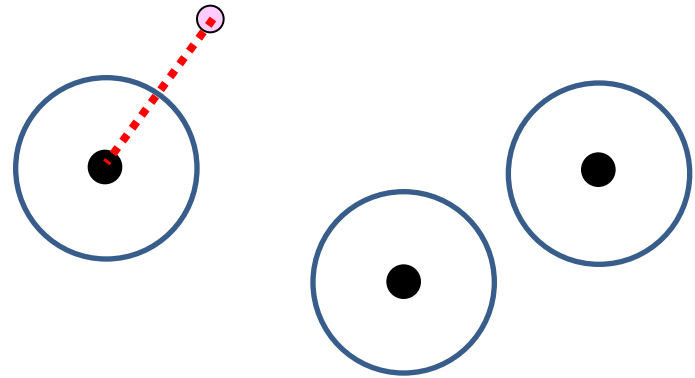
1. Initialize cluster parameters
2. For each data point  $x$ , find the distance from each cluster
  - $d_{cluster} = d(x, cluster)$



$$d(v, j) = 0.5 \log((2\pi)^D |C_j|) + 0.5(v - m_j)^T C_j^{-1} (v - m_j) - \log(w_i)$$

# K-means

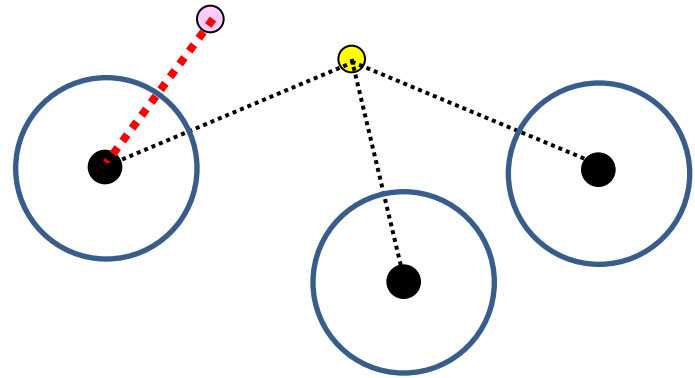
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3. Put data point in the cluster of the closest centroid
  - Cluster for which  $d_{cluster}$  is minimum



$$d(v, j) = 0.5 \log \left( (2\pi)^D |C_j| \right) + 0.5 (v - m_j)^T C_j^{-1} (v - m_j) - \log(w_i)$$

# K-means

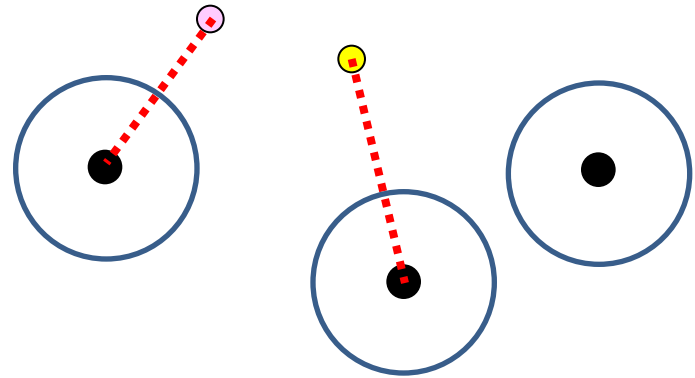
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# K-means

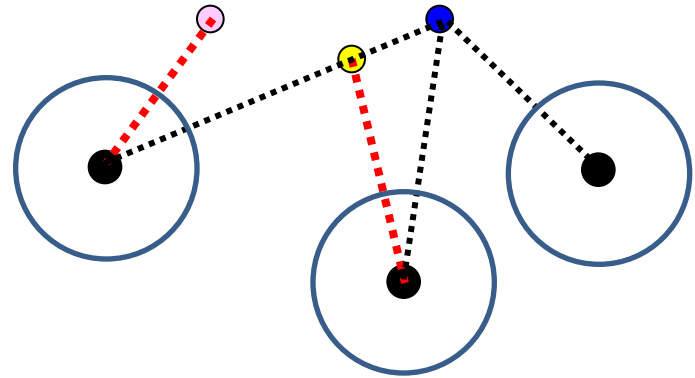
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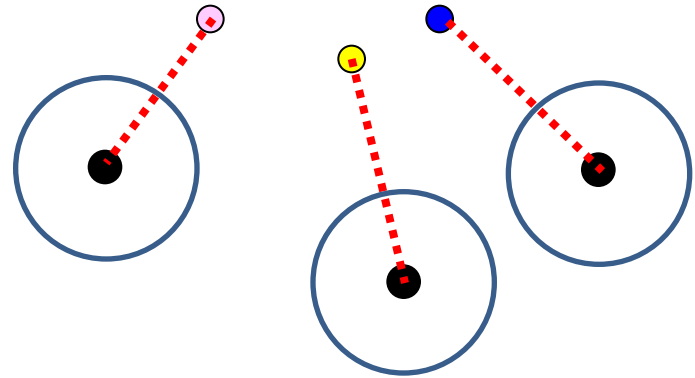
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# K-means

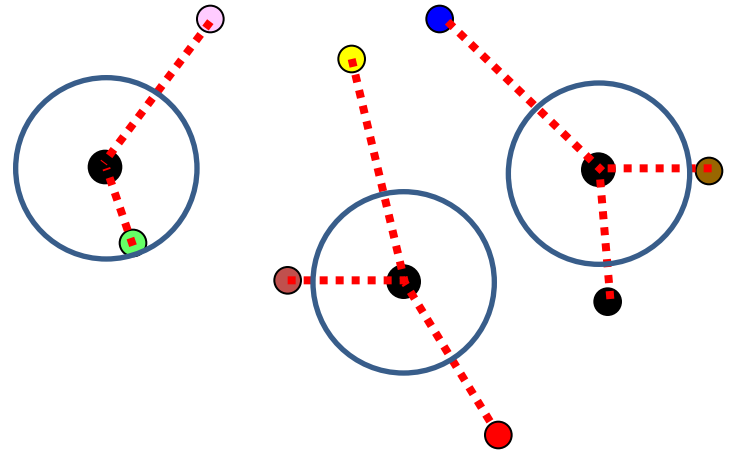
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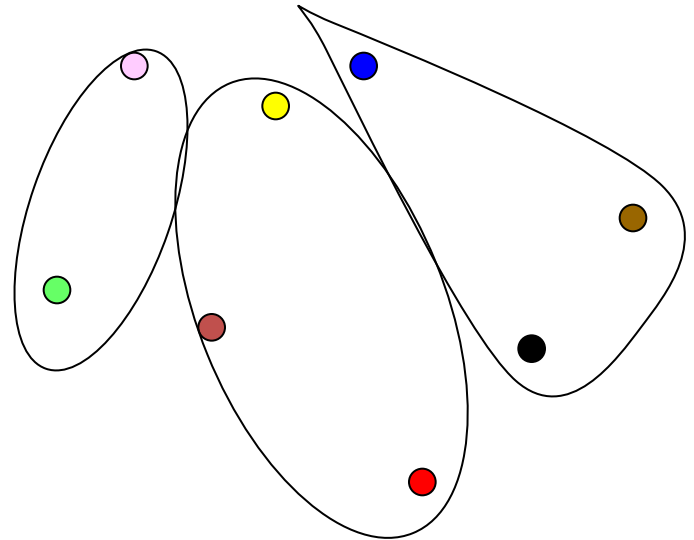
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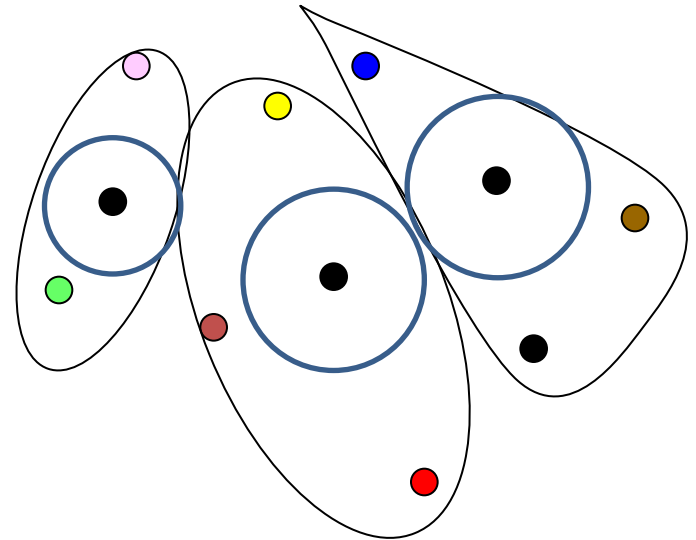


$$d(v, j) = 0.5 \log((2\pi)^D |C_j|) + 0.5(v - m_j)^T C_j^{-1} (v - m_j) - \log(w_i)$$



# K-means

1. Initialize cluster parameters
2. For each data point  $x$ , find the distance from each cluster
  - $d_{cluster} = d(x, cluster)$
3. Put data point in the cluster of the closest centroid
  - Cluster for which  $d_{cluster}$  is minimum
4. When all data points clustered, recompute cluster parameters
  - Means, variances, weights



$$m_j = \frac{1}{N_j} \sum_{v:j(v)=j} v$$

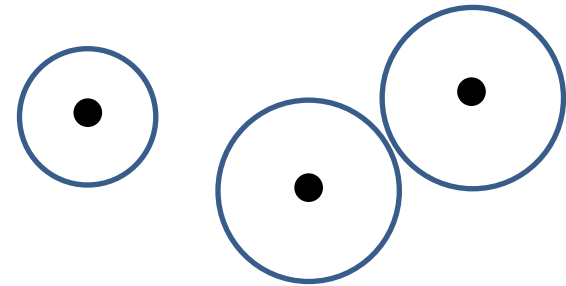
$$C_j = \frac{1}{N_j} \sum_{v:j(v)=j} (v - m_j)(v - m_j)^T$$

$$w_j = \frac{N_j}{N}$$

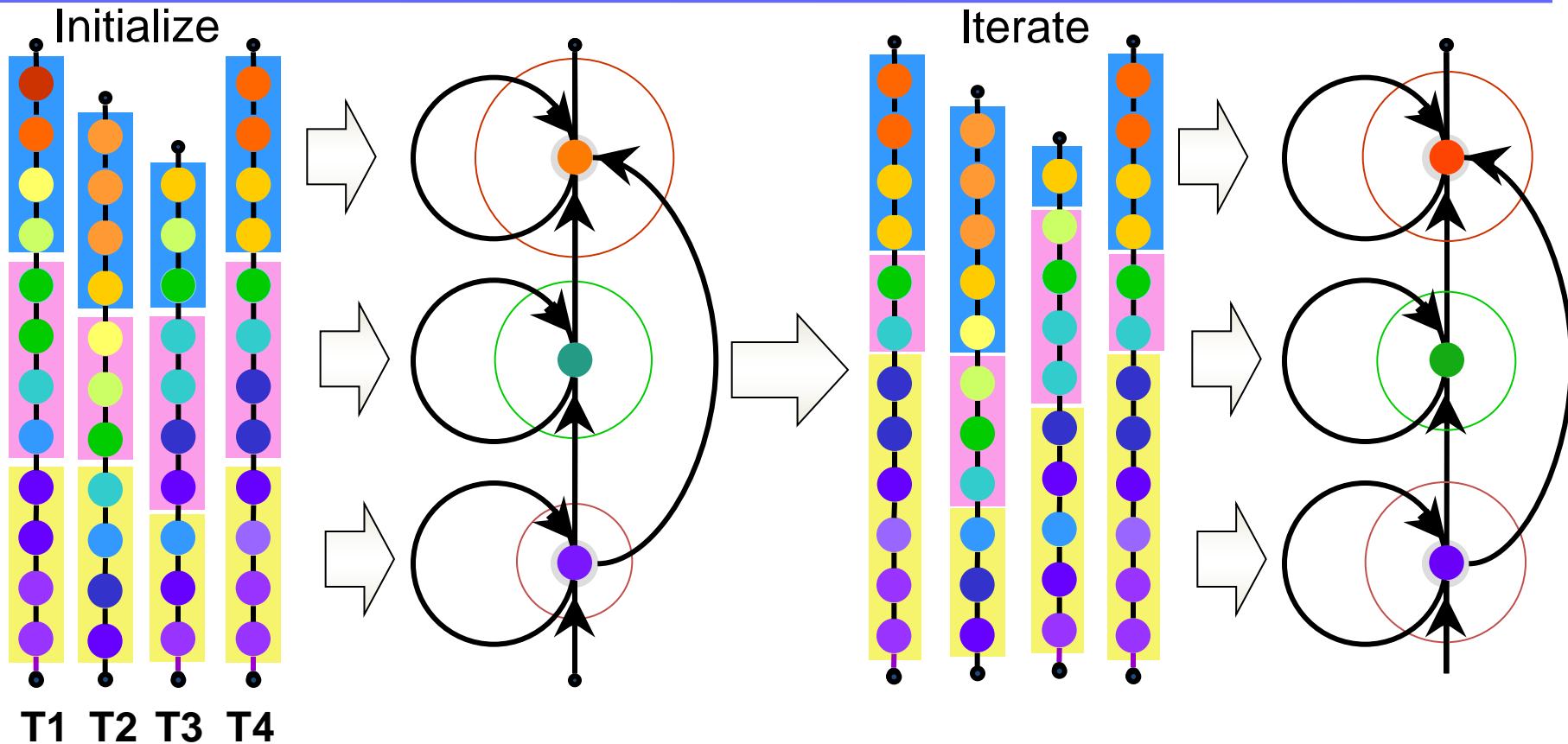
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---

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  - $d_{cluster} = d(x, cluster)$
3. Put data point in the cluster of the closest centroid
  - Cluster for which  $d_{cluster}$  is minimum
4. When all data points clustered, recompute cluster parameters
  - Means, variances, weights
5. If not converged, go back to 2



# Overall Segmental Kmeans



Identical to what we learned before, with one change:  
At each iteration, learn a Gaussian *mixture* distribution for each state.

# More on Gaussian Mixtures

---

- More common procedure:
  1. Train HMMs with 1-Gaussian per state HMMs using a first-pass of segmental K-means
    - This is a trivial Gaussian mixture with only one Gaussian
  2. *Split* the Gaussians in the state output distributions to obtain a larger Gaussian mixture at each state
  3. Run segmental K-means to convergence with updated Gaussian Mixtures
  4. If desired number of Gaussians not obtained for each state, return to 2
- What is “splitting”? What is the ideal no. of Gaussians?
  - We get to this shortly

# A Better Technique

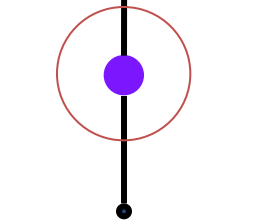
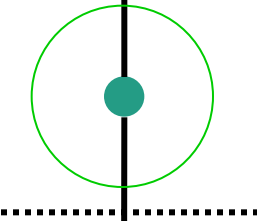
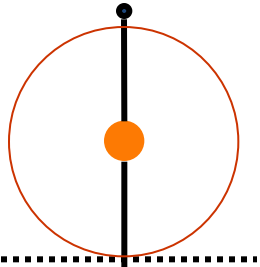
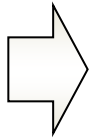
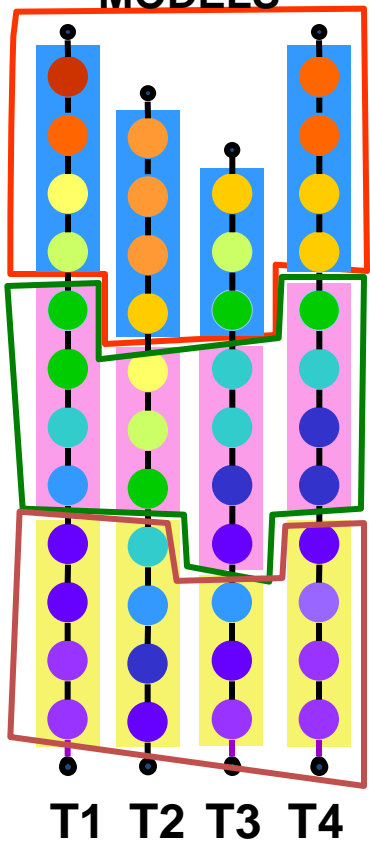
---

- The Segmental K-means technique uniquely assigns each observation to one state
- However, this is only an estimate and may be wrong
- A better approach is to take a “soft” decision
  - Assign each observation to *every* state with a probability

# Training by segmentation:

## Soft Assignment (1 Gaussian case)

MODELS



$$\mu_j = \frac{1}{\sum_{i \in \text{All vectors}} f_{i,j}} \sum_{i \in \text{All vectors}} f_{i,j} x_i$$

$$\sum_{j \in \text{all segments}} f_{i,j} = 1$$

Assignment is *fractioned*:  
Every segment gets a piece of  
every vector

Means and variances are computed  
from fractioned vectors

Where do the fractions come from?

# The “probability” of a state

---

- The probability assigned to any state  $s$ , for any observation  $x_t$  is the probability that the process was at  $s$  when it generated  $x_t$

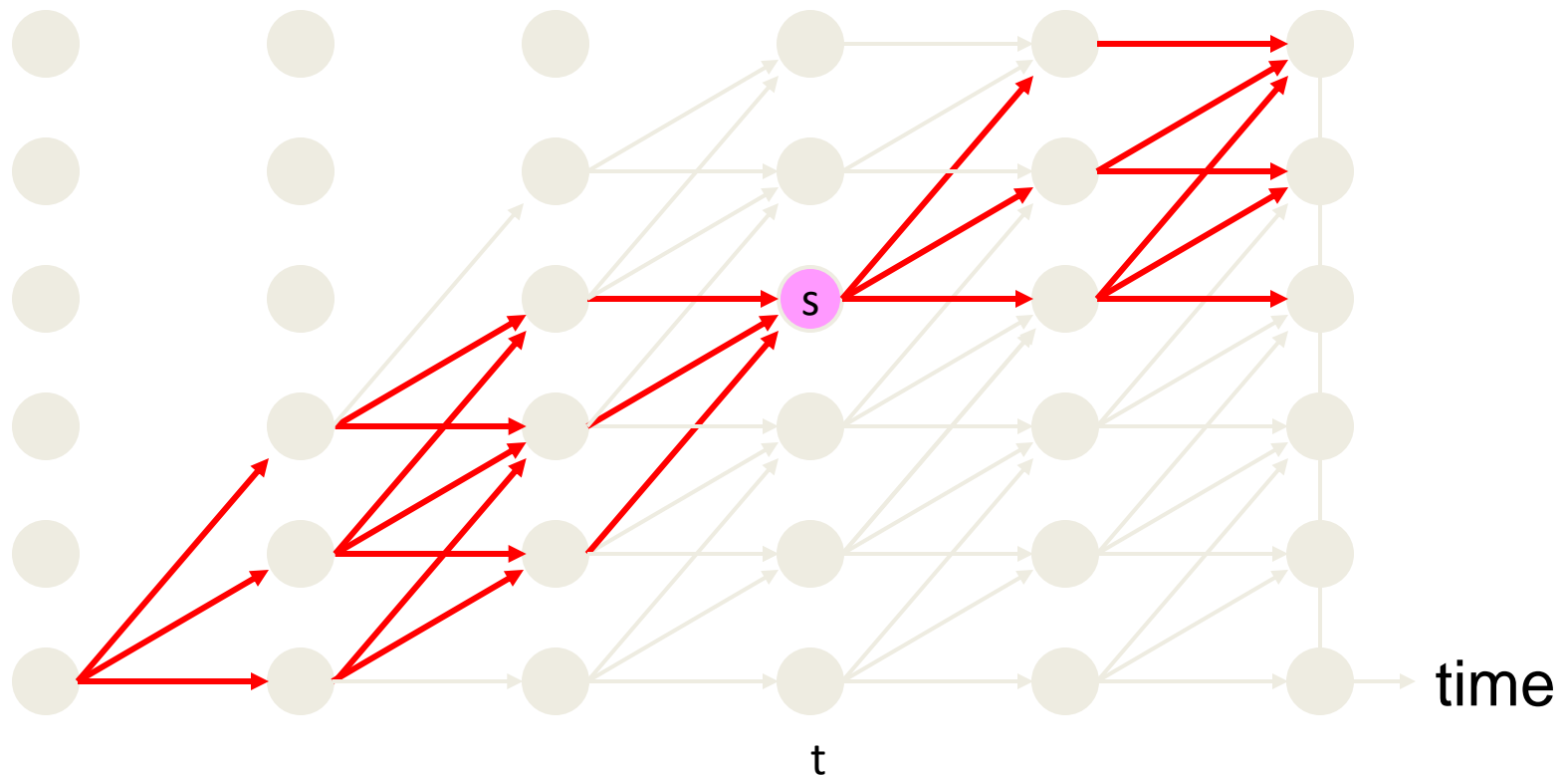
- We want to compute

$$P(\text{state}(t) = s \mid x_1, x_2, \dots, x_T) \propto P(\text{state}(t) = s, x_1, x_2, \dots, x_T)$$

- We will compute  $P(\text{state}(t) = s, x_1, x_2, \dots, x_T)$  first
  - This is the probability that the process visited  $s$  at time  $t$  while producing the entire observation

# Probability of Assigning an Observation to a State

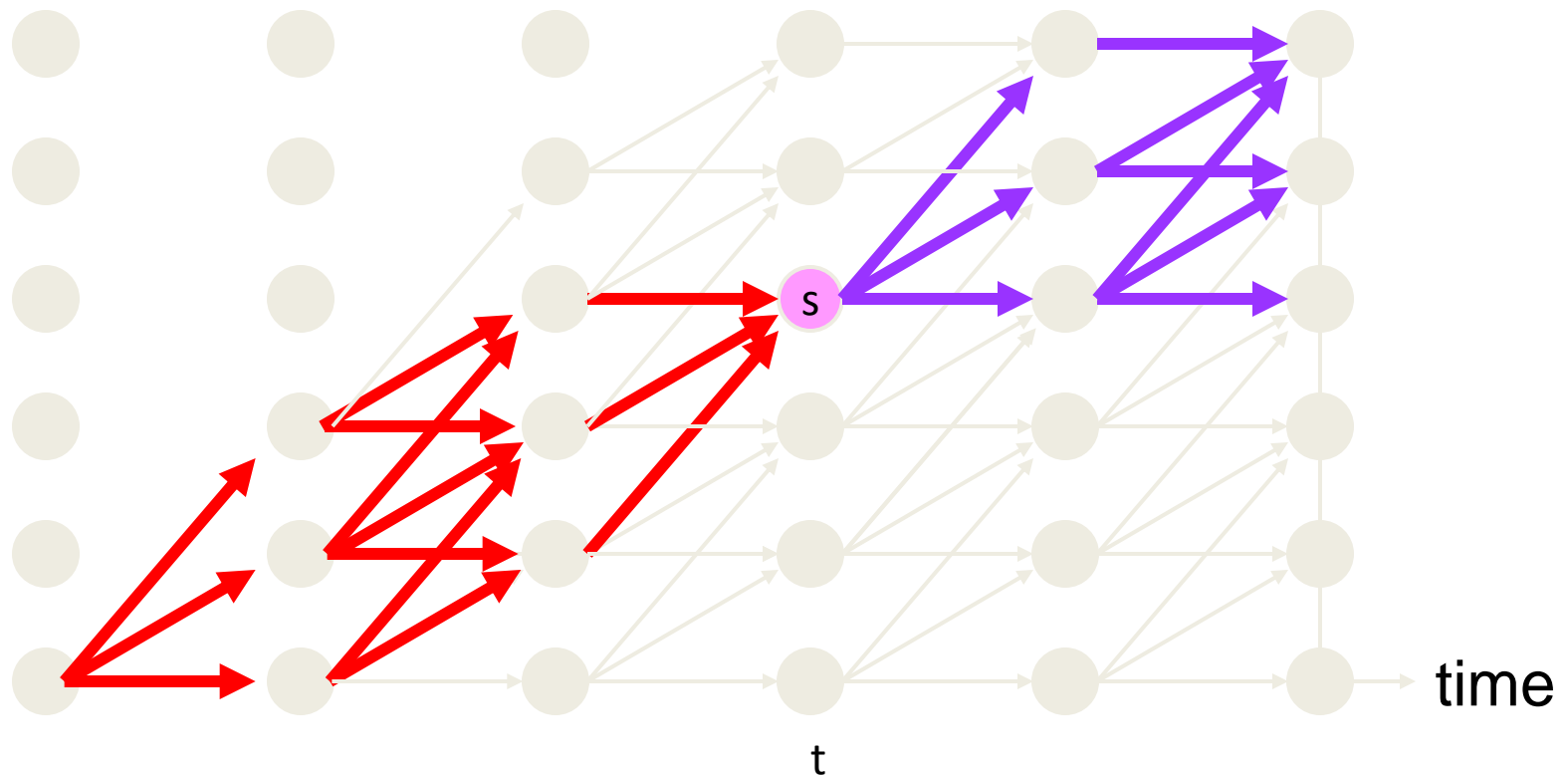
- The probability that the HMM was in a particular state  $s$  when generating the observation sequence is the probability that it followed a state sequence that passed through  $s$  at time  $t$





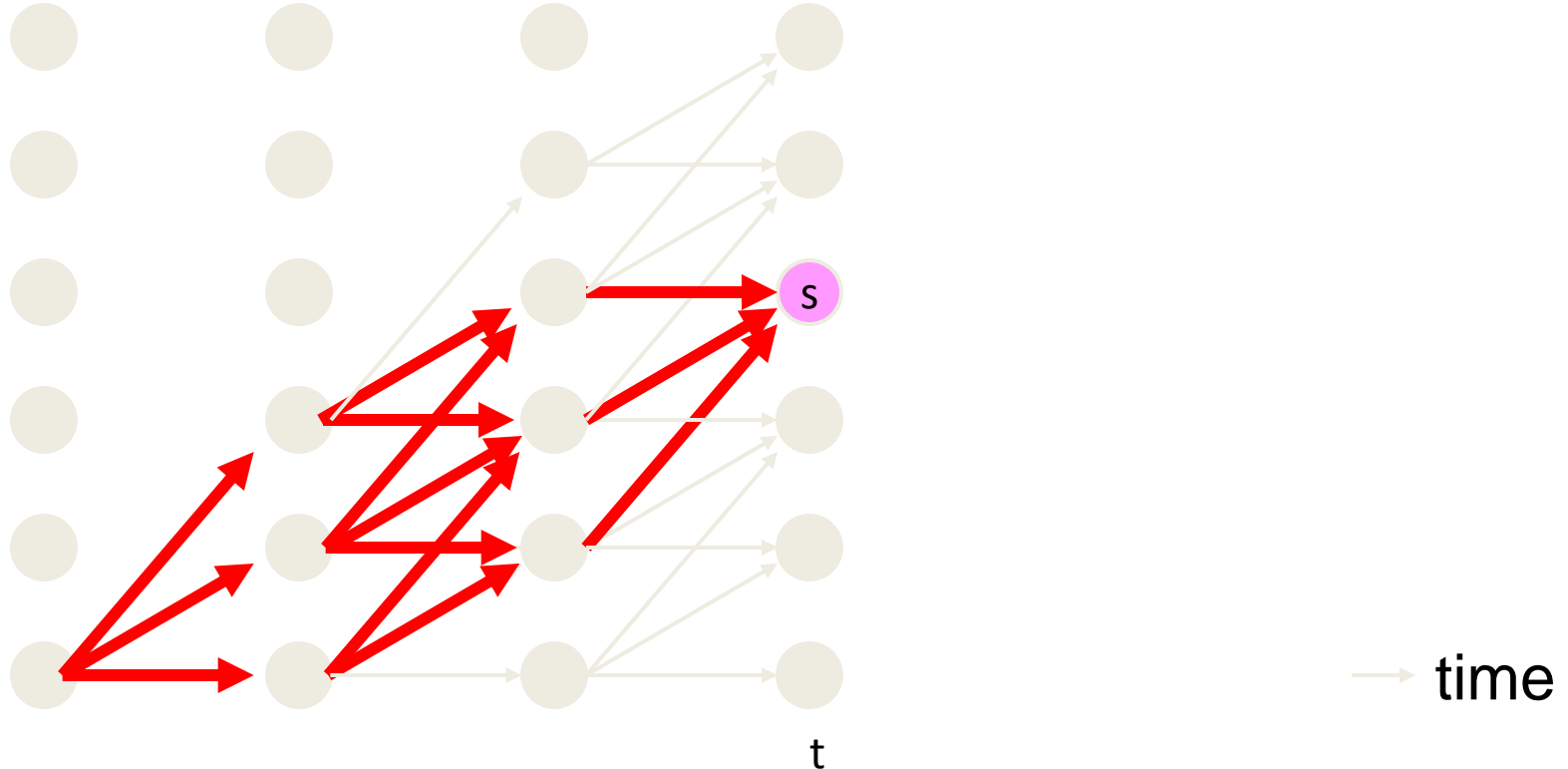
# Probability of Assigning an Observation to a State

- This can be decomposed into two multiplicative sections
  - The section of the lattice leading into state  $s$  at time  $t$  and the section leading out of it



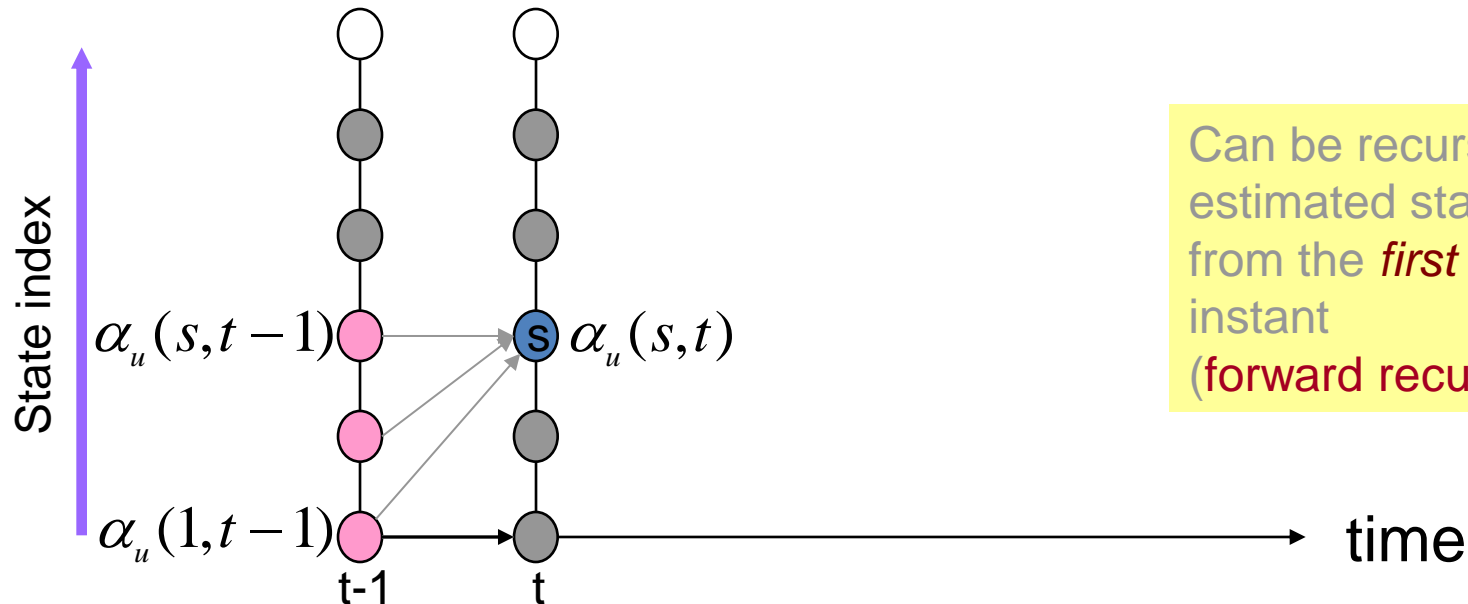
# Probability of Assigning an Observation to a State

- The probability of the red section is the total probability of all state sequences ending at state  $s$  at time  $t$ 
  - This is simply  $\alpha(s, t)$
  - Can be computed using the forward algorithm



# The forward algorithm

$$\alpha_u(s, t) = P(x_{u,1}, x_{u,2}, \dots, x_{u,t}, \text{state}(t) = s | \lambda)$$

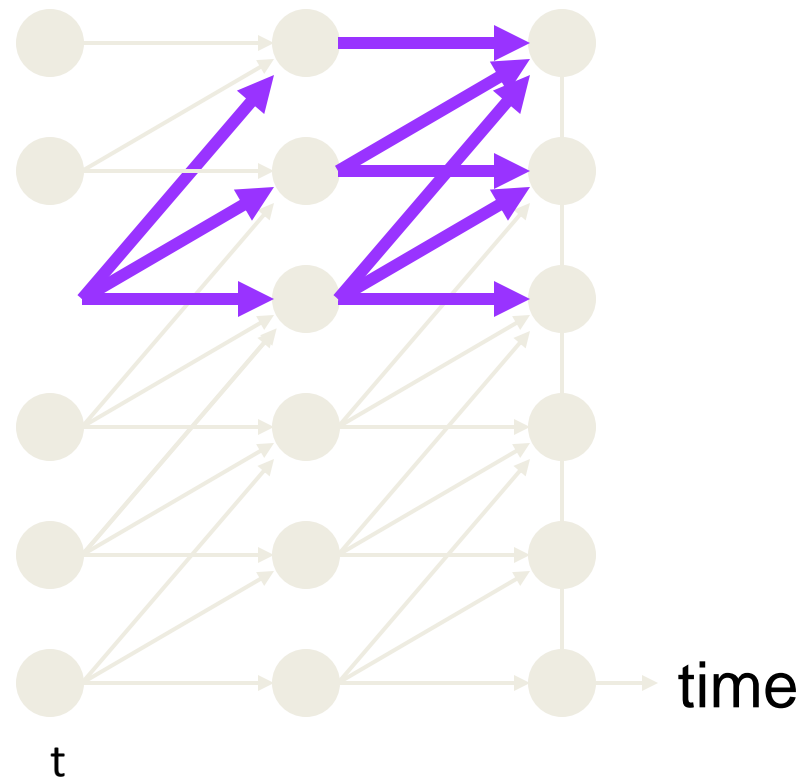


$$\alpha_u(s, t) = \sum_{s'} \alpha_u(s', t-1) P(s|s') P(x_{u,t} | s)$$

$\lambda$  represents the complete current set of HMM parameters

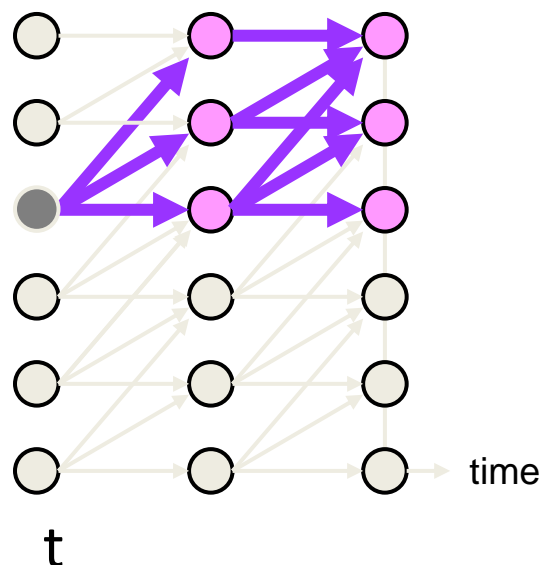
# The Future Paths

- The blue portion represents the probability of all state sequences that began at state  $s$  at time  $t$ 
  - Like the red portion it can be computed using a *backward recursion*



# The Backward Recursion

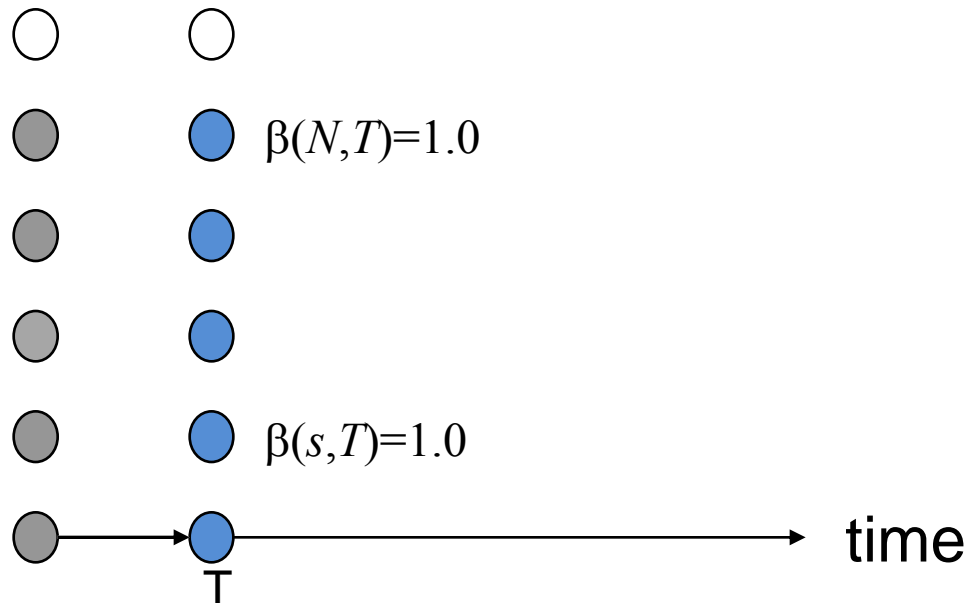
$$\beta_u(s, t) = P(x_{u,t+1}, x_{u,t+2}, \dots, x_{u,T} | \text{state}(t) = s, \lambda)$$



- $\beta_u(s, t)$  is the total probability of ALL state sequences that depart from  $s$  at time  $t$ , and all observations after  $x_t$

# The Backward Recursion

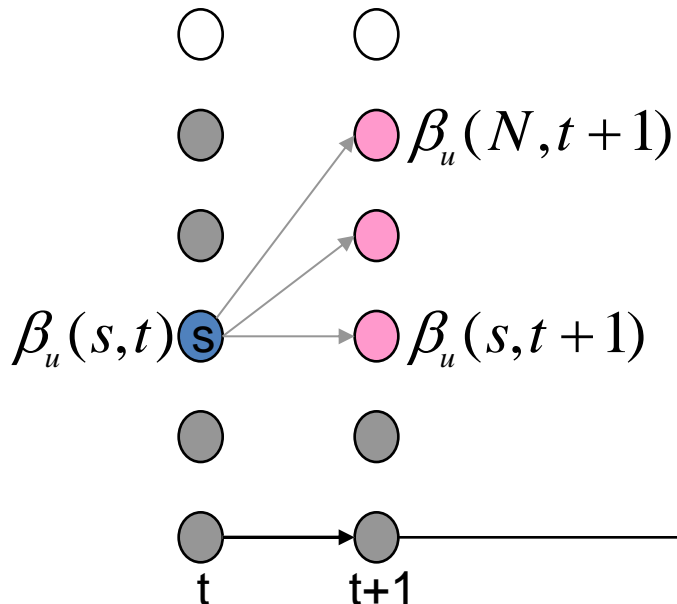
$$\beta_u(s, t) = P(x_{u,t+1}, x_{u,t+2}, \dots, x_{u,T} | \text{state}(t) = s, \lambda)$$



- $\beta(s, T) = 1$  at the final time instant for all valid final states
  - Since the future is a deterministic nothing..

# The Backward Recursion

$$\beta_u(s, t) = P(x_{u,t+1}, x_{u,t+2}, \dots, x_{u,T} | \text{state}(t) = s, \lambda)$$



Can be recursively estimated starting from the *final* time instant (backward recursion)

$$\beta_u(s, t) = \sum_{s'} \beta_u(s', t+1) P(s'|s) P(x_{u,t+1}|s')$$

- Note: Beta for any time  $t$  does *not* include the contribution of the observation at that time
  - $\beta(s, t)$  does not factor in  $P(x_t|s)$

# The backward algorithm

---

1. Initialize all beta terms at  $t=T$ :

$$\beta(s, T) = 1$$

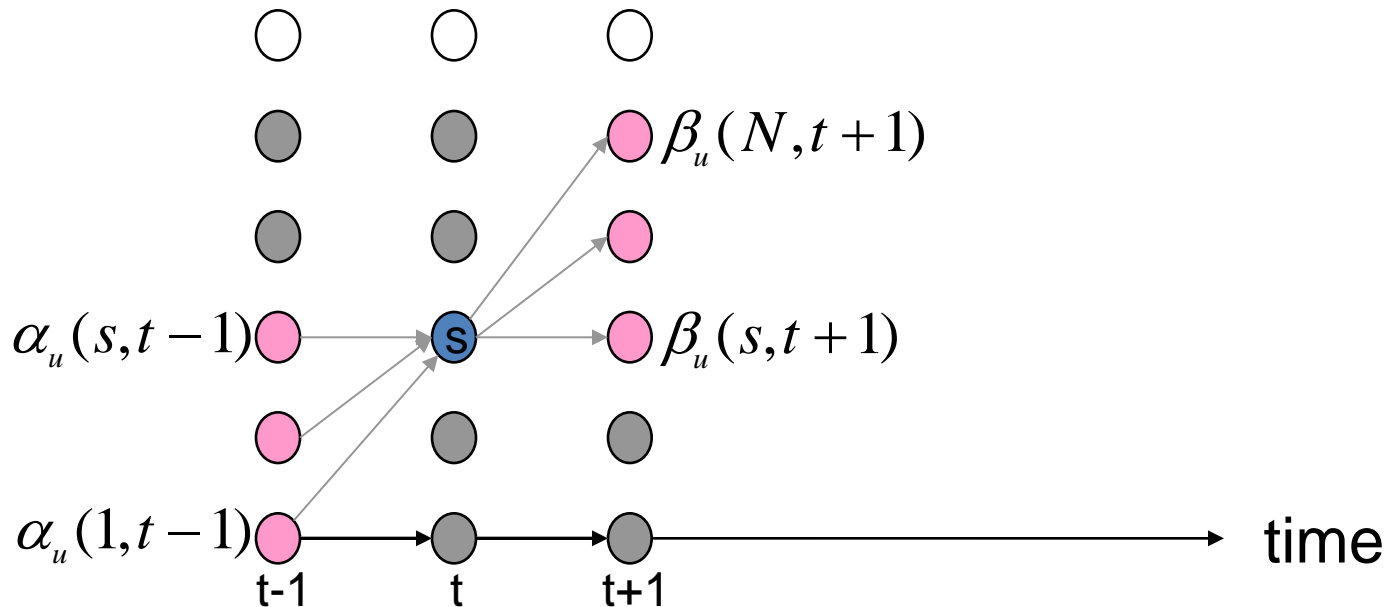
2. Recursively estimate betas for all prior time steps

$$\beta_u(s, t) = \sum_{s'} \beta_u(s', t+1) P(s'|s) P(x_{u,t+1}|s')$$



# The complete probability

$$\alpha_u(s, t) \beta_u(s, t) = P(x_{u,1}, x_{u,2}, \dots, x_{u,T}, \text{state}(t) = s | \lambda)$$



$$= P(\mathbf{X}_u, \text{state}(t) = s | \lambda)$$

# Posterior probability of a state

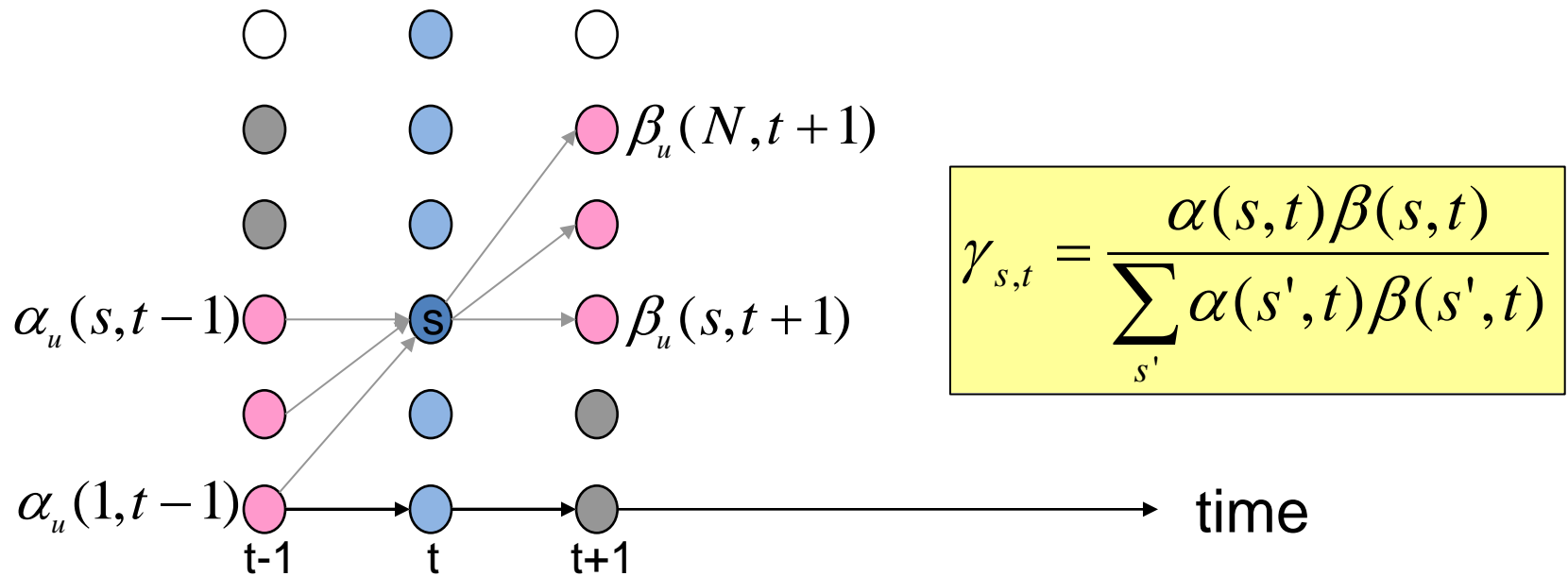
---

- The probability that the process was in state  $s$  at time  $t$ , given that we have observed the data is obtained by simple normalization

$$P(state(t) = s | \mathbf{X}_u, \lambda) = \frac{P(\mathbf{X}_u, state(t) = s | \lambda)}{\sum_{s'} P(\mathbf{X}_u, state(t) = s' | \lambda)} = \frac{\alpha_u(s, t) \beta_u(s, t)}{\sum_{s'} \alpha_u(s', t) \beta_u(s', t)}$$

- This term is often referred to as the gamma term and denoted by  $\gamma_{s,t}$

# The complete probability



- The gamma at any state at any time is obtained by normalizing the product of alphas and betas to sum to 1.0 over the corresponding column of the trellis

# Update Rules

---

- Once we have the state probabilities (the gammas) the update rules are obtained through a simple modification of the formulae used for segmental K-means
  - This new learning algorithm is known as the Baum-Welch learning procedure
- Case1: State output densities are Gaussians

# Update Rules

$$\mu_s = \frac{1}{N_s} \sum_{x \in s} x$$

$$C_s = \frac{1}{N_s} \sum_{x \in s} (x - \mu_s)^T (x - \mu_s)$$

Segmental K-means

$$\mu_s = \frac{\sum_u \sum_t \gamma_{u,s,t} x_{u,t}}{\sum_u \sum_t \gamma_{u,s,t}}$$

$$C_s = \frac{\sum_u \sum_t \gamma_{u,s,t} (x - \mu_s)^T (x - \mu_s)}{\sum_u \sum_t \gamma_{s,u,t}}$$

Baum Welch

- A similar update formula reestimates transition probabilities
- The *initial* state probabilities  $P(s)$  also have a similar update rule

## Case 2: State output densities are Gaussian Mixtures

---

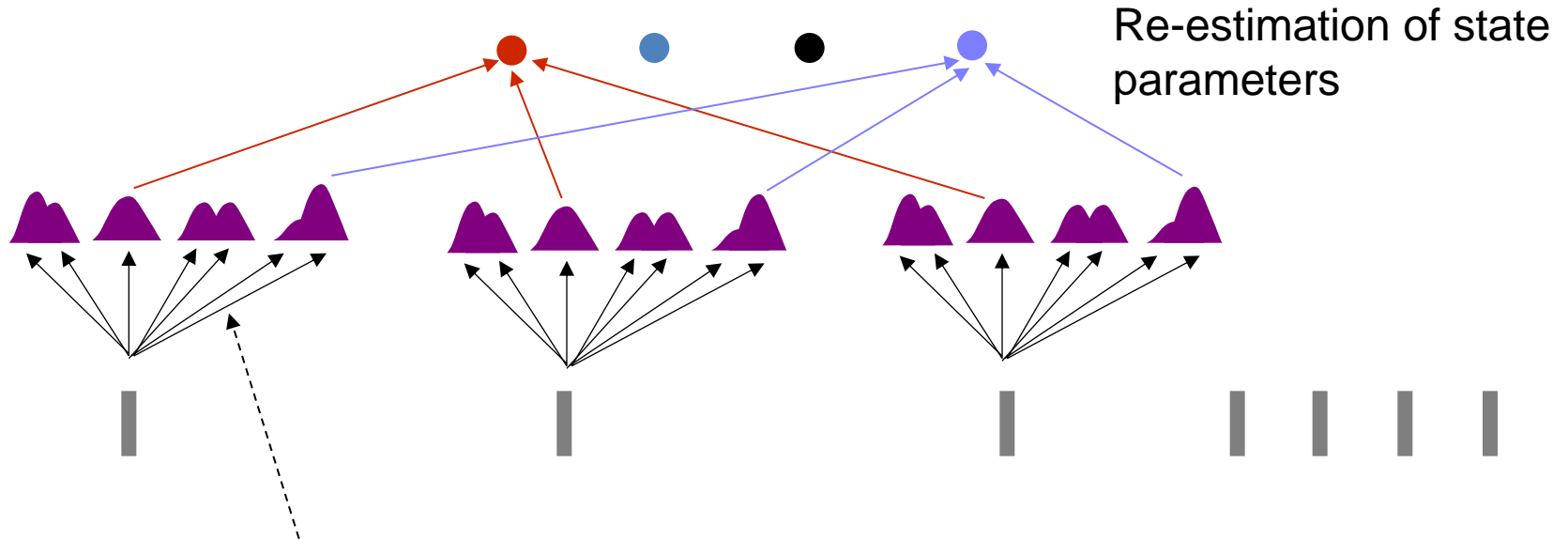
- When state output densities are Gaussian *mixtures*, more parameters must be estimated

$$P(x | s) = \sum_{i=0}^{K-1} w_{s,i} \textit{Gaussian}(x; \mu_{s,i}, C_{s,i})$$

- The mixture weights  $w_{s,i}$ , mean  $\mu_{s,i}$  and covariance  $C_{s,i}$  of every Gaussian in the distribution of each state must be estimated

# Splitting the Gamma

We split the gamma for any state among all the Gaussians at that state



A posteriori probability that the  $t^{\text{th}}$  vector was generated by the  $k^{\text{th}}$  Gaussian of state  $s$

$$\gamma_{k,s,u,t} = P(state(t) = s | \mathbf{X}_u, \lambda) P(k^{th}.Gaussian | state(t) = s, x_{u,t}, \lambda)$$

# Splitting the Gamma among Gaussians

---

A posteriori probability that the  $t^{\text{th}}$  vector was generated by the  $k^{\text{th}}$  Gaussian of state  $s$

$$\gamma_{k,s,t} = P(\text{state}(t) = s \mid \mathbf{X}, \lambda) P(k^{\text{th}} \text{ Gaussian} \mid \text{state}(t) = s, x_t, \lambda)$$

$$\gamma_{k,s,t} = \gamma_{s,t} \frac{w_{k,s} \frac{1}{\sqrt{(2\pi)^D |C_{k,s}|}} e^{-\frac{1}{2}(x_t - \mu_{k,s})^T C_{k,s}^{-1} (x_t - \mu_{k,s})}}{\sum_{k'} w_{k',s} \frac{1}{\sqrt{(2\pi)^D |C_{k',s}|}} e^{-\frac{1}{2}(x_t - \mu_{k',s})^T C_{k',s}^{-1} (x_t - \mu_{k',s})}}$$



# Updating HMM Parameters

$$\tilde{\mu}_{k,s} = \frac{\sum_u \sum_t \gamma_{k,s,u,t} x_{u,t}}{\sum_u \sum_t \gamma_{k,s,u,t}}$$

$$\tilde{\mathbf{C}}_{k,s} = \frac{\sum_u \sum_t \gamma_{k,s,u,t} (x_{u,t} - \tilde{\mu}_{k,s})(x_{u,t} - \tilde{\mu}_{k,s})^T}{\sum_u \sum_t \gamma_{k,s,u,t}}$$

$$\tilde{w}_{k,s} = \frac{\sum_u \sum_t \gamma_{k,s,u,t}}{\sum_u \sum_t \sum_j \gamma_{j,s,u,t}}$$

- Note: Every observation contributes to the update of parameter values of every Gaussian of every state

# Overall Training Procedure: Single Gaussian PDF

---

- Determine a topology for the HMM
  - Initialize all HMM parameters
    - Initialize all allowed transitions to have the same probability
    - Initialize all state output densities to be Gaussians
      - We'll revisit initialization
1. Over all utterances, compute the “sufficient” statistics  $\sum_u \sum_t \gamma_{u,s,t}$   $\sum_u \sum_t \gamma_{u,s,t} x_{u,t}$   $\sum_u \sum_t \gamma_{u,s,t} (x - \mu_s)^T (x - \mu_s)$
  2. Use update formulae to compute new HMM parameters
  3. If the overall probability of the training data has not converged, return to step 1

# An Implementational Detail

- Step1 computes “buffers” over all utterance

$$\sum_u \sum_t \gamma_{u,s,t} = \sum_{u \in U_1} \sum_t \gamma_{u,s,t} + \sum_{u \in U_2} \sum_t \gamma_{u,s,t} + ..$$

Assuming  
1 gaussian/stt  
on this slide

$$\sum_u \sum_t \gamma_{u,s,t} x_{u,t} = \sum_{u \in U_1} \sum_t \gamma_{u,s,t} x_{u,t} + \sum_{u \in U_2} \sum_t \gamma_{u,s,t} x_{u,t} + ..$$

$$\sum_u \sum_t \gamma_{u,s,t} (x - \mu_s)^T (x - \mu_s) = \sum_{u \in U_1} \sum_t \gamma_{u,s,t} (x - \mu_s)^T (x - \mu_s) + \sum_{u \in U_2} \sum_t \gamma_{u,s,t} (x - \mu_s)^T (x - \mu_s) + ..$$

- This can be split and parallelized
  - $U_1, U_2$  etc. can be processed on separate machines

Machine 1

$$\sum_{u \in U_1} \sum_t \gamma_{u,s,t} \quad \sum_{u \in U_1} \sum_t \gamma_{u,s,t} x_{u,t}$$

$$\sum_{u \in U_1} \sum_t \gamma_{u,s,t} (x - \mu_s)(x - \mu_s)^T$$

Machine 2

$$\sum_{u \in U_2} \sum_t \gamma_{u,s,t} \quad \sum_{u \in U_2} \sum_t \gamma_{u,s,t} x_{u,t}$$

$$\sum_{u \in U_2} \sum_t \gamma_{u,s,t} (x - \mu_s)(x - \mu_s)^T$$

# An Implementational Detail

- Step2 *aggregates and adds* buffers before updating the models

$$\sum_u \sum_t \gamma_{u,s,t} = \sum_{u \in U_1} \sum_t \gamma_{u,s,t} + \sum_{u \in U_2} \sum_t \gamma_{u,s,t} + \dots$$

$$\sum_u \sum_t \gamma_{u,s,t} x_{u,t} = \sum_{u \in U_1} \sum_t \gamma_{u,s,t} x_{u,t} + \sum_{u \in U_2} \sum_t \gamma_{u,s,t} x_{u,t} + \dots$$

$$\sum_u \sum_t \gamma_{u,s,t} (x - \mu_s)^T (x - \mu_s) = \sum_{u \in U_1} \sum_t \gamma_{u,s,t} (x - \mu_s)^T (x - \mu_s) + \sum_{u \in U_2} \sum_t \gamma_{u,s,t} (x - \mu_s)^T (x - \mu_s) + \dots$$

$$\tilde{\mu}_{k,s} = \frac{\sum_u \sum_t \gamma_{k,s,u,t} x_{u,t}}{\sum_u \sum_t \gamma_{k,s,u,t}}$$

$$\tilde{\mathbf{C}}_{k,s} = \frac{\sum_u \sum_t \gamma_{k,s,u,t} (x_{u,t} - \tilde{\mu}_{k,s})(x_{u,t} - \tilde{\mu}_{k,s})^T}{\sum_u \sum_t \gamma_{k,s,u,t}}$$

$$\tilde{w}_{k,s} = \frac{\sum_u \sum_t \gamma_{k,s,u,t}}{\sum_u \sum_t \sum_j \gamma_{j,s,u,t}}$$

# An Implementational Detail

- Step2 *aggregates and adds* buffers before updating the models

$$\sum_u \sum_t \gamma_{u,s,t} = \sum_{u \in U_1} \sum_t \gamma_{u,s,t} + \sum_{u \in U_2} \sum_t \gamma_{u,s,t} + \dots$$

$$\sum_u \sum_t \gamma_{u,s,t} x_{u,t} = \sum_{u \in U_1} \sum_t \gamma_{u,s,t} x_{u,t} + \sum_{u \in U_2} \sum_t \gamma_{u,s,t} x_{u,t} + \dots$$

$$\sum_u \sum_t \gamma_{u,s,t} (x - \mu_s)^T (x - \mu_s) = \sum_{u \in U_1} \sum_t \gamma_{u,s,t} (x - \mu_s)^T (x - \mu_s) + \sum_{u \in U_2} \sum_t \gamma_{u,s,t} (x - \mu_s)^T (x - \mu_s) + \dots$$

$$\tilde{\mu}_{k,s} = \frac{\sum_u \sum_t \gamma_{k,s,u,t} x_{u,t}}{\sum_u \sum_t \gamma_{k,s,u,t}}$$

$$\tilde{\mathbf{C}}_{k,s} = \frac{\sum_u \sum_t \gamma_{k,s,u,t} (x_{u,t} - \tilde{\mu}_{k,s})(x_{u,t} - \tilde{\mu}_{k,s})^T}{\sum_u \sum_t \gamma_{k,s,u,t}}$$

Computed by  
machine 1

Computed by  
machine 2

$$\tilde{w}_{k,s} = \frac{\sum_u \sum_t \gamma_{k,s,u,t}}{\sum_u \sum_t \sum_j \gamma_{j,s,u,t}}$$

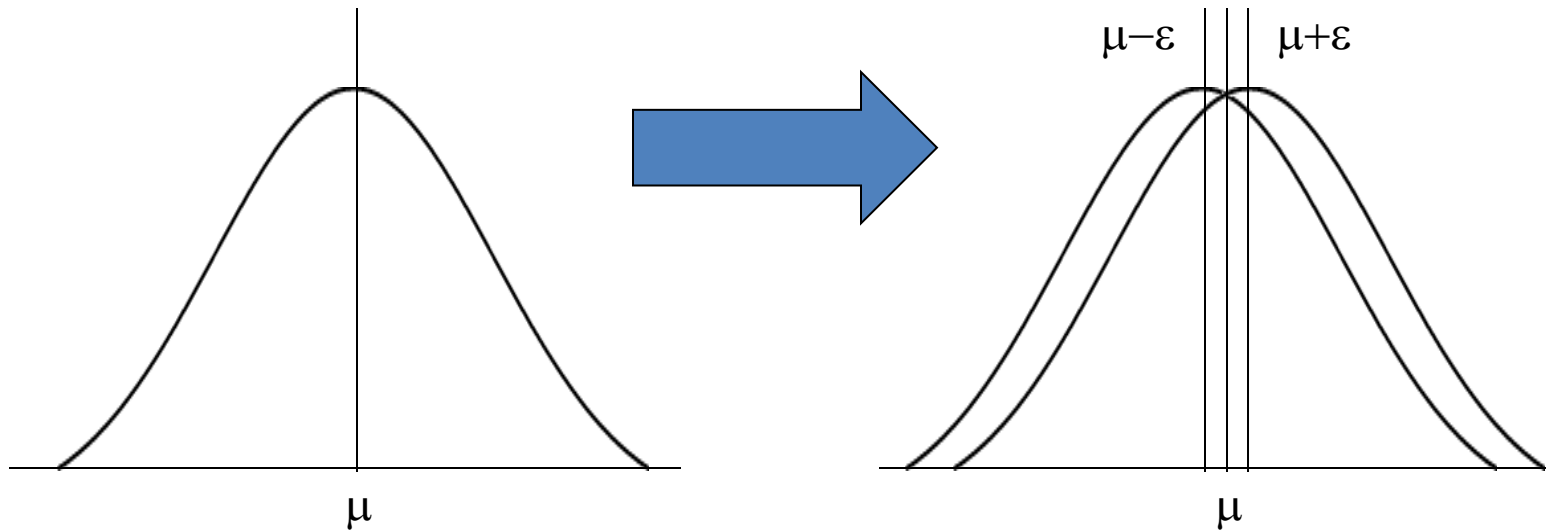
# Training for HMMs with *Gaussian Mixture* State Output Distributions

---

- Gaussian *Mixtures* are obtained by splitting
  1. Train an HMM with (single) Gaussian state output distributions
  2. Split the Gaussian with the largest variance
    - Perturb the mean by adding and subtracting a small number
    - This gives us 2 Gaussians. Partition the mixture weight of the Gaussian into two halves, one for each Gaussian
    - A mixture with  $N$  Gaussians now becomes a mixture of  $N+1$  Gaussians
  3. Iterate BW to convergence
  4. If the desired number of Gaussians not obtained, return to 2

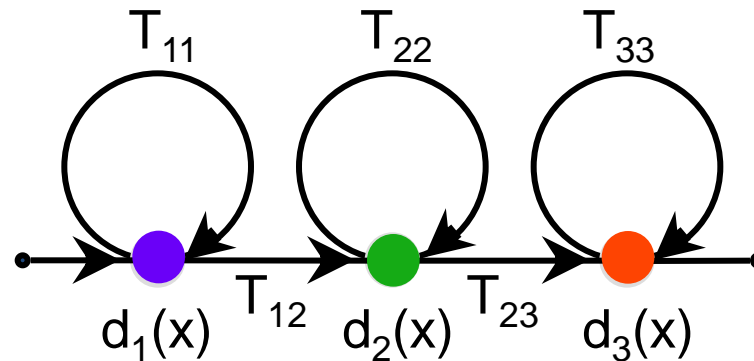
# Splitting a Gaussian

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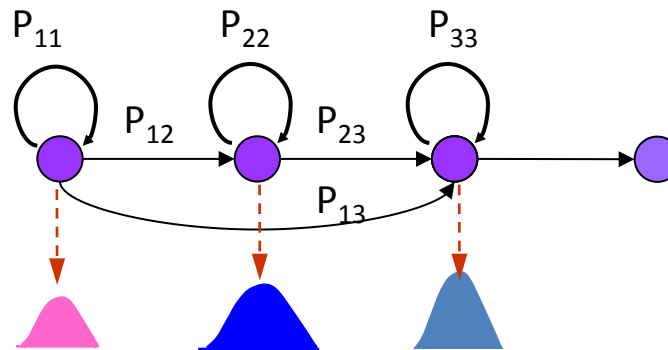


- The mixture weight  $w$  for the Gaussian gets shared as  $0.5w$  by each of the two split Gaussians

# Transition Probabilities



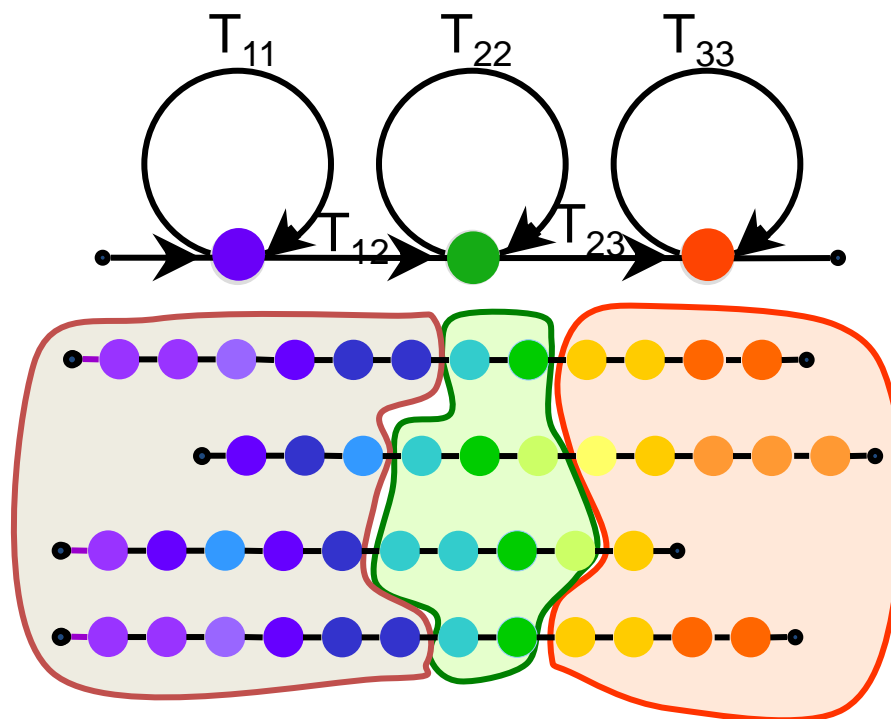
- We have seen how to compute transition penalties for templates



- How about transition probabilities in an HMM?
  - “Hard” estimation – by counting, as for templates
  - “Soft” estimation – need soft counts

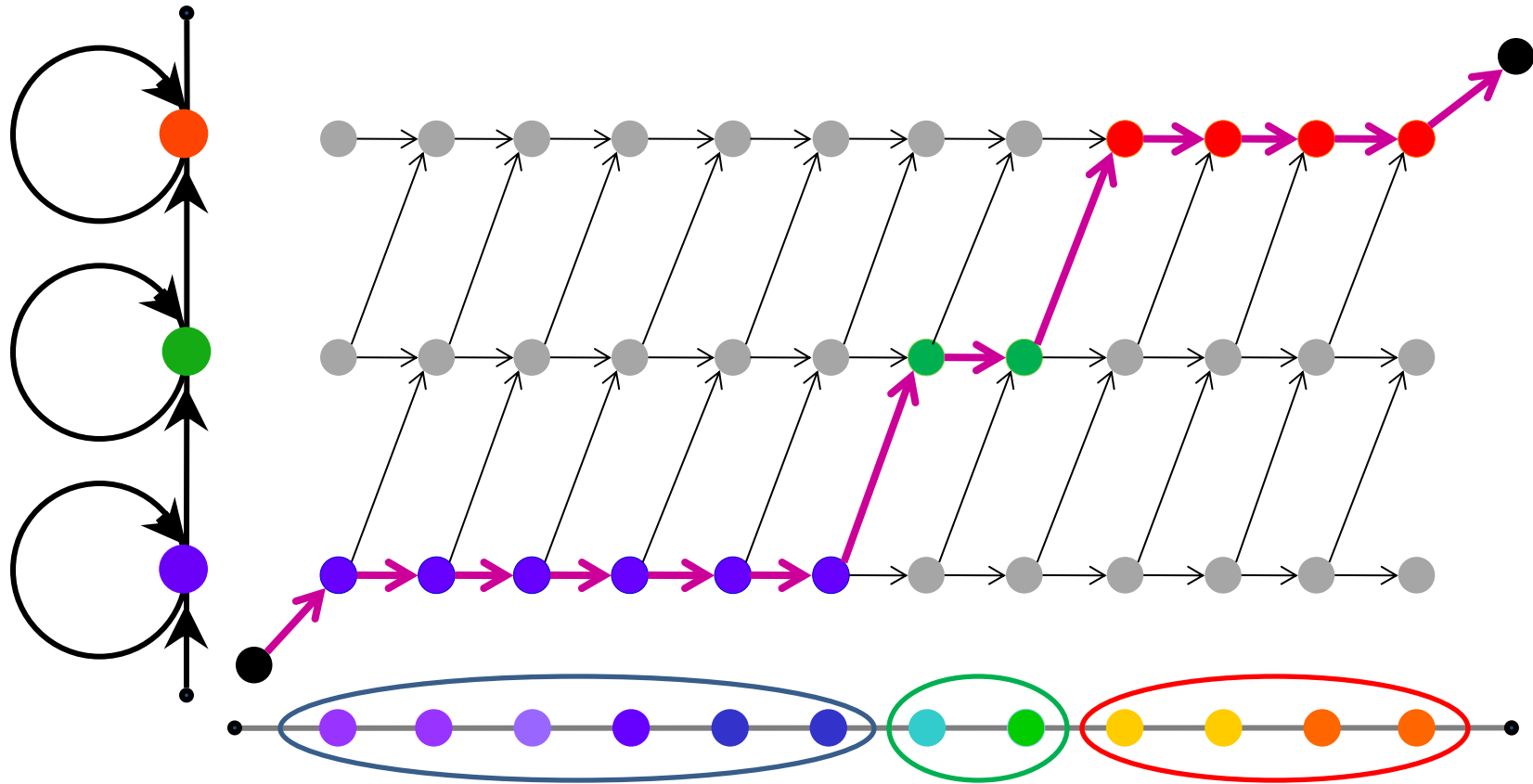


# Transition penalties by counting



- 20 vectors in state 1
  - 16 are followed by vectors in state 1
  - 4 are followed by vectors in state 2
- $P_{11} = 16/20 = 0.8 \rightarrow T_{11} = -\log(P_{11}) = -\log(0.8)$
- $P_{12} = 4/20 = 0.2 \rightarrow T_{12} = -\log(P_{12}) = -\log(0.2)$

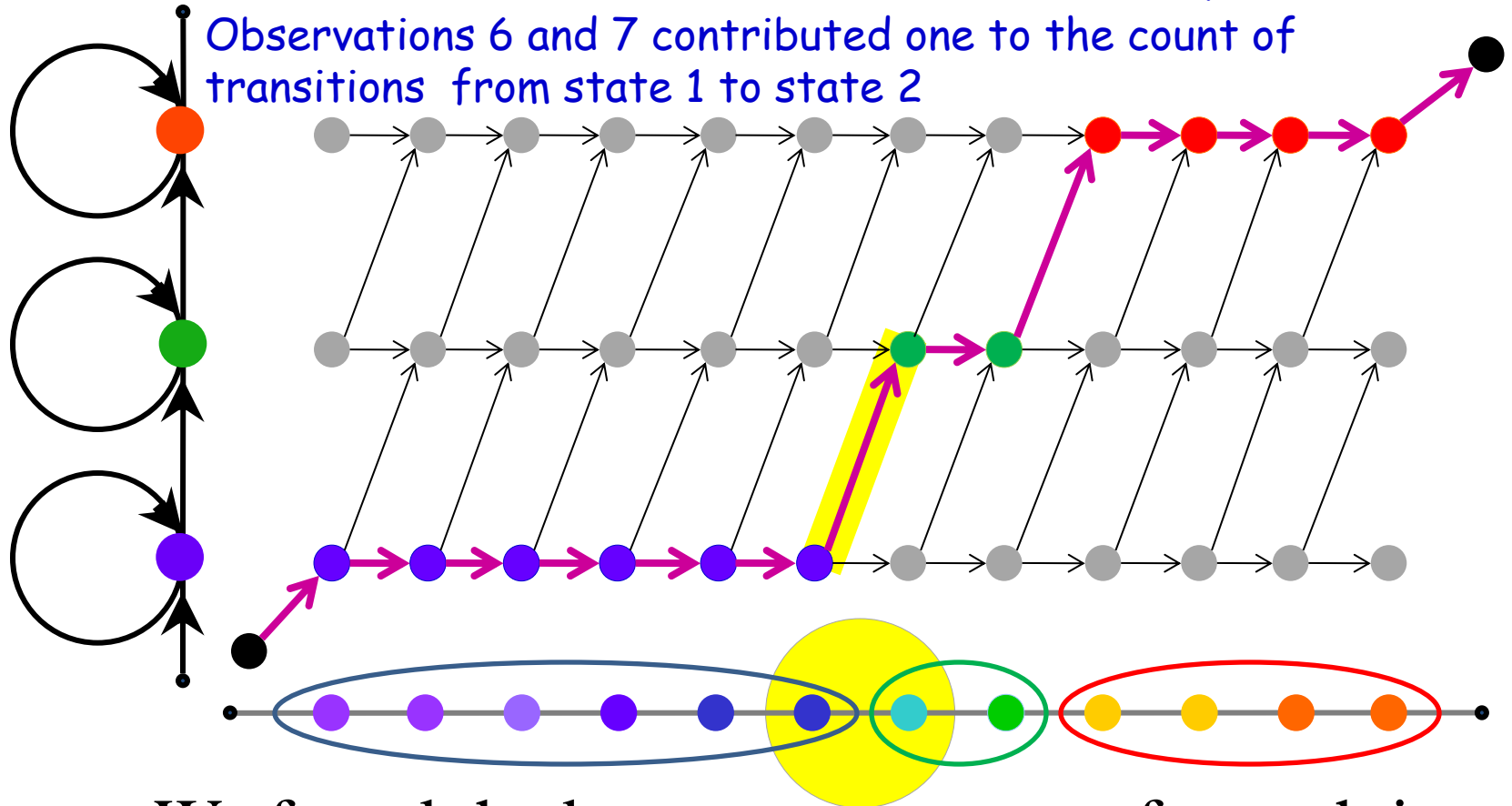
# Transitions by counting



- We found the best state sequence for each input
  - And counted transitions

# Transitions by counting

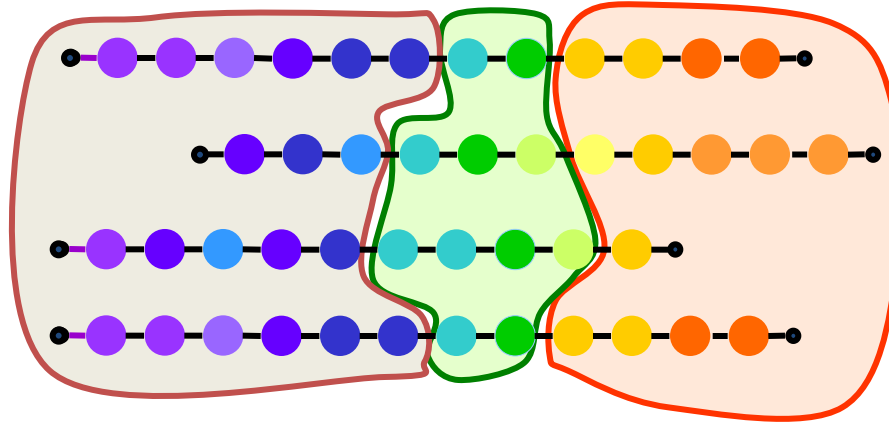
Observation no. 6 contributed one to the count of occurrences of state 1  
Observations 6 and 7 contributed one to the count of transitions from state 1 to state 2



- We found the best state sequence for each input
  - And counted transitions

# Probability of transitions

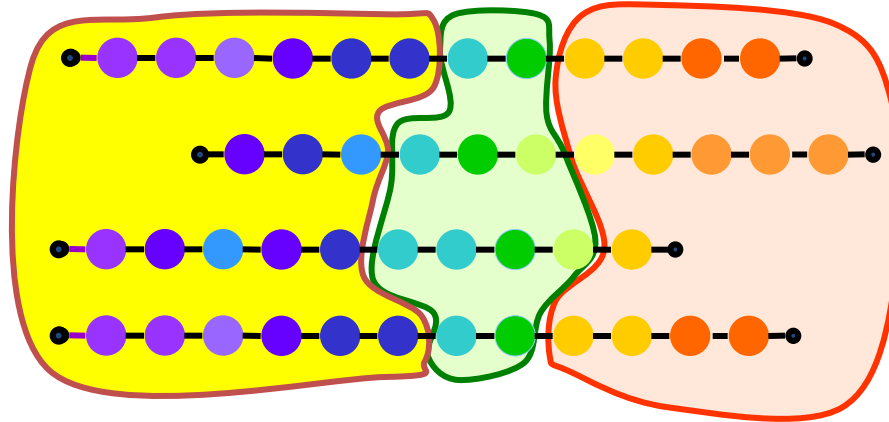
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- $P(\text{transition state } I \rightarrow \text{state } J) =$ 
  - $\text{Count transitions}(I, J) / \text{count instances}(I)$

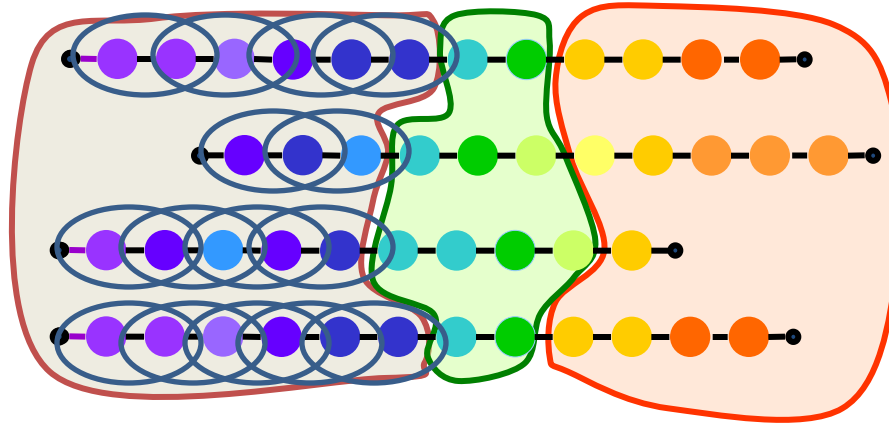
# Probability of transitions

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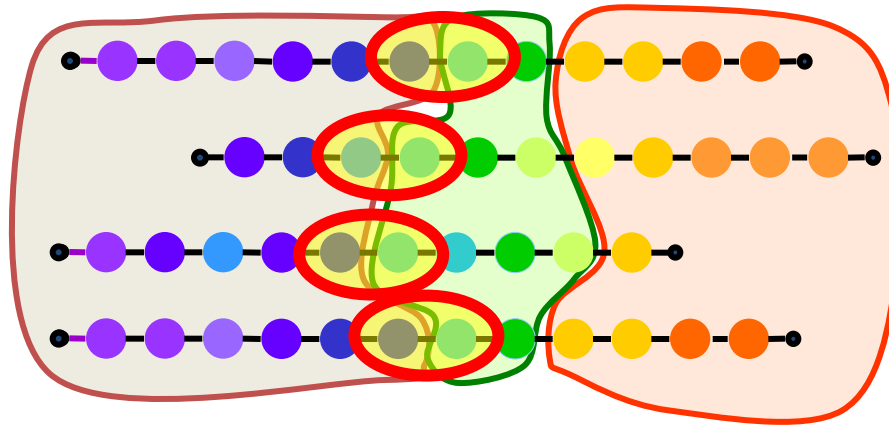
- $P(\text{transition state } I \rightarrow \text{state } J) =$ 
  - $\text{Count transitions}(I, J) / \text{count instances}(I)$
  - $\text{Count instances}(1) = 20$

# Probability of transitions



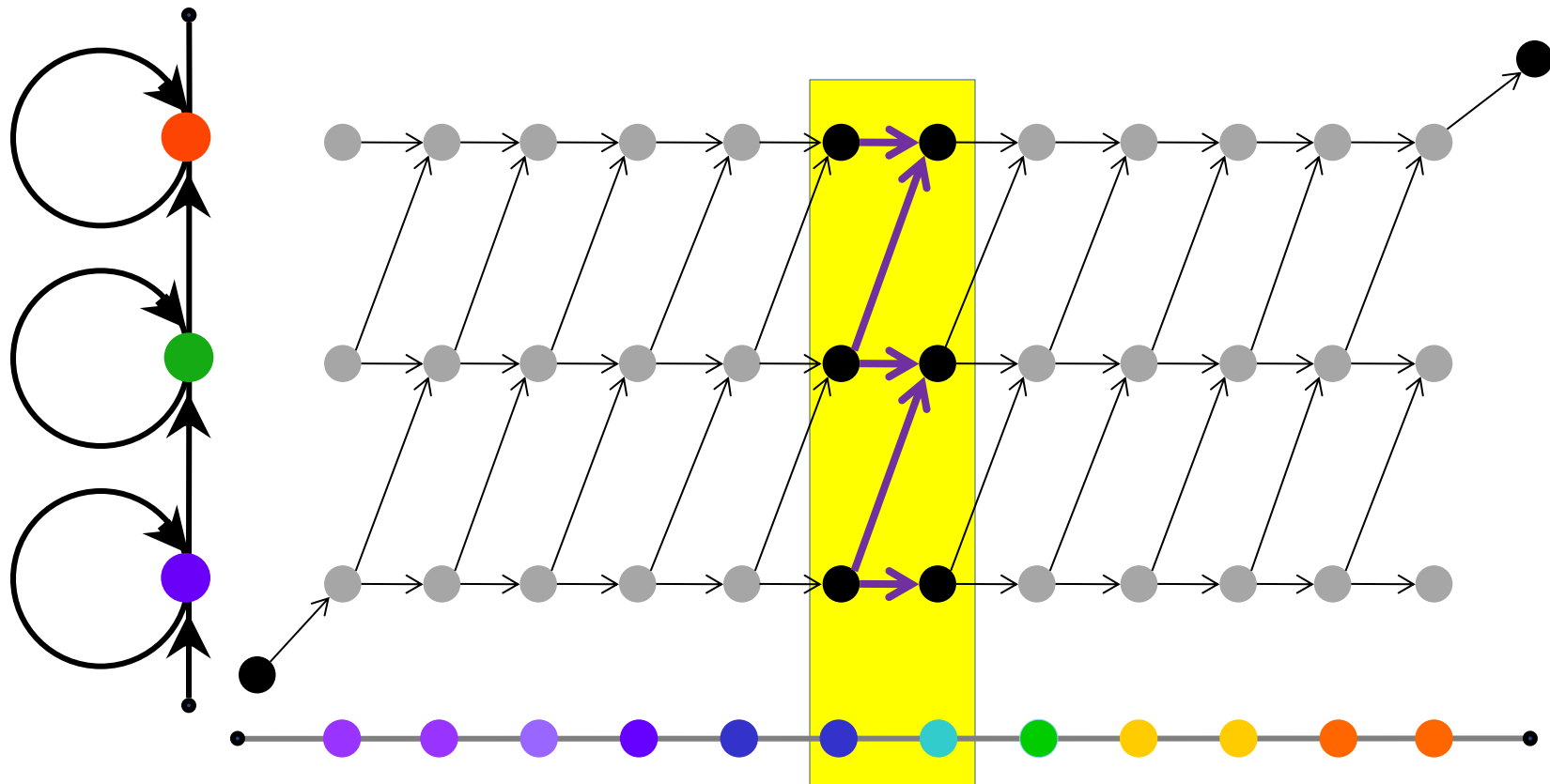
- $P(\text{transition state } I \rightarrow \text{state } J) =$ 
  - $\text{Count transitions}(I, J) / \text{count instances}(I)$
  - $\text{Count instances}(1) = 20$ 
    - $\text{Count transitions}(1, 1) = 16$
    - $P(\text{transition state } 1 \rightarrow \text{state } 1) = 0.8$

# Probability of transitions



- $P(\text{transition state } I \rightarrow \text{state } J) =$ 
  - $\text{Count transitions}(I, J) / \text{count instances}(I)$
  - $\text{Count instances}(1) = 20$ 
    - $\text{Count transitions}(1, 2) = 4$
    - $P(\text{transition state } 1 \rightarrow \text{state } 2) = 0.2$

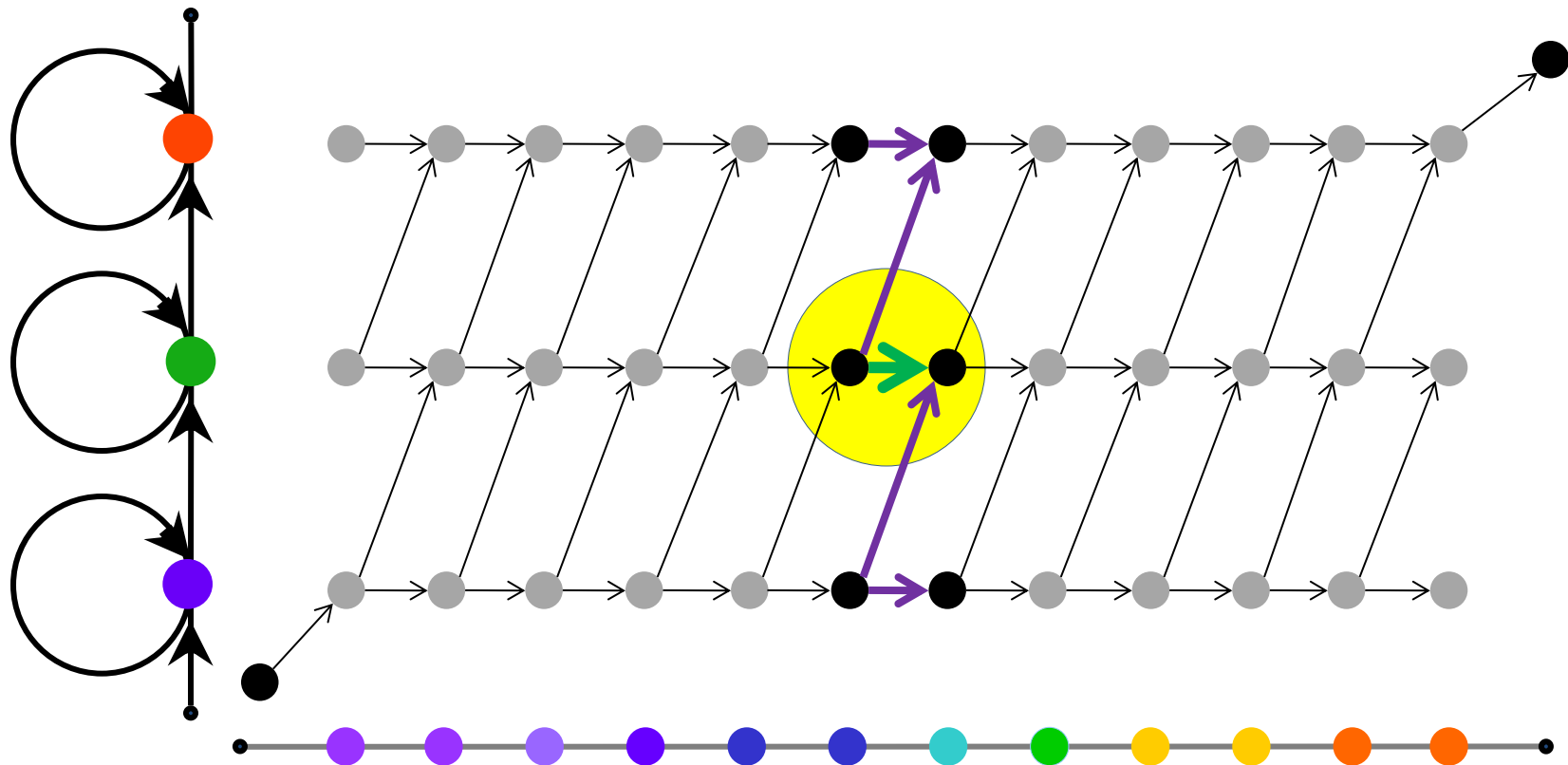
# Transitions by *soft* counting



- Each observation pair contributes to every transition
  - E.g. observations 6,7 contribute counts to all of the following:
    - Transition  $(1 \rightarrow 1)$ , Transition  $(1 \rightarrow 2)$ , Transition  $(2 \rightarrow 2)$ ,  
Transition  $(2 \rightarrow 3)$ , Transition  $(3 \rightarrow 3)$

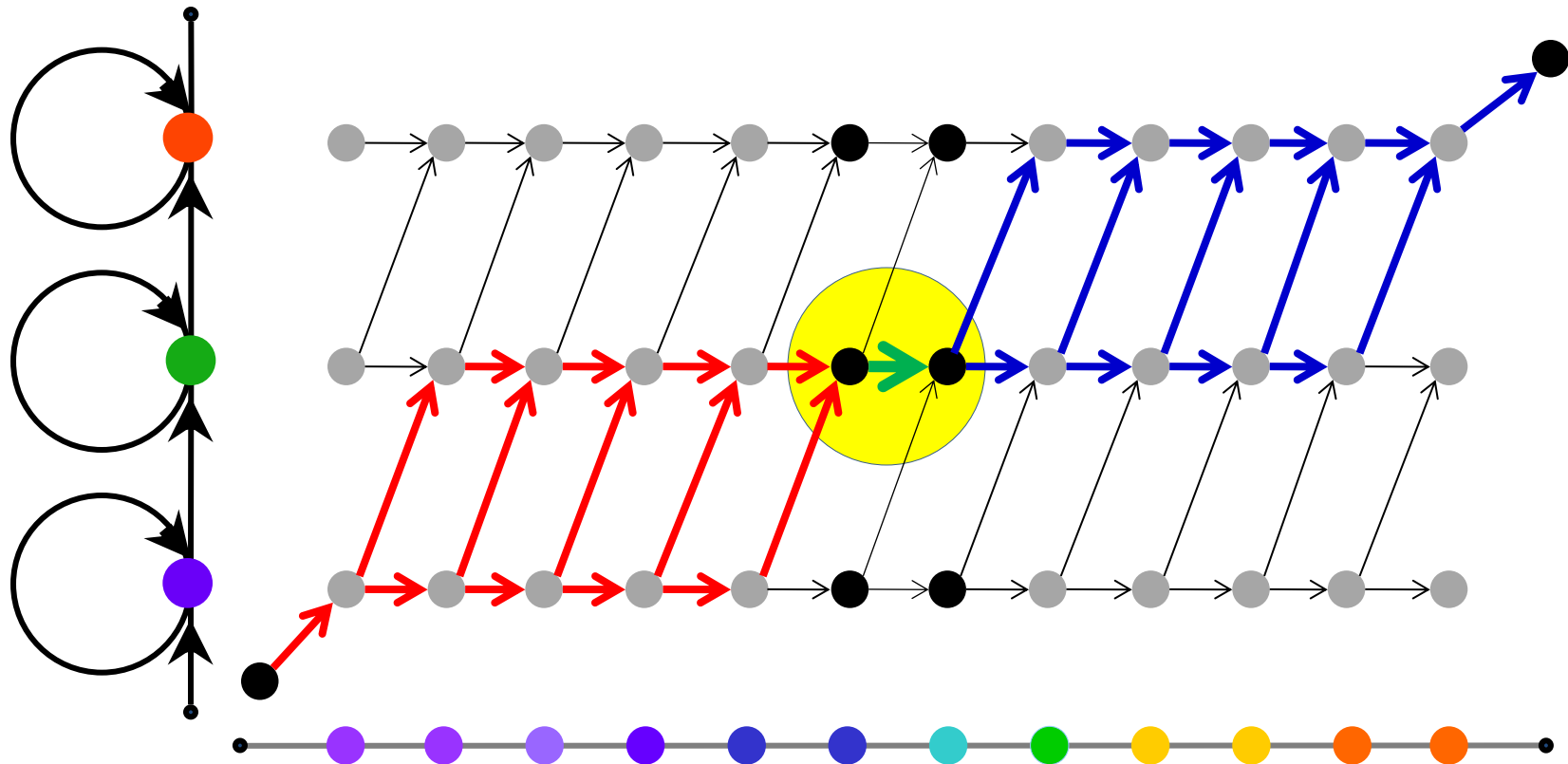


# Transitions by *soft* counting



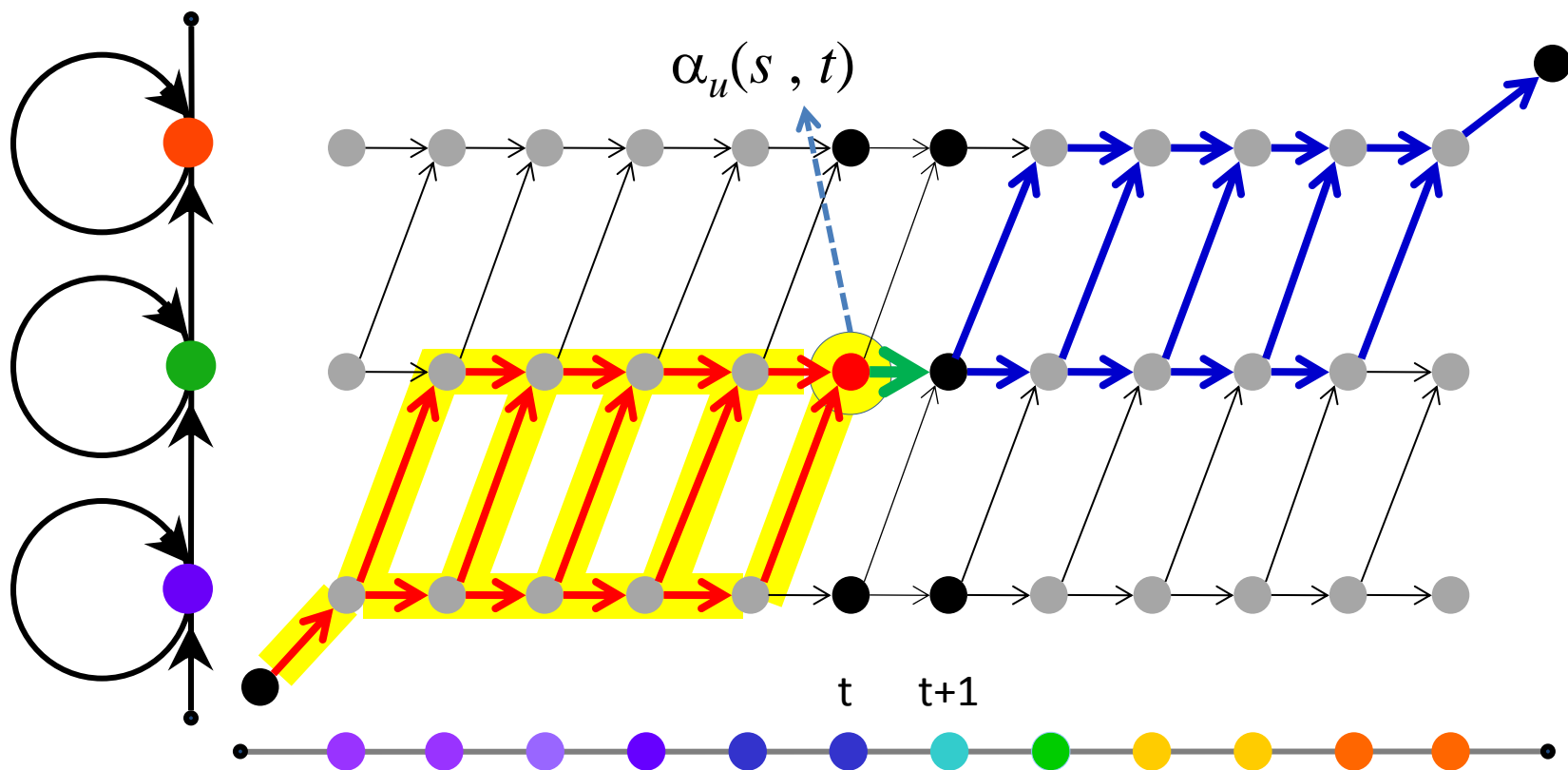
- Contribution of any transition to the count is the *a posteriori* probability of the count
  - This is a fraction
  - The fractions for all possible transitions at any time sum to 1

# Transitions by *soft* counting



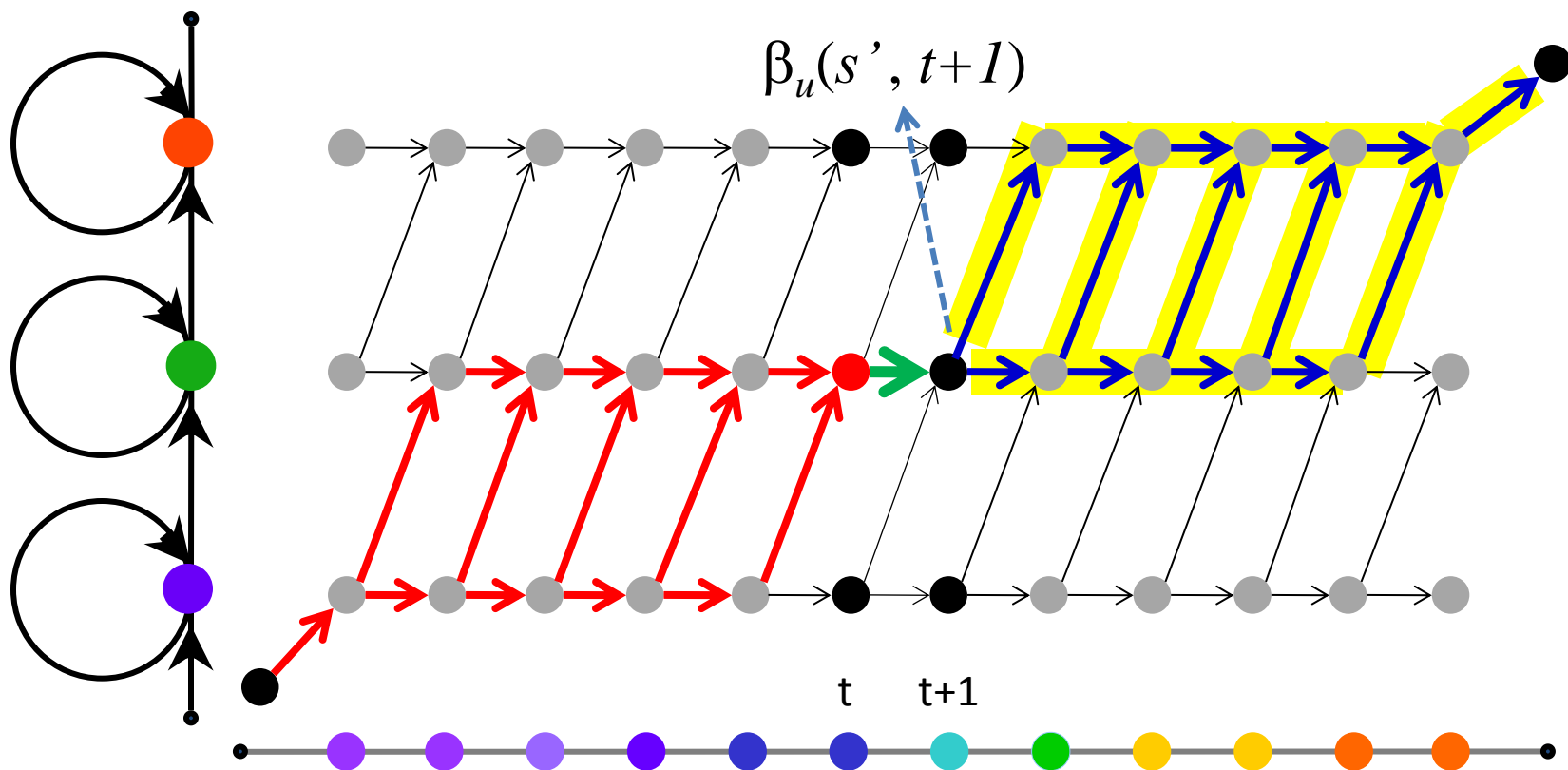
- Probability of a transition is the total probability of all paths that include the transition

# Transitions by *soft* counting



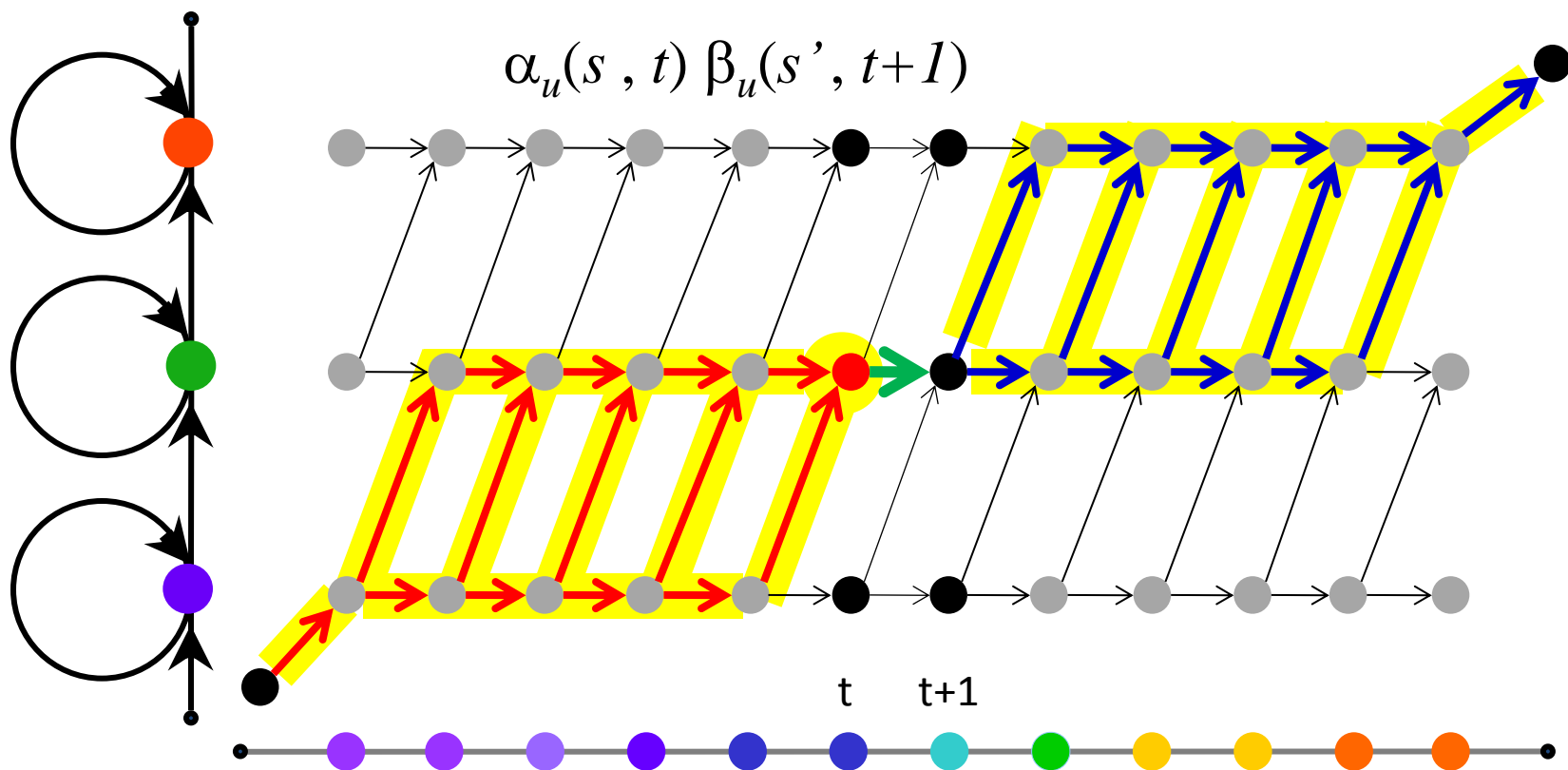
- The forward probability of the source state at  $t$  accounts for all incoming paths at time  $t$ 
  - *including* the  $t$ -th observation  $x_t$

# Transitions by *soft* counting



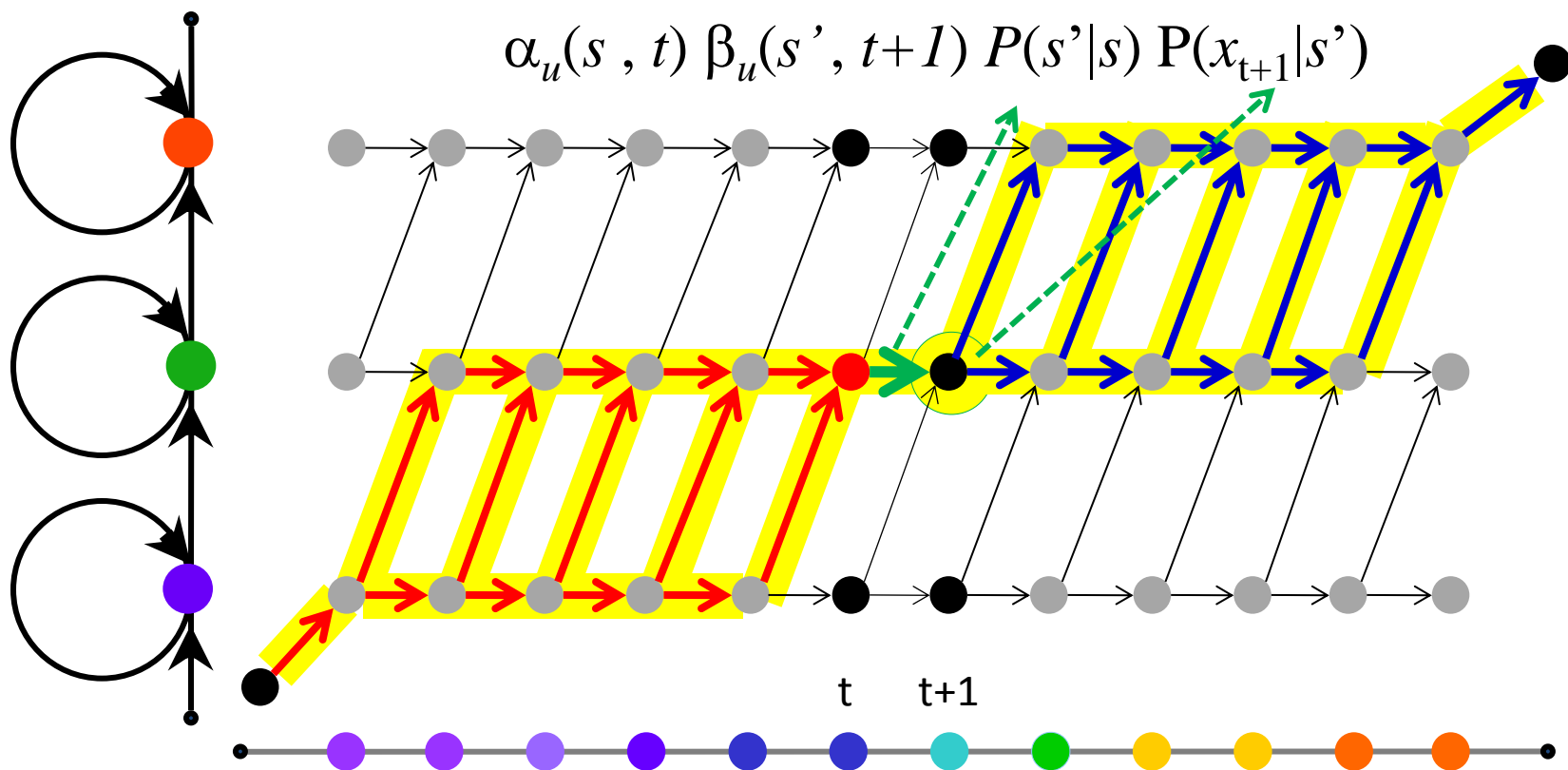
- The *backward* probability of the *destination* state at  $t+1$  accounts for all *outgoing* paths from the state at time  $t+1$ 
  - NOT including** the  $t+1$ -th observation  $x_{t+1}$

# Transitions by *soft* counting



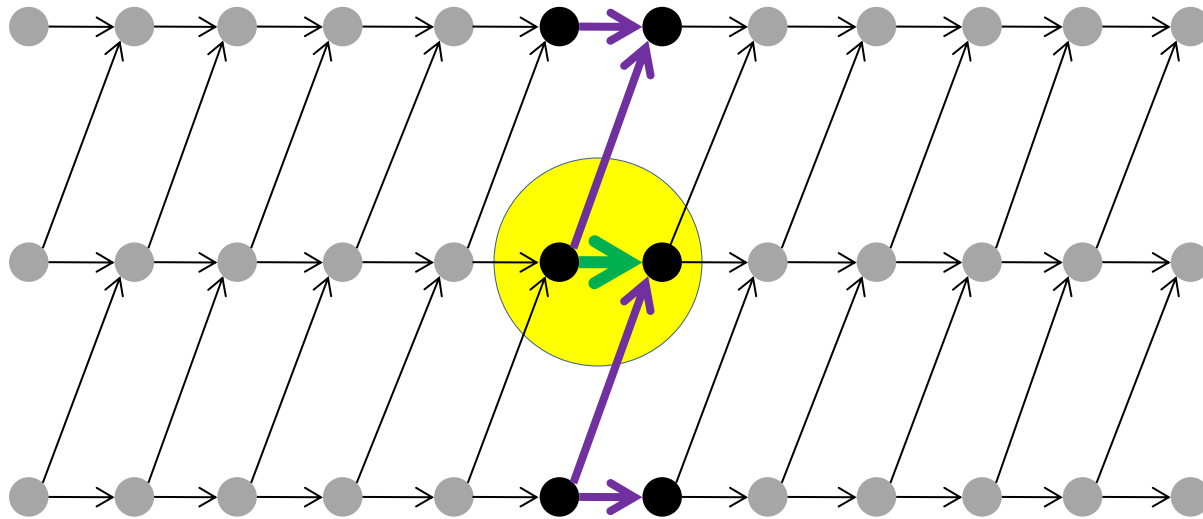
- The product of the forward probability of  $s$  at  $t$  and  $s'$  at  $t+1$  accounts for all paths TO state  $s$  at  $t$ , and all paths FROM  $s'$  at  $t+1$ 
  - But not the *transition* from  $s$  to  $s'$  or the observation at  $t+1$

# Transitions by *soft* counting



- By factoring in the transition probability and observation probabilities, the total *probability* is obtained

# From probability to *a posteriori* probability



- The *a posteriori* probability of a transition is the ratio of its probability to the sum of all transitions at the same time

# A posteriori probability of a transition

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- Probability of a transition

$$P(\text{state}(t) = s, \text{state}(t+1) = s', x_1, x_2, \dots, x_N) = \alpha_u(s, t) P(s' | s) P(x_t + 1 | s') \beta_u(s', t+1)$$

- *A posteriori* probability of a transition

$$\gamma_{u,s,t,s',t+1} = \frac{P(\text{state}(t) = s, \text{state}(t+1) = s', x_1, x_2, \dots, x_N)}{\sum_{S, S'} P(\text{state}(t) = S, \text{state}(t+1) = S', x_1, x_2, \dots, x_N)}$$

$$\gamma_{u,s,t,s',t+1} = \frac{\alpha_u(s, t) P(s' | s) P(x_t + 1 | s') \beta_u(s', t+1)}{\sum_{S, S'} \alpha_u(S, t) P(S' | S) P(x_t + 1 | S') \beta_u(S', t+1)}$$



# Estimate of Transition Probabilities

---

$$P(s' | s) = \frac{\sum_u \sum_t \gamma_{u,s,t,s',t+1}}{\sum_u \sum_t \gamma_{u,s,t}}$$

- Numerator is total “soft” count of transitions from state  $s$  to  $s'$
- Denominator is total “soft” count of instances of state  $s$

# Implementation of BW: underflow

- **Arithmetic underflow is a problem**

$$\alpha_u(s, t) = \sum_{s'} \alpha_u(s', t-1) \underbrace{P(s|s') P(x_{u,t}|s)}_{\text{probability terms}}$$

$\alpha_u(s, t)$  is labeled as a **probability term** with a downward arrow.

- The alpha terms are a recursive product of probability terms
  - As  $t$  increases, an increasingly greater number probability terms are factored into the alpha
- All probability terms are less than 1
  - State output probabilities are actually probability densities
  - Probability density values *can* be greater than 1
  - On the other hand, for large dimensional data, probability density values are usually *much* less than 1
- With increasing time, alpha values decrease
- Within a few time instants, they underflow to 0
  - Every alpha goes to 0 at some time  $t$ . All future alphas remain 0
  - As the dimensionality of the data increases, alphas goes to 0 faster

# Underflow: Solution

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- One method of avoiding underflow is to scale all alphas at each time instant
  - Scale with respect to the largest alpha to make sure the largest scaled alpha is 1.0
  - Scale with respect to the sum of the alphas to ensure that all alphas sum to 1.0
  - Scaling constants must be appropriately considered when computing the final probabilities of an observation sequence
- An alternate method: Compute alphas and betas in log domain
  - How? (Not obvious)

# Implementation of BW: underflow

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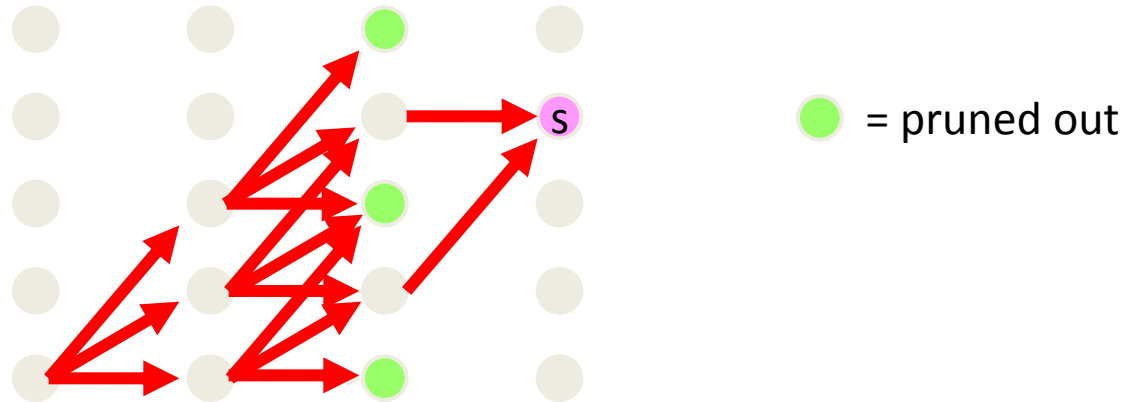
- Similarly, arithmetic underflow can occur during beta computation

$$\beta_u(s, t) = \sum_{s'} \beta_u(s', t+1) P(s' | s) P(x_{u, t+1} | s')$$

- The beta terms are also a recursive product of probability terms and can underflow
- Underflow can be prevented by
  - Scaling: Divide all beta terms by a constant that prevents underflow
  - By performing beta computation in the log domain (now? Not obvious..)
- **QUESTION: HOW DOES SCALING AFFECT THE ESTIMATION OF GAMMA TERMS**
  - For Gaussian parameter updates?
  - For transition probability updates?

# Implementation of BW: pruning

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- The forward backward computation can get very expensive
- Solution: Prune
- Pruning in the forward backward algorithm raises some additional issues
  - Pruning from forward pass can be employed by backward pass
  - Convergence criteria and tests may be affected
  - More later

# Building a recognizer for isolated words

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- Now have all necessary components to build an HMM-based recognizer for isolated words
  - Where each word is spoken by itself in isolation
  - E.g. a simple application, where one may either say “Yes” or “No” to a recognizer and it must recognize what was said

# Isolated Word Recognition with HMMs

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- Assuming all words are equally likely
- Training
  - Collect a set of “training” recordings for each word
  - Compute feature vector sequences for the words
  - Train HMMs for each word
- Recognition:
  - Compute feature vector sequence for test utterance
  - Compute the forward probability of the feature vector sequence from the HMM for each word
    - Alternately compute the best state sequence probability using Viterbi
  - Select the word for which this value is highest

# Issues

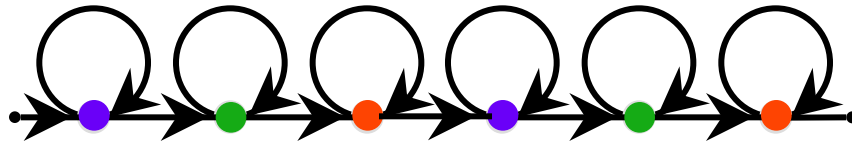
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- What is the topology to use for the HMMs
  - How many states
  - What kind of transition structure
  - If state output densities have Gaussian Mixtures: how many Gaussians?

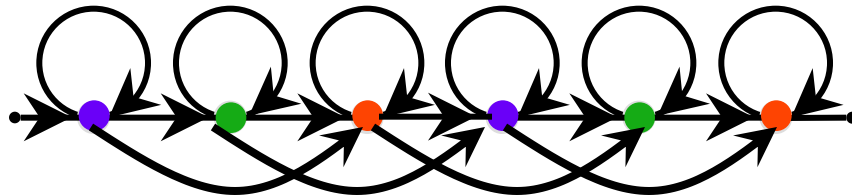


# HMM Topology

- For speech a left-to-right topology works best
  - The “Bakis” topology
  - Note that the initial state probability  $P(s)$  is 1 for the 1<sup>st</sup> state and 0 for others. This need not be *learned*



- States may be skipped



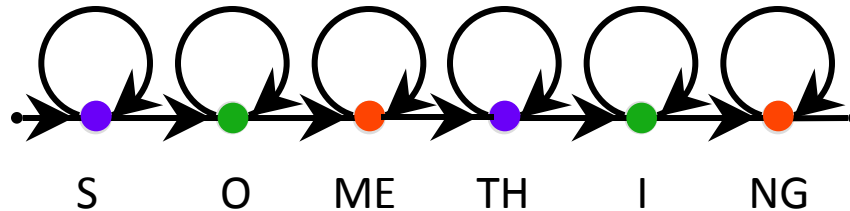
# Determining the Number of States

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- How do we know the number of states to use for any word?
  - We do not, really
  - Ideally there should be at least one state for each “basic sound” within the word
    - Otherwise widely differing sounds may be collapsed into one state
    - The average feature vector for that state would be a poor representation
- For computational efficiency, the number of states should be small
  - These two are conflicting requirements, usually solved by making some educated guesses

# Determining the Number of States

- For small vocabularies, it is possible to examine each word in detail and arrive at reasonable numbers:



- For larger vocabularies, we may be forced to rely on some *ad hoc* principles
  - *E.g.* proportional to the number of letters in the word
    - Works better for some languages than others
    - Spanish and Indian languages are good examples where this works as almost every letter in a word produces a sound

# How many Gaussians

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- No clear answer for this either
- The number of Gaussians is usually a function of the amount of training data available
  - Often set by trial and error
  - A minimum of 4 Gaussians is usually required for reasonable recognition

# Implementation of BW: initialization of alphas and betas

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- Initialization for alpha:  $\alpha_u(s, 1)$  set to 0 for all states except the first state of the model.  $\alpha_u(s, 1)$  set to 1 for the first state
  - All observations *must* begin at the first state
- Initialization for beta:  $\beta_u(s, T)$  set to 0 for all states except the terminating state.  $\beta_u(s, t)$  set to 1 for this state
  - All observations *must* terminate at the final state

# Initializing State Output Density Parameters

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1. Initially only a single Gaussian per state assumed
  - Mixtures obtained by splitting Gaussians
2. For Bakis-topology HMMs, a good initialization is the “flat” initialization
  - Compute the *global* mean and variance of all feature vectors in all training instances of the word
  - Initialize *all Gaussians* (i.e all state output distributions) with this mean and variance
  - Their means and variances will converge to appropriate values automatically with iteration
  - Gaussian splitting to compute Gaussian mixtures takes care of the rest

# Isolated word recognition: Final thoughts

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- All relevant topics covered
  - How to compute features from recordings of the words
    - We will not explicitly refer to feature computation in future lectures
  - How to set HMM topologies for the words
  - How to train HMMs for the words
    - Baum-Welch algorithm
  - How to select the most probable HMM for a test instance
    - Computing probabilities using the forward algorithm
    - Computing probabilities using the Viterbi algorithm
      - Which also gives the state segmentation

# Questions

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- ?